

# Analyzing Time Series with Long-Range Dependencies 2: Using Exponential Smoothing to Model Complex Periodic Patterns



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## Abstract

Traditional seasonal time series models estimate short-range regularities (e. g., days of the week), but are not equipped to address long-range dependencies, nested seasonal cycles, or modeling periods including fractions, such as the 365.25 days per year in the Gregorian calendar (De Livera, Hyndman, & Snyder, 2012). The *forecast* package (ibid.) in R makes such analyses possible. This poster shows how with two datasets: 1. Daily high school attendance rates in one New York City High School (2009-2014), modeling autocorrelations over a school year of 177 days, and 2. Daily recordings of births to teens in the state of Texas (1964-1999), requiring a model that estimates weekly as well as annual dependencies. R script for *forecast* is provided.

## Rationale for the Study

Until recently, the estimation of long-range regularities in time series was cumbersome at best. This poster illustrates:

- how the Trigonometric Box-Cox ARMA Trend Seasonal (TBATS) model addresses this problem, and
- how the *forecast* package in R implements this model to analyze long-range dependencies statistically.

## Datasets

- Daily high school attendance rates in one New York City high school (School 2) from 2009 to 2014;
- Daily recordings of births to teens in the state of Texas from 1964 to 1999 (Hamilton *et al.*, 1997).

## Plan of the Analysis

- Initial exploration of the data, including stationarity tests;
- Outlier removal (attendance data only);
- Estimating short-range processes and long-range irregularity;
- Conventional ARIMA estimates with  $d=1$  and weekly; seasonal estimates (teen birth data only);
- Estimation of long-range regularity with TBATS;
- Analysis of residuals.

## The TBATS Model

The Trigonometric Box-Cox ARMA Trend Seasonal (TBATS) model can be expressed as follows:

$$Y_t^{(\omega)} = (Y_t^\omega - 1)/\omega \quad (1a)$$

$$Y_t^{(\omega)} = \ell_{t-1} + \phi b_{t-1} + \sum_{i=1}^T s_{t-1}^{(i)} + d_t \quad (1b)$$

$$\ell_t = \ell_{t-1} + \phi b_{t-1} + \alpha \delta_t \quad (1c)$$

$$b_t = (1 - \phi)b + \phi b_{t-1} + \beta d_t \quad (1d)$$

$$s_t^{(i)} = \sum_{j=1}^{k_i} s_{j,t}^{(i)} \quad (1e)$$

$$s_{j,t}^{(i)} = s_{j,t-1}^{(i)} \cos \lambda_j^{(i)} + s_{j,t-1}^{*(i)} \sin \lambda_j^{(i)} + \gamma_1^{(i)} d_t \quad (1f)$$

$$s_{j,t}^{*(i)} = -s_{j,t-1}^{(i)} \sin \lambda_j^{(i)} + s_{j,t-1}^{*(i)} \cos \lambda_j^{(i)} + \gamma_2^{(i)} d_t \quad (1g)$$

$$d_t = \sum_{i=1}^p \varphi_i d_{t-i} + \sum_{i=1}^q \theta_i \varepsilon_{t-i} + \varepsilon_t \quad (1h)$$

The Box-Cox transformation (1a) stabilizes the variance by  $\omega$ ;  $\ell_t$  estimates the local level at  $t$ ,  $b_t$  the short-range trend at  $t$ , and  $b$  the long-range trend across the series;  $d_t$  represents the ARMA  $(p, q)$  process;  $s_{j,t}^{(i)}$  models the seasonal component as a Fourier series with  $\lambda_j^{(i)} = 2\pi j/m_i$  with  $m_i$  representing the seasonal period;  $s_{j,t}^{*(i)}$  captures the level variance at the  $i^{th}$  seasonal cycle, and  $s_{j,t}^{(i)}$  models the change in seasonal variability over time;  $\alpha, \beta, \gamma_1^{(i)}$  and  $\gamma_2^{(i)}$  are smoothing parameters, and  $\varepsilon_t$  is Gaussian white noise with zero mean and constant variance  $\sigma^2$  (De Livera, Hyndman, & Snyder, 2012).

## Results

**Table 1.**  
Summary Statistics and Stationarity Tests:  
Daily Attendance in School 2 (N = 885)

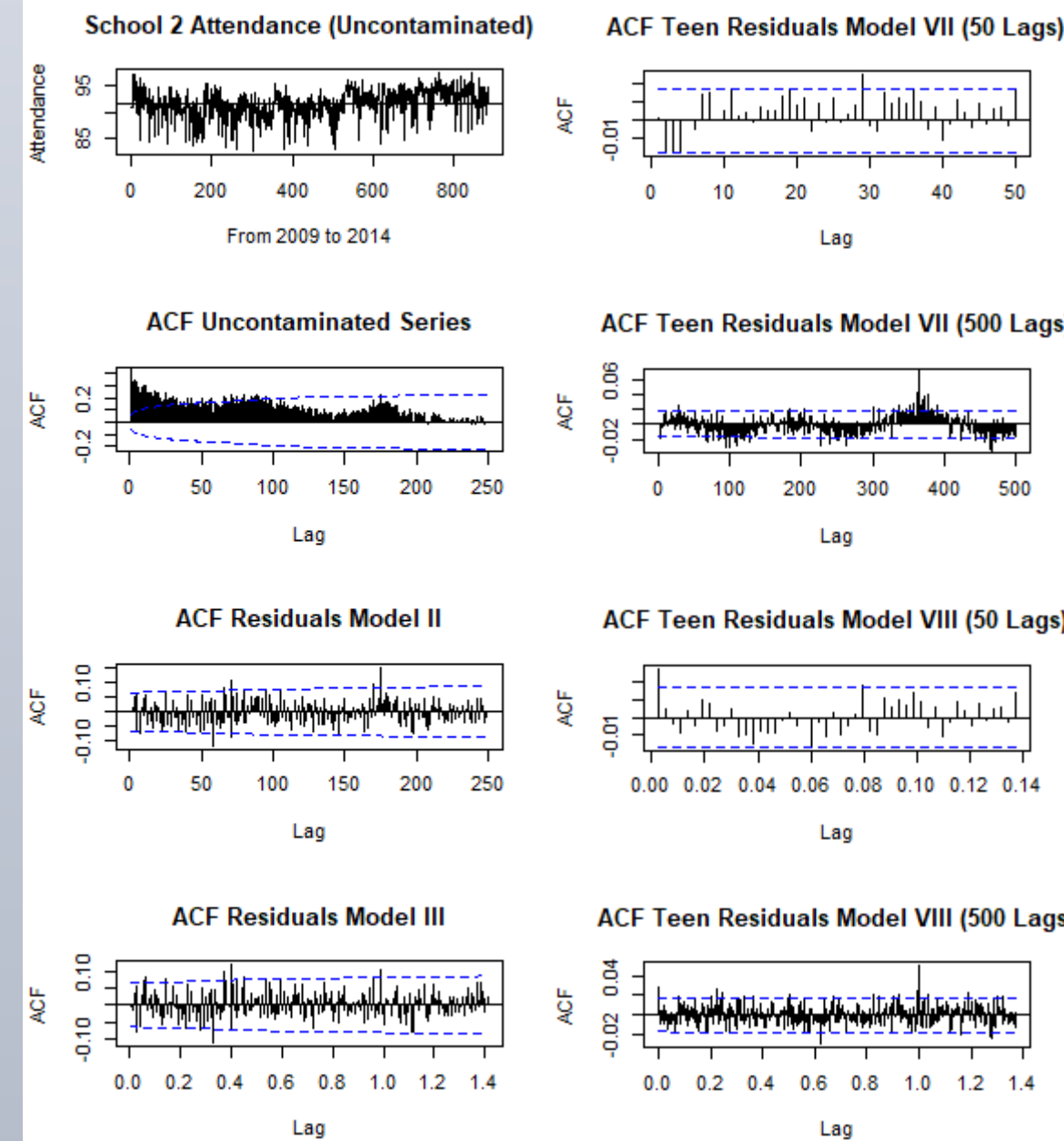
Summary Statistics	Uncontaminated Series
Mean	91.65
Standard Deviation	2.63
Minimum	82.60
First Quartile	90.28
Median	91.84
Third Quartile	93.56
Maximum	97.47
<b>Stationarity Tests</b>	
Augmented Dickey Fuller Test	-6.07*
Fuller Test	Lag Order = 9
KPSS Test	
Level	4.32*
Trend	0.66*
	Lag Order = 6

\* p < .01. A rejection of the null hypothesis implies stationarity in all three tests.

**Table 2.**  
Model Fitting with TBATS:  
Attendance in School 2 (Uncontaminated)

Model	ARMA (p, q)	Period	$\sigma^2$	AIC	LB Test
I	--	--	5.14	7460.77	34.45
II	(0, 1)	--	5.04	7446.82	21.55
III	(0, 1)	177	4.12	7383.55	25.82

P < .05  
Ljung-Box (LB) Portmanteau tested under a  $\chi^2$  distribution at df = 12  
**Preferred Model in Boldface**



**Table 3.**  
Summary Statistics and Stationarity Tests:  
Texas Teen Births Data (N = 13,149)

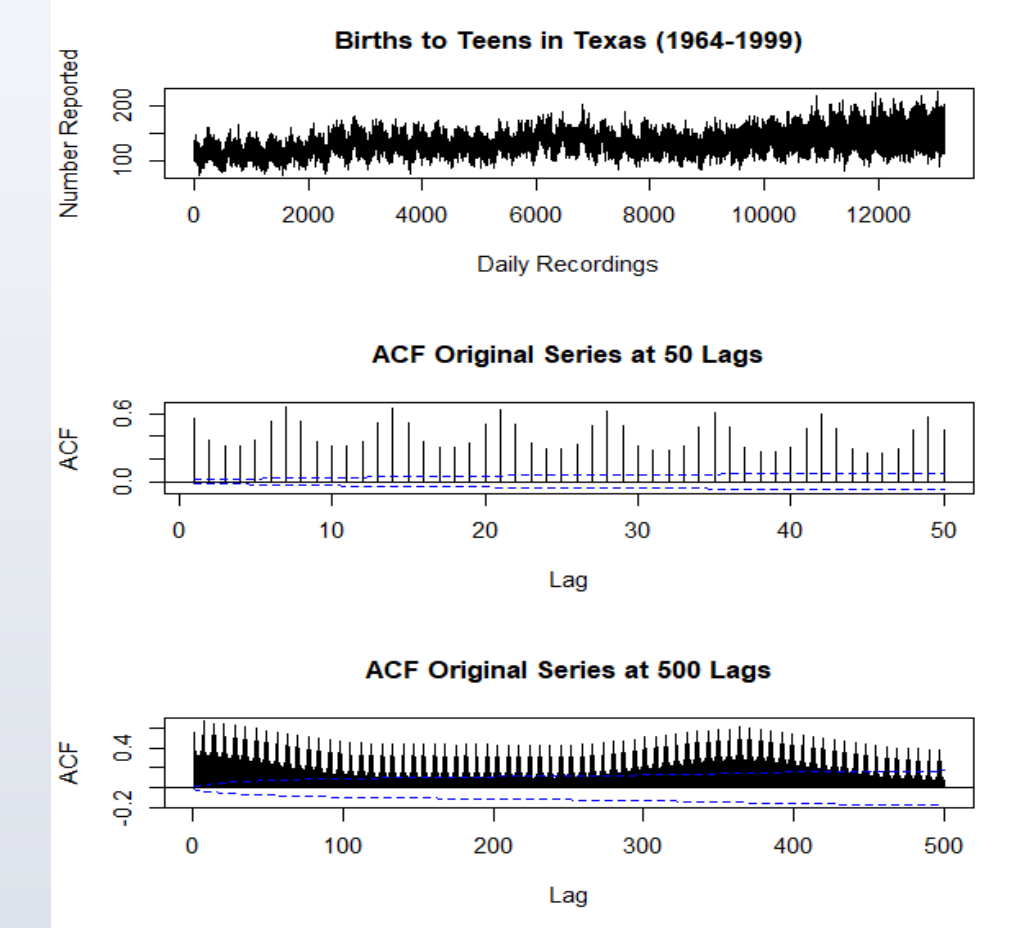
Summary Statistics	
Mean	132.2
Standard Deviation	21.04
Minimum	73.0
First Quartile	117.0
Median	131.0
Third Quartile	145.0
Maximum	226.0
<b>Stationarity Tests</b>	
Augmented Dickey Fuller Test	-8.25*
	Lag Order = 23
KPSS Test	
Level	24.23*
Trend	1.08*
	Lag Order = 26

\* p < .01. A rejection of the null hypothesis implies stationarity in all three tests.

**Table 4.**  
Model Fitting with Fractional Differencing and TBATS:  
Texas Teen Births Data

Model	Specification	$\sigma^2$	LB
I	(0, 0, 0)	442.50	29,486.00*
II	(1, 0, 1)	264.40	6,719.10*
III	(0, d, 0)	269.96	4,989.70*
IV	(1, d, 1)	266.53	3,690.90*
V	(1, 0, 1) X (0, 0, 1) <sub>7</sub>	234.30	2,004.80*
VI	(1, 0, 1) X (0, 1, 1) <sub>7</sub>	161.80	42.56*
VII	(1, 1, 1) X (0, 1, 1) <sub>7</sub>	162.50	20.87*
VIII	Residuals Model VII Seasonal Period = 365.25	158.36	17.70

\* p < .05  
Ljung-Box (LB) Portmanteau tested under a  $\chi^2$  distribution at df = 12  
**Preferred Model in Boldface**



## References

- De Livera, A. M., Hyndman, R. J., & Snyder, R. D. (2012). Forecasting time series with complex seasonal patterns using exponential smoothing. *Journal of the American Statistical Association*, 106, 1513-1527.
- Hamilton, P., Pollock, J. E., Mitchell, D. A., Vincenzi, A. E., & West, B. J. (1997). The application of nonlinear dynamics to nursing research. *Nonlinear Dynamics, Psychology, and Life Sciences*, 1, 237-261.

## R Script for Modeling Daily Attendance (School 2)

```
>library(forecast) ##call the forecast package
>attach(school 2) ##call the dataset for school 2
>summery(rate) ##generate attendance summary statistics
>tsoutliers(rate,iterate = 2) ##identify outliers
>urate<-tsclean(rate,replace.missing = TRUE)
##replace outliers
>urate.m1<-tbats(urate,use.box.cox = T, use.trend = F,
use.damped.trend = F, use.arma.errors = F)
>urate.m1 ##fit model 1 and generate output
>checkresiduals(urate.m1)
>urate.m2<-tbats(urate, use.box.cox = T, use.trend = F,
use.damped.trend = F, use.arma.errors = T)
>urate.m2 ##fit model 2 and generate output
>checkresiduals(urate.m2)
>urate.msts<-msts(urate, seasonal.periods=177)
##adjustment for annual cycle
>urate.m3<-tbats(urate.msts, use.box.cox=T,
use.trend=F, use.damped.trend=F, use.arma.errors = T)
>urate.m3 ##fit model 3 and generate output
>checkresiduals(urate.m3)
```