

*Assessing Health Disparities in Intensive Longitudinal Data:  
Gender Differences in Granger Causality Between Primary Care  
Provider and Emergency Room Usage,  
Assessed with Medicaid Insurance Claims*

May 24, 2017

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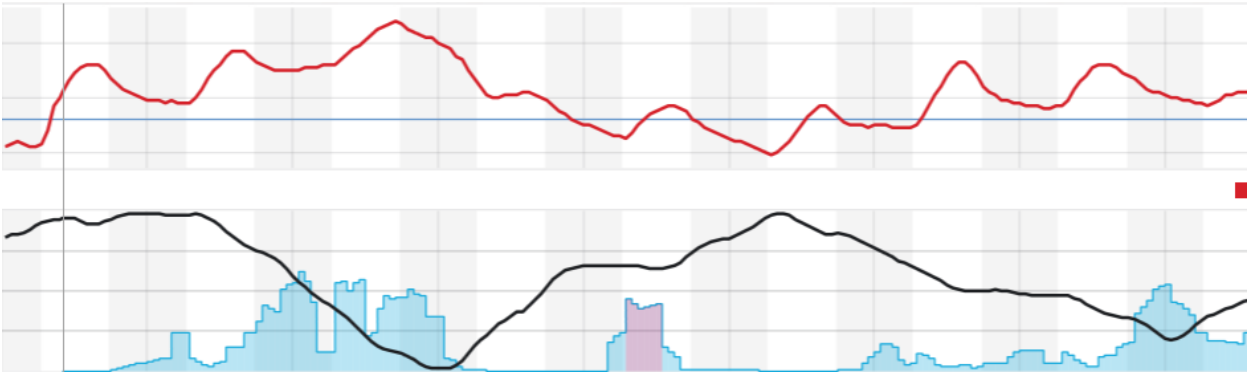
Granger causality

Goals

1. Introduce an analytical method to 'extract direction of causal effects' from time series/panel data/intensive longitudinal data
2. Show software code and output
3. Interpret results

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## A daily example



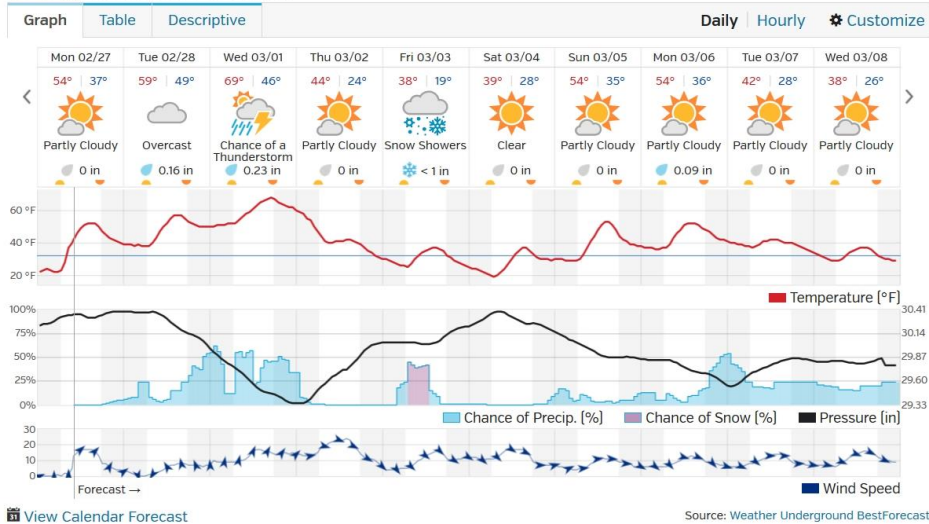
Guess?

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## A daily example

[https://www.wunderground.com/q/zmw06001.9.99999?sp=MC0728&utm\\_source=HomeCard&utm\\_content=Button&cm\\_ven=HomeCar](https://www.wunderground.com/q/zmw06001.9.99999?sp=MC0728&utm_source=HomeCard&utm_content=Button&cm_ven=HomeCar) DOI: 10.1056/NEIMsa1610524

### 10-Day Weather Forecast



1. Wunderground

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## Chicken-egg

“To conclude that one of the two "came first," we must find unidirectional causality from one to the other.

In other words, we must reject the *noncausality* of the one to the other and at the same time fail to reject the *noncausality* of the other to the one.”

1. Thurman, W. N., & Fisher, M. E. (1988). Chickens, eggs, and causality, or which came first. *American journal of agricultural economics*, 70(2), 237-238.

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## Granger *noncausality*

"Granger (1969, 1980, 1988) proposed a notion of causality that is based on the following important points: the cause occurs before the effect and the cause must contain unique information that is otherwise not available and helps to predict the effect.“

Eichler, M., & Didelez, V. (2010). On Granger causality and the effect of interventions in time series. *Lifetime Data Analysis*, 16(1), 3-32. doi:10.1007/s10985-009-9143-3

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## Granger *noncausality*

“We say that X Granger-causes Y if the current value of Y can be better predicted from the past values of all three series X, Y, and Z than from the past values of the two processes Y and Z alone.

Here, "better predicted" means a smaller mean square prediction error.

We note that the definition depends on the set of variables Z included in the analysis.”

Tschacher W, Ramseyer F: Modeling psychotherapy processes by time-series panel analysis (TSPA). *Psychother Res* 2009, 19:469-481

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## Granger *noncausality*

*Example 2.2* Consider the following multivariate Gaussian process X with

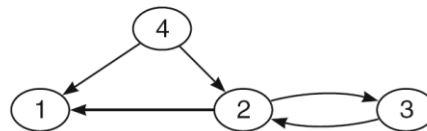
$$X_1(t) = \alpha_1 X_4(t - 2) + \beta_{12} X_2(t - 1) + \varepsilon_1(t),$$

$$X_2(t) = \alpha_2 X_4(t - 1) + \beta_{23} X_3(t - 1) + \varepsilon_2(t),$$

$$X_3(t) = \beta_{32} X_2(t - 1) + \varepsilon_3(t),$$

$$X_4(t) = \varepsilon_4(t)$$

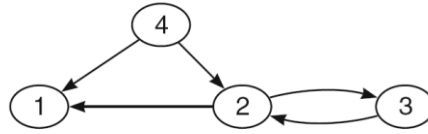
where  $\varepsilon_i$ ,  $i = 1, 2, 3, 4$ , are independent Gaussian white noise processes with mean 0 and variance  $\sigma^2$ .



Eichler, M., & Didelez, V. (2010). On Granger causality and the effect of interventions in time series. *Lifetime Data Analysis*, 16(1), 3-32. doi:10.1007/s10985-009-9143-3

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## Granger noncausality



“It is **immediately obvious** that with respect to the set  $\{X_1, X_2, X_3, X_4\}$ , the component  $X_1$  is Granger noncausal for all other variables, and for instance  $X_3$  is Granger noncausal for  $X_1$ .

However, with respect to the reduced set  $\{X_1, X_3, X_4\}$  it cannot be assumed anymore that  $X_3$  is Granger noncausal for  $X_1$ .

It is less intuitive, but also clear from the full model that  $X_3(t)$  is also not Granger noncausal anymore for  $X_1$  with respect to  $\{X_1, X_2, X_3\}$  due to the selection effect when conditioning on the past of  $X_2$  which induces an association between  $X_3$  and  $X_4$ .”

Eichler, M., & Didelez, V. (2010). On Granger causality and the effect of interventions in time series. *Lifetime Data Analysis*, 16(1), 3-32. doi:10.1007/s10985-009-9143-3

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## Chicken-egg real example

The following equation was estimated by OLS:

$$Eggs_t = \mu + \sum_{i=1}^L \alpha_i Eggs_{t-i} = \sum_{i=1}^L \beta_i Chickens_{t-i} + \epsilon_t;$$

$H_0 : \beta_1 = \dots = \beta_L = 0$  (chickens do not Granger cause eggs).

<u><math>L = \text{no. of lags}</math></u>	<u><math>F\text{-statistic}</math></u>	<u><math>P\text{-value}</math></u>	<u><math>R^2</math> of the regression</u>
1	.04	.85	.96
2	1.71	.19	.97
3	1.10	.36	.97
4	.79	.54	.97

“The validity of our test statistic requires lack of serial correlation, homoskedasticity, and normality of the disturbances in the distributed lag equations, which we of course assume.”

1. Thurman, W. N., & Fisher, M. E. (1988). Chickens, eggs, and causality, or which came first. *American journal of agricultural economics*, 70(2), 237-238.

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## Chicken-egg

### Part 2: Did the Egg Come First?

The following equation was estimated by OLS:

$$Chickens_t = \mu + \sum_{i=1}^L \alpha_i Chickens_{t-i} + \sum_{i=1}^L \beta_i Eggs_{t-i};$$

$H_0: \beta_1 = \dots = \beta_L = 0$  (eggs do not Granger cause chickens).

“The validity of our test statistic requires lack of serial correlation, homoskedasticity, and normality of the disturbances in the distributed lag equations, which we of course assume.”

<u>L = no. of lags</u>	<u>F-statistic</u>	<u>P-value</u>	<u>R<sup>2</sup> of the regression</u>
1	1.23	.27	.73
2	10.36	.0002	.81
3	5.85	.0019	.81
4	4.71	.0032	.82

Data source: U.S. Department of Agriculture, 1983 and others.  
Note: The data are annual, 1930–83.

1. Thurman, W. N., & Fisher, M. E. (1988). Chickens, eggs, and causality, or which came first. *American journal of agricultural economics*, 70(2), 237-238.

## Chicken-egg

“While our test is agnostic regarding this instantaneous causality, we suspect that **eggs are endogenous** in the sense that chickens cause eggs within the sampling period.

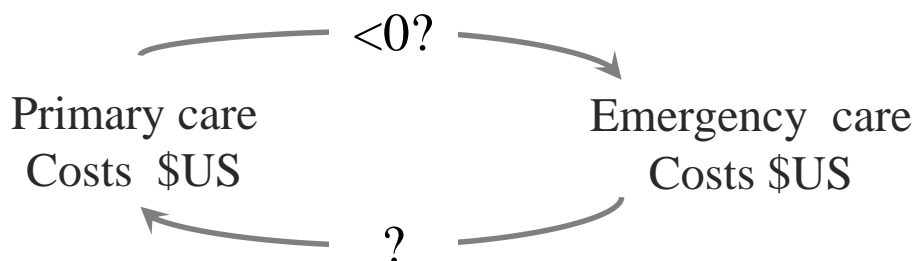
A Wu-Hausman test of the predeterminedness of eggs could address the issue and would require a valid instrumental variable (correlated with eggs and uncorrelated with the chicken forecast error), perhaps bacon.”

### “Suggestions for Future Research

The **structural** implications of our results are not yet clear. To draw them out fully will require collaboration between economists and poultry scientists. The potential here is great.”

1. Thurman, W. N., & Fisher, M. E. (1988). Chickens, eggs, and causality, or which came first. *American journal of agricultural economics*, 70(2), 237-238.

## Illustration



Medicaid costs:

Does spending for primary health care  
*reduce* later on emergency care costs?

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## Granger noncausality – applied in Stata

"Time-series vector autoregression ( VAR ) models originated in the macroeconometrics literature as an alternative to multivariate simultaneous equation models (Sims 1980)."

"Panel VAR model selection, estimation, and inference in a generalized method of moments ( GMM ) framework."

"Model selection, estimation, and inference about the homogeneous panel VAR model above can be implemented with the new commands *pvar*, *pvarsoc*, *pvargranger*, *pvarstable*, *pvarirf*, and *pvarfevd*. The syntax and outputs are closely patterned after Stata's built-in *var* commands to easily switch between panel and time-series VAR."

1. Abrigo, M. R., & Love, I. (2016). Estimation of panel vector autoregression in Stata. *Stata Journal*, 16(3), 778-804.

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## Granger noncausality – applied in Stata

"*pvar* fits homogeneous panel VAR models by fitting a multivariate panel regression of each dependent variable on lags of itself, lags of all other dependent variables, and lags of exogenous variables, if any. The estimation is by GMM. The command is implemented using the interactive version of Stata's *gmm* command with analytic derivatives."

"*pvarsoc* provides various summary measures to aid the process of panel VAR model selection."

"The postestimation command *pvargranger* performs Granger causality Wald tests for each equation of the underlying panel VAR model. It provides a convenient alternative to Stata's built-in test command."

"The postestimation command *pvarstable* checks the stability condition of panel VAR estimates by calculating the modulus of each eigenvalue of the fitted model."

"The postestimation command *pvarirf* calculates and plots IRF s(impulse–response functions)."

"The postestimation command *ichecks* the stability condition of panel VAR estimates by calculating the modulus of each eigenvalue of the fitted model."

"The postestimation command *pvarfevd* computes FEVD based on a Cholesky decomposition of the residual covariance matrix of the underlying panel VAR model. Standard errors and confidence intervals based on Monte Carlo simulation may be optionally computed."

1. Abrigo, M. R., & Love, I. (2016). Estimation of panel vector autoregression in Stata. *Stata Journal*, 16(3), 778-804.

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## Medicaid claims data

14,358 patients in CT (238404=63.79% female)

386,613 claims, up to 45 months.

Average age 44 years old

```
. xtsum age
```

Variable		Mean	Std. Dev.	Min	Max	Observations
age	overall	44.12122	17.60198	0	100	N = 373725
	between		18.80543	0	100	n = 14358
	within		0	44.12122	44.12122	T-bar = 26.029

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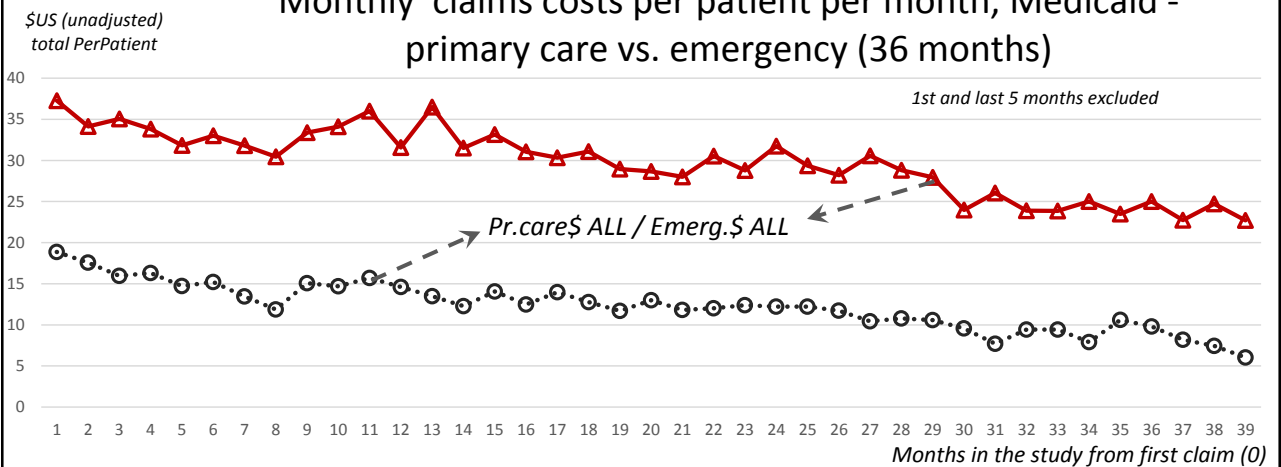
# Medicaid claims Primary care (1) vs. Emergency (2)

92 Family Nurse Practitioner	1
99 Community Health Nurse Practitioner	1
124 Primary Care Nurse Practitioner	1
306 Preventative Medicine	1
318 General Practice Medicine	1
345 General Pediatrics	1
521 Medical FQHC & Tribal Svs Medical FQHC	1
997 Primary Care Physician Assistant	1
100 Critical Care Nurse Practitioner	2
102 Neonatal Critical Care Nurse Practitioner	2
104 Pediatric Critical Care Nurse Practitioner	2
260 Ambulance	2
261 Air Ambulance	2
262 Critical Care Helicopter	2
315 Emergency Medicine	2
611 Pediatric Emergency Department Medicine	2
612 Pediatric Emergency Medicine	2
621 Pediatric Critical Care Medicine	2

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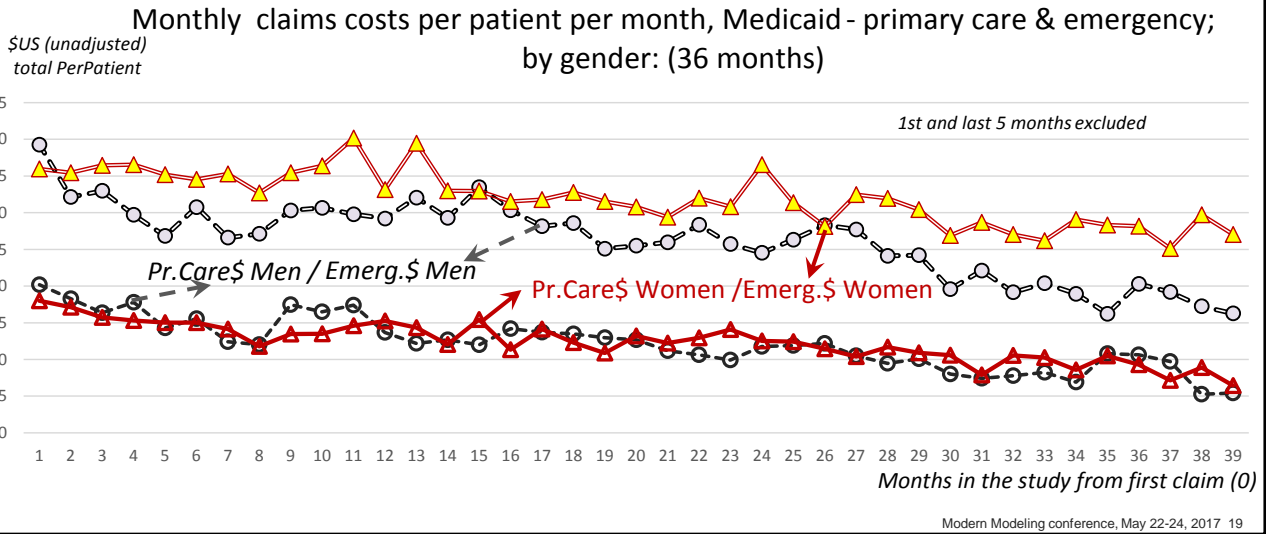
# Medicaid claims Primary care vs. Emergency

Monthly claims costs per patient per month, Medicaid - primary care vs. emergency (36 months)

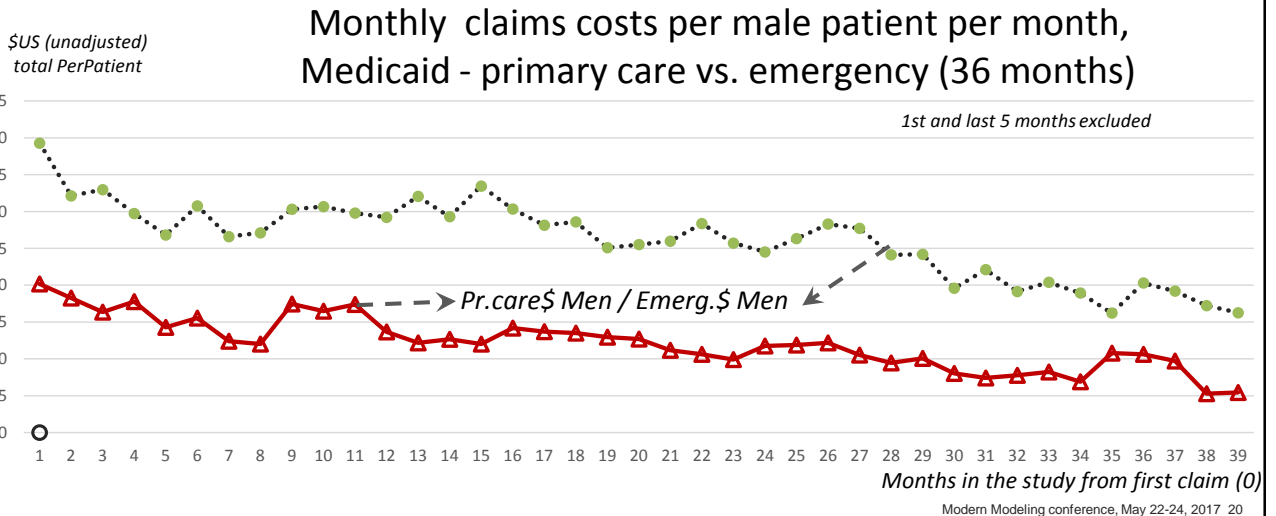


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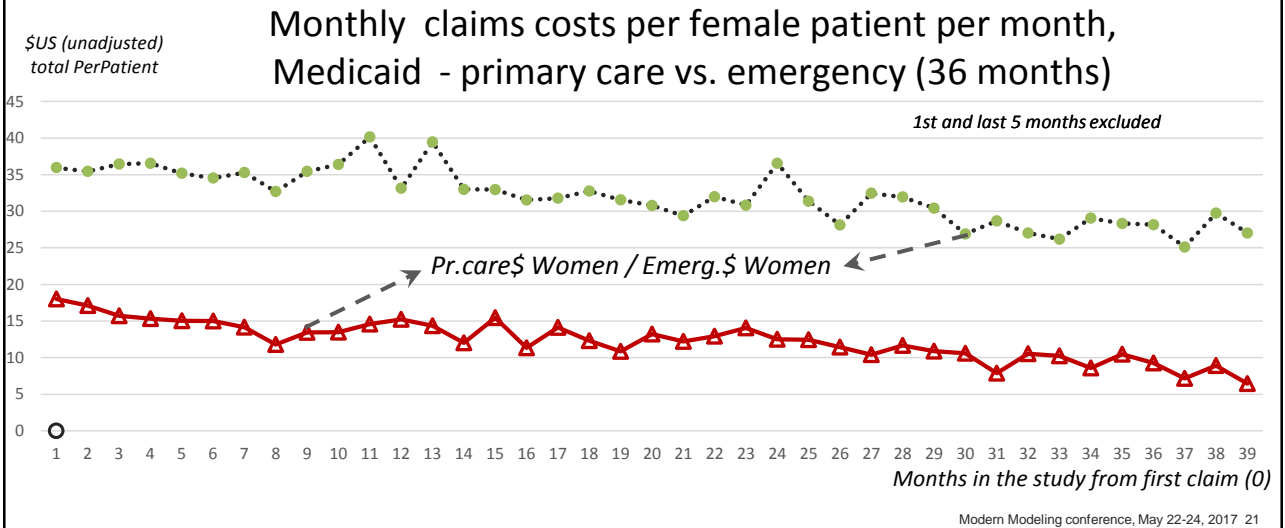
## Medicaid claims by gender



## Medicaid claims - males



## Medicaid claims - females



## Panel Regression

Emergency \$s-[-.018/.001\*]->Primary care \$s

```
xtreg dolprimc [Primary care $s] dolemerng [Emergency $s]
```

Random-effects GLS regression

Group variable: id

R-sq:

within = 0.0004

between = 0.0007

overall = 0.0004

Number of obs = 373,725

Number of groups = 14,358

Obs per group:

min = 1

avg = 26.0

max = 45

Wald chi2(1) = 146.57

Prob > chi2 = 0.0000

corr(u\_i, X) = 0 (assumed)

	Coef.	Std. Err.	z	P> z	[95% Conf. Interval]
dolemerng	-.0181991	.0015032	-12.11	0.000	-.0211454 - .0152528
_cons	24.33821	.5940643	40.97	0.000	23.17387 25.50256

## Panel Regression

Primary care \$s -[-.021/.002\*]->Emergency \$s

```

xtreg dolemerg [Emergency $s] dolprimc [Primary care $s]
Random-effects GLS regression      Number of obs   =   373,725
Group variable: id                 Number of groups =   14,358
R-sq:                               Obs per group:
    within = 0.0004                    min =           1
    between = 0.0007                    avg =          26.0
    overall = 0.0004                    max =           45
                                     Wald chi2(1)     =   146.31
                                     Prob > chi2      =    0.0000
corr(u_i, X) = 0 (assumed)

```

	Coef.	Std. Err.	z	P> z	[95% Conf. Interval]
dolemerg					
dolprimc	<b>-.0214857</b>	.0017763	-12.10	0.000	-.0249672 - .0180043
_cons	47.62394	.6888659	69.13	0.000	46.27378 48.97409

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*pvar* Emergency \$s ->Primary care \$s

```

. pvar dolprimc dolemerg , lags(6) No. of obs   =   168681
No. of panels =   9644 Ave. no. of T =   17.491

```

	Coef.	Std. Err.	z	P> z	[95% Conf. Interval]
dolprimc					
dolprimc					
L1.	.2746698	.1290072	2.13	0.033	.0218204 .5275192
L2.	-.0485268	.0453748	-1.07	0.285	-.1374597 .0404062
L3.	.0190768	.0153223	1.25	0.213	-.0109544 .049108
L4.	-.0008771	.0096073	-0.09	0.927	-.0197071 .0179529
L5.	-.0072563	.0102543	-0.71	0.479	-.0273544 .0128418
L6.	.0043254	.0049237	0.88	0.380	-.005325 .0139757
dolemerg					
L1.	<b>.0052949</b>	.0023697	2.23	<b>0.025</b>	.0006504 .0099395
L2.	.001567	.0016548	0.95	0.344	-.0016764 .0048104
L3.	.0003862	.0009709	0.40	0.691	-.0015167 .0022891
L4.	<b>.0017168</b>	.0009404	1.83	<b>0.068</b>	-.0001262 .0035599
L5.	.001087	.0010514	1.03	0.301	-.0009737 .0031477
L6.	.0012177	.0009675	1.26	0.208	-.0006785 .0031139

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## Primary care \$s ->Emergency \$s

```
. pvar dolprimc dolemerg , lags(6) No. of obs = 168681
No. of panels = 9644 Ave. no. of T = 17.491
      Coef.   Std. Err.      z    P>|z|    [95% Conf. Interval]
dolemerg |
  dolprimc |
    L1. | .0047209   .0031455    1.50  0.133   -.0014441   .0108859
    L2. | .006232   .003401    1.83  0.067  -.0004337   .0128978
    L3. | .0032212   .0028914    1.11  0.265   -.0024459   .0088882
    L4. | .0098327   .004332    2.27  0.023   .0013422   .0183233
    L5. | -.0036448   .0039193   -0.93  0.352   -.0113265   .0040368
    L6. | .0245599   .0061225    4.01  0.000   .0125601   .0365598
dolemerg |
  L1. | .1230638   .0233041    5.28  0.000   .0773886   .1687391
  L2. | .0796443   .0134472    5.92  0.000   .0532882   .1060004
  L3. | .0649163   .010089    6.43  0.000   .0451422   .0846904
  L4. | .0522287   .0124232    4.20  0.000   .0278798   .0765777
  L5. | .0411237   .0113239    3.63  0.000   .0189292   .0633181
  L6. | .0303653   .0088422    3.43  0.001   .0130349   .0476957
```

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## *pvargranger* Primary care \$s -!\*->Emergency \$s

```
. pvargranger
```

```
panel VAR-Granger causality Wald test
```

```
Ho: Excluded variable does not Granger-cause Equation variable
```

```
Ha: Excluded variable Granger-causes Equation variable
```

```
+-----+
| Equation \ Excluded |      chi2    df    Prob > chi2 |
+-----+-----+
| dolprimc           |             |             |
|      dolemerg      |      9.425    6      0.151 |
|           ALL      |      9.425    6      0.151 |
+-----+-----+
| dolemerg           |             |             |
|      dolprimc      |     19.265    6      0.004 |
|           ALL      |     19.265    6      0.004 |
+-----+-----+
```

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## Males *neither* vs. Females PrC-!\*->Em.

+-----+ MALES				
Equation \ Excluded	chi2	df	Prob >	chi2
dolprimc				
dolemerg	6.326	6	0.388	
ALL	6.326	6	0.388	
dolemerg				
dolprimc	4.375	6	0.626	
ALL	4.375	6	0.626	
+-----+ FEMALES				
Equation \ Excluded	chi2	df	Prob >	chi2
dolprimc				
dolemerg	7.462	6	0.280	
ALL	7.462	6	0.280	
dolemerg				
dolprimc	113.310	6	0.000	
ALL	113.310	6	0.000	

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## Primary care \$\$ <-> Emergency \$\$ by gender

```
. pvar dolprimc dolemerg , lags(6)
[ONLY approaching sig. shown]
```

### MALES

	Coef.	Std. Err.	z	P> z	[95% Conf. Interval]	
--	-------	-----------	---	------	----------------------	--

PrimC\$\$ = DV

  Emerg\$\$

L1.		.01222	.007332	1.67	0.096	-.0021505	.0265905
-----	--	--------	---------	------	-------	-----------	----------

### FEMALES

PrimC\$\$ = DV

  Emerg\$\$

L3.		.0018045	.0009995	1.81	0.071	-.0001545	.0037635
L4.		.0017593	.0010084	1.74	0.081	-.0002171	.0037357

Emerg\$\$ =Dv

  PrimC\$\$

L2.		.0169106	.0061186	2.76	0.006	.0049184	.0289028
L6.		.0295973	.0028833	10.27	0.000	.0239461	.0352485

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## [ Conclusions – email for >: [comanus@gmail.com](mailto:comanus@gmail.com) ]

1. Causal dynamic processes captured by costs of care may differ by gender
2. The pattern of differences merits further investigations
3. Causality with such intensive data offers some opportunities

"As a limitation, we have to emphasize that Granger causality is not true causality."