Applying Modern Methods for Missing Data Analysis to the Social Relations Model

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Social Relations Model

- Dyadic data from round-robin design (Kenny, 1994)
 - e.g., all students in a classroom rate each other
 - Block design: heterosexual speed-daters rate opposite sex
- Originally used ANOVA to partition ratings into personand dyad-level components: $x_{ij} = \mu + P_i + T_j + R_{ij}$
 - \cap μ = average rating
 - P_i = perceiver *i*'s tendency to rate above/below μ
 - T_i = target j's tendency to elicit ratings above/below μ
 - R_{ij} = residual, contains dyadic relationship effect and error

• All participants rate their perceptions of one another

Targets \rightarrow \checkmark Perceivers	Alice	Betty	Cathy	Daria	Ellen	X _{Row}	Perceiver Effect
Alice	Self	11	6	9	7	8.25	3.25
Betty	2	Self	2	5	4	3.25	-1.75
Cathy	3	11	Self	8	4	6.50	1.50
Daria	4	4	1	Self	2	2.75	-2.25
Ellen	4	7	2	4	Self	4.25	-0.75
\overline{X}_{Column}	3.25	8.25	2.75	6.50	4.25	$\overline{X} = 5$	
Target Effect	-1.75	3.25	-2.25	1.50	-0.75		

• μ is estimated from the average of all ratings (\overline{X})

Targets \rightarrow \checkmark Perceivers	Alice	Betty	Cathy	Daria	Ellen	\overline{X}_{Row}	Perceiver Effect
Alice	Self	11	6	9	7	8.25	3.25
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Cathy	3	11	Self	8	4	6.50	1.50
Daria	4	4	1	Self	2	2.75	-2.25
Ellen	4	7	2	4	Self	4.25	-0.75
<u></u> X _{Column}	3.25	8.25	2.75	6.50	4.25	$\overline{X} = 5$	
Target Effect	-1.75	3.25	-2.25	1.50	-0.75		

• Row means indicate each perceiver's average rating

Targets \rightarrow \checkmark Perceivers	Alice	Betty	Cathy	Daria	Ellen	\overline{X}_{Row}	Perceiver Effect
Alice	Self	11	6	9	7	8.25	3.25
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Cathy	3	11	Self	8	4	6.50	1.50
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<u></u> X _{Column}	3.25	8.25	2.75	6.50	4.25	$\overline{X} = 5$	
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• Perceiver effects (P_i) are their averages relative to the grand mean

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\overline{X}_{Column}	3.25	8.25	2.75	6.50	4.25	$\overline{X} = 5$	
Target Effect	-1.75	3.25	-2.25	1.50	-0.75		

• Column means indicate the average rating received by each target

Targets → ↓ Perceivers	Alice	Betty	Cathy	Daria	Ellen	\overline{X}_{Row}	Perceiver Effect
Alice	Self	11	6	9	7	8.25	3.25
Betty	2	Self	2	5	4	3.25	-1.75
Cathy	3	11	Self	8	4	6.50	1.50
Daria	4	4	1	Self	2	2.75	-2.25
Ellen	4	7	2	4	Self	4.25	-0.75
\overline{X}_{Column}	3.25	8.25	2.75	6.50	4.25	$\overline{X} = 5$	
Target Effect	-1.75	3.25	-2.25	1.50	-0.75		

• Target effects (T_j) are their averages relative to the grand mean

Targets → ↓ Perceivers	Alice	Betty	Cathy	Daria	Ellen	\overline{X}_{Row}	Perceiver Effect
Alice	Self	11	6	9	7	8.25	3.25
Betty	2	Self	2	5	4	3.25	-1.75
Cathy	3	11	Self	8	4	6.50	1.50
Daria	4	4	1	Self	2	2.75	-2.25
Ellen	4	7	2	4	Self	4.25	-0.75
\overline{X}_{Column}	3.25	8.25	2.75	6.50	4.25	$\overline{X} = 5$	
Target Effect	-1.75	3.25	-2.25	1.50	-0.75		

• Residuals (R_{ij}) are the differences between observed and expected ratings, given μ , P_i , and T_j

∩ $R_{\text{Alice} \rightarrow \text{Betty}} = 11 - (5 + 3.25 + 3.25) = 11 - 11.5 = -0.5$

Targets → ↓ Perceivers	Alice	Betty	Cathy	Daria	Ellen	\overline{X}_{Row}	Perceiver Effect
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\bar{X}_{Column}	3.25	8.25	2.75	6.50	4.25	$\overline{X} = 5$	
Target Effect	-1.75	3.25	-2.25	1.50	-0.75		

• Self-ratings (or expert observations) can also be recorded, to calculate self-other agreement (or accuracy)

Targets → ↓ Perceivers	Alice	Betty	Cathy	Daria	Ellen	$ar{X}_{Row}$	Perceiver Effect
Alice	Self	11	6	9	7	8.25	3.25
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\overline{X}_{Column}	3.25	8.25	2.75	6.50	4.25	$\overline{X} = 5$	
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Estimating SRM with ML

- More recently, maximum likelihood estimation is available by specifying SRM as a random-effects / multilevel model (Snijders & Kenny, 1999)
 - Requires many dummy codes and tedious equality constraints in MLwiN (unavailable in most other multilevel software)
 - Restrictive assumptions in SAS PROC MIXED
- Can also be specified as an "n-level SEM" (Brunson et al., 2016) in R package xxM (Mehta, 2013)
- Restricted ML recently proposed (Nestler, 2016)

Estimating SRM with MCMC

- Bayesian estimation first proposed by Hoff (2005), more recently by Lüdtke et al. (2013)
- Using MCMC estimation, random effects of perceiver (P_i) and target (T_j) are estimated as parameters, along with variance-component hyperparameters
 - Data augmentation
 - Differs from marginalizing over random effects in ML estimation of random-effects / multilevel models
- Major advantage: posterior distribution of (functions of) all parameters

Missing Data

- ANOVA calculations using all available data might work if data are missing completely at random (MCAR)
 Very restrictive assumption
- State-of-the-art missing-data methods are maximum likelihood and multiple imputation
 - Only assumes data are missing at random (MAR), given the observed data in the model
 - More defensible if the model includes variables that explain missingness or correlate with the missing values
 - Auxiliary variables: not of theoretical interest, but useful to justify MAR assumption

Missing Data

- Multiple imputation has a Bayesian foundation
- Can be done by augmenting observed data with missing data, just like estimating latent variables
 - e.g., random effects, factor scores
 - Usually an "unrestricted" imputation model (e.g., NORM)
 - Freely estimated mean vector and covariance matrix
- Can easily incorporate into the SRM model
 - More efficient than unrestricted imputation model (Merkle, 2011)

Missing Data

- ML simply evaluates the likelihood function using all available data
 - Advantage: no need to "do anything" with the missing data, but may need to incorporate auxiliaries
 - Disadvantage: exogenous predictors must be complete
- Handling of missing data has been touted as an advantage of ML and MCMC estimation
 - Brunson et al. (2016), Hoff (2005), Lüdtke et al. (2013), Nestler (2016), Snijders & Kenny (1999)
 - No one yet described how to do so in MCMC, so I submitted an application to Social Networks (under review)

• From the SRM equation $Y_{ij} = \mu + P_i + T_j + E_{ij}$

• The expected value of the vector of both observations within a dyad is

$$\widehat{\mathbf{Y}}_{\{ij\}} = \begin{bmatrix} \widehat{Y}_{ij} \\ \widehat{Y}_{ji} \end{bmatrix} = \begin{bmatrix} \mu + P_i + T_j \\ \mu + P_j + T_i \end{bmatrix}$$

 Data might be missing for one or both observations within a dyad

• The likelihood of a dyad's vector is bivariate normal with mean equal to the expected value of the vector, and covariance matrix of residuals:

$$\begin{bmatrix} Y_{ij} \\ Y_{ji} \end{bmatrix} \sim \mathcal{N} \left(\begin{bmatrix} \widehat{Y}_{ij} \\ \widehat{Y}_{ji} \end{bmatrix}, \begin{bmatrix} \sigma_E^2 & \\ \rho_E \sigma_E & \sigma_E^2 \end{bmatrix} \right)$$

- Observed data can be augmented with estimates of missing values by using this likelihood of observed data as the prior for missing-data estimates
 - Assumes MAR, conditional on expected values

- Even complete data are augmented with estimates of random effects, distributed bivariate normally: $\begin{bmatrix}
 P_i \\
 T_i
 \end{bmatrix} \sim \mathcal{N}\left(\begin{bmatrix}0 \\
 0\end{bmatrix}, \begin{bmatrix}\sigma_P^2 \\ \sigma_P^2 \\ \rho_{PT}\sigma_P\sigma_T & \sigma_T^2\end{bmatrix}\right)$
- This can be extended to include 1 or more auxiliary covariates (X_i) :

 $\begin{bmatrix} P_i \\ T_i \\ X_i \end{bmatrix} \sim \mathcal{N} \left(\begin{bmatrix} 0 \\ 0 \\ \mu_X \end{bmatrix}, \begin{bmatrix} \sigma_P^2 & & \\ \rho_{TP} \sigma_T \sigma_P & \sigma_T^2 & \\ \rho_{XP} \sigma_X \sigma_P & \rho_{XT} \sigma_X \sigma_T & \sigma_X^2 \end{bmatrix} \right)$

 Covariates might also be substantively interesting as predictors of the random effects:

 $P_i = \beta_1^P X_i + \varepsilon_i \qquad , \qquad T_j = \beta_1^T X_j + \delta_j$

• In which case the covariate(s) would be independent of the residual perceiver and target effects:

$$\begin{bmatrix} \varepsilon_i \\ \delta_i \\ X_i \end{bmatrix} \sim \mathcal{N} \left(\begin{bmatrix} 0 \\ 0 \\ \mu_X \end{bmatrix}, \begin{bmatrix} \sigma_{\varepsilon}^2 & & \\ \sigma_{\varepsilon\delta} & \sigma_{\delta}^2 & \\ 0 & 0 & \sigma_X^2 \end{bmatrix} \right)$$

- The MAR assumption is more defensible if explanatory or auxiliary covariates can be included that either:
 - Explain missingness
 - Correlate with missing values
- Dyad-level covariates can also be incorporated in the model, can could be either:
 - Constant within dyad ($V_{ij} = V_{ji}$; e.g., "How long have you known each other?")
 - Vary within dyad (W_{ij} ≠ W_{ji}; e.g., "How well does [this friend] know you?")

• Dyad-level covariates can be merely auxiliary:

 $\begin{bmatrix} Y_{ij} \\ Y_{ji} \\ Y_{ij} \\ W_{ij} \\ W_{ij} \\ W_{ii} \end{bmatrix} \sim \mathcal{N} \left(\begin{bmatrix} \hat{Y}_{ij} \\ \hat{Y}_{ji} \\ \mu_V \\ \mu_W \\ \mu_$

- So including them would make the MAR assumption 0 more tenable
- Note that the equality constraints are unnecessary if 0 persons *i* and *j* have roles (e.g., men rating women)

• Dyad-level covariates can also be explanatory: $\begin{bmatrix} \hat{Y}_{ii} \end{bmatrix} \begin{bmatrix} \mu + \gamma_1 V_{\{ii\}} + \gamma_2 W_{ii} + P_i + T_i \end{bmatrix}$

$$\mathbf{\hat{Y}}_{\{ij\}} = \begin{bmatrix} i \\ \hat{Y}_{ji} \end{bmatrix} = \begin{bmatrix} i & i \\ \mu + \gamma_1 V_{\{ij\}} + \gamma_2 W_{ji} + P_j + T_i \end{bmatrix}$$

- In which case they should not correlate with residuals: $\begin{bmatrix}
 Y_{ij} \\
 Y_{ji} \\
 V_{ij} \\
 W_{ij} \\
 W_{ji}
 \end{bmatrix} \sim \mathcal{N} \left(\begin{bmatrix}
 \hat{Y}_{ij} \\
 \hat{Y}_{ji} \\
 \mu_V \\
 \mu_W \\
 \mu_W
 \end{bmatrix}, \begin{bmatrix}
 \sigma_R^2 \\
 \rho_R \sigma_R & \sigma_R^2 \\
 0 & 0 & \sigma_V^2 \\
 0 & 0 & \rho_{WV} \sigma_W \sigma_V & \sigma_W^2 \\
 0 & 0 & \rho_{WV} \sigma_W \sigma_V & \sigma_W^2
 \end{bmatrix} \right)$
- Again, equality constraints on slopes and (co)variances unnecessary if *i* and *j* have roles

Evaluation in MCMC and ML

- I implemented the methods described above, mimicking a real-data analysis of sorority data
 - Described in APS slides, which can be downloaded from my Open Science Framework account, along with sorority data and syntax for xxM and RStan
 - https://osf.io/fmhg6/
- To evaluate the frequency properties of the method, I used parameter estimates as population values to simulate 200 samples of the same size, and imposed the same missing-data pattern (MCAR)
 - Also compared Bayes to (FI)ML using xxM package, specifying the SRM as a multilevel SEM



Full-Information ML (in xxM)

Parameter	True θ	$\widehat{\boldsymbol{\theta}}$	Bias	RMSE	Coverage	Power / a
μ (intercept)	2.00	1.99	-0.01	0.07	95%	100%
Perceiver σ^2	0.25	0.24	-0.01	0.05	95%	100%
β_1	0.30	0.30	0.00	0.07	95%	98%
β_2	0	0.00	0.00	0.07	94%	6%
Target o ²	0.12	0.16	0.04	0.05	75%	100%
β_1	0	0.00	0.00	0.06	96%	4%
β_2	0.30	0.30	0.00	0.06	98%	100%
Generalized p	-0.02	-0.01	0.01	0.04	95%	7%
Residual σ ²	0.49	0.49	0.00	0.03	95%	100%
Dyadic ρ	0.15	0.00	-0.15	0.15	3%	5%

Impute/Augment Data (in RStan)

Parameter	True θ	$\widehat{\boldsymbol{\theta}}$	Bias	RMSE	Coverage	Power / α
μ (intercept)	2.00	1.93	-0.07	0.07	0%	100%
Perceiver σ^2	0.25	0.27	0.02	0.06	94%	100%
β_1	0.30	0.29	-0.01	0.07	95%	95%
β_2	0	0.00	0.00	0.06	94%	7%
Target o ²	0.12	0.13	0.01	0.03	94%	100%
β_1	0	0.00	0.00	0.07	96%	4%
β_2	0.30	0.29	-0.01	0.05	97%	100%
Generalized p	-0.02	-0.02	0.00	0.03	96%	6%
Residual σ ²	0.49	0.50	0.01	0.03	95%	100%
Dyadic ρ	0.15	0.14	-0.01	0.05	96%	75%

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