

# Bayesian multilevel factor analysis of the Emotion Regulation Checklist

## EXAMINING CROSS-LEVEL AND LONGITUDINAL MEASUREMENT INVARIANCE

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### Discussion

The ERC was developed with single-level samples and single-level analyses. Our work is the first to examine its psychometric properties in a multilevel sample with multilevel analysis. We constructed two factors (regulatory skills and temperament) that demonstrate good reliability and metric invariance.

Single level analyses of parents' ERC ratings of the children yielded a similar factor structure with longitudinal metric invariance for regulatory skills and temperament (results omitted for brevity). This increases the utility of the measures we created.

Software limits prevented directly testing longitudinal scalar invariance. Equivalence tests (Mascha & Sessler, 2011) on thresholds via inequality hypotheses (van de Schoot et al., 2012) may be a viable alternative; we will investigate this option soon.

Cross-level scalar invariance implies perfect between-classroom reliability ( $\omega = 1$ ) because residual variances must be zero (Geldhof et al., 2014; Jak et al., 2014). So, our high reliabilities ( $\omega \geq .78$ ) suggest that the cluster bias caused by threshold non-invariance across classrooms is small enough to proceed with examining intervention effects.

### Research Design and Methods

Multi-site, cluster randomized longitudinal design with two 2-year cohorts.

- 8 sites (Head Start programs serving racially/ethnically diverse populations).
- 66 classrooms randomly assigned to intervention vs. control. Ten classrooms dropped out before child data collection, leaving 56 for this analysis.
- 522 children were rated by their teachers. Each teacher rated multiple children (Mean  $N = 9.3$ ).
- 2 sets of ratings per child (Time 1 in Fall, Time 2 in Spring).

Both teachers and parents rated children on the 24-item Emotion Regulation Checklist (ERC; Shields & Cicchetti, 1997, 1998; Miller et al., 2004).

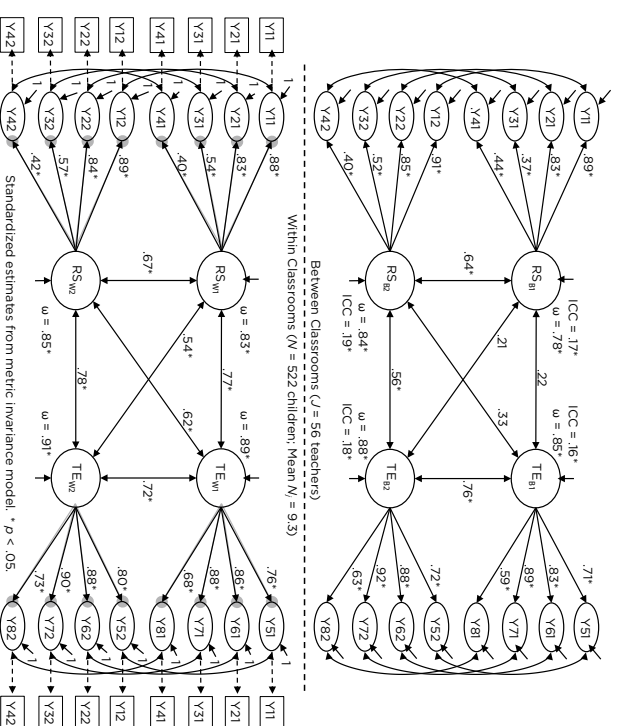
- 4-point Likert response format (treated as ordinal data).
- We could not replicate the published factor structure, so we used item content to select 8 core items for further analysis.

Multilevel EFA with half the Time 1 data suggested two 4-item factors: regulatory skills and temperament. We report the final multilevel CFA on the full dataset of teachers' ratings here.

- Used Bayesian estimation to better handle small classroom sample size (Hox, Van de Schoot, & Matuhjisse, 2012; Muthén & Asparouhov, 2012) and ordinal indicators (Liang & Yang, 2014).
- Estimated ICC for latent factors and items (Jak, Oort, & Dolan, 2014).
- Estimated level-specific reliability ( $\omega$ ) of factors (Geldhof, Preacher, & Zychner, 2014; Raykov & Marcoulides, 2011).

We tested measurement invariance longitudinally (Coertjens et al., 2012) and across levels (Jak et al., 2014) simultaneously.

- Longitudinal residual correlations among parallel items are method effects.
- Tested for exact configural invariance first (loadings free across time & level).
- Tested for exact metric invariance (equal loadings across time & level).
- Bayesian estimator doesn't support threshold constraints in multilevel models, preventing tests of exact longitudinal scalar invariance.



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Parameter Type	Priors	Constraints
Factor loadings	N(0,1)*	Free for within time 1; within time 2; between = within.
Latent & res. variances	W(1,3)**	Within res. variances = 1 (theta parameterization).
Latents' covariances	W(0,3)**	Free.
Latent means	N/A	Fixed = 0 for identification.
Res. covariances	W(0,3)**	Free longitudinally; else fixed = 0.
Thresholds	N(0.5)**	Free (constraints not supported).

\*Informative; lowers bias (Asparouhov & Muthén, 2010). \*\*Non-informative (Liang & Yang, 2014).

### Findings

Model fit was adequate, but far from ideal. Adding approximate zero cross-loadings with informative priors may improve model fit (Muthén & Asparouhov, 2012).

- Our model of regulatory skills and temperament measures demonstrates:
  - Substantial non-independence due to classroom-level variability (ICCs  $\geq .15$  for all individual items and for the latent factors).
  - Simultaneous longitudinal and cross-level metric invariance with teachers' ratings of children for both factors.
  - High composite reliability across time and level for both factors.
- Bayes factors testing whether between-classroom residual variances  $< 0.01$  all suggest thresholds are not invariant across classrooms

Model	Sampler	Chains	BITER	Thin	Convergence	# Param.	95% CI	$\chi^2$	$\chi^2_r$	PPP
Configural MI	Gibbs PX1	4	(10,000)	100	0.01	132	[-14.58, 142.10]	0.056		0.056
Metric MI	Gibbs PX1	4	(10,000)	100	0.01	108	[-12.16, 141.96]	0.051		0.051

Factor/Item	Time 1		Time 2		TI<->T2
	ICC	R <sup>2</sup> <sub>w</sub> R <sup>2</sup> <sub>b</sub>	ICC	R <sup>2</sup> <sub>w</sub> R <sup>2</sup> <sub>b</sub>	
<b>Regulatory skills (RS)</b>					
1. Transitions well from one activity to another	.17*	.78 .80	.18*	.80 .83	.33* .21
2. Can recover quickly from upset or distress	.18*	.68 .69	.19*	.71 .73	.29* .34
3. Is able to delay gratification	.31*	.30 .14	.22*	.33 .27	.22* .76*
4. Can say when s/he is feeling sad, angry or mad, fearful or afraid	.15*	.16 .19	.20*	.18 .16	.33* .32
<b>Temperament (TE)</b>					
5. [Does not] exhibit wide mood swings	.18*	.57 .51	.21*	.64 .52	.16 .32
6. [Is not] easily frustrated	.17*	.73 .68	.18*	.78 .77	.00 .24
7. [Is not] prone to angry outbursts/ tantrums easily	.16*	.77 .78	.17*	.82 .85	.30 .02
8. [Is not] impulsive	.20*	.46 .34	.22*	.53 .39	.47* .60*

Note: Likert format: 1. Rarely/never; 2. Sometimes; 3. Often; 4. Almost always. \* $p < .05$ .



### CHILD CHARACTERISTICS (N = 522)

Variable	n (%)
Female	262 (50.2%)
Hispanic	125 (24%)
Race	
AA	74 (16%)
AI/NA	23 (5%)
White	247 (52%)
Other	127 (27%)
Age (years)	3.7 (SD = .48)

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