

# Difference in differences: A methodological illustration

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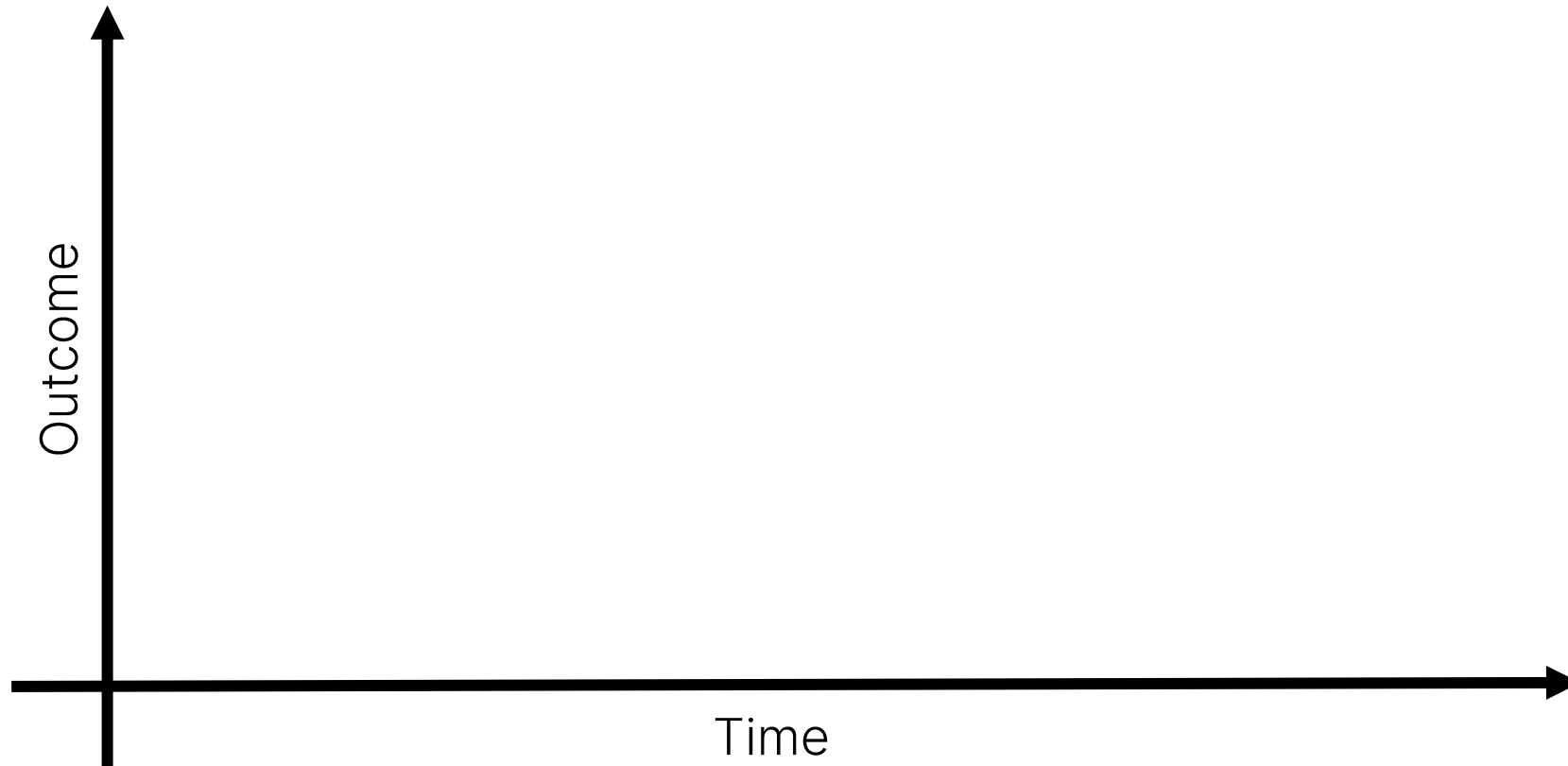
# Outline

- Conceptual introduction
- Motivating example
- Difference-in-differences analysis
  - Model formulation
  - Standard errors
  - Diagnostics
- Heterogeneous difference in differences
  - Model formulation
  - Aggregating results
  - Diagnostics

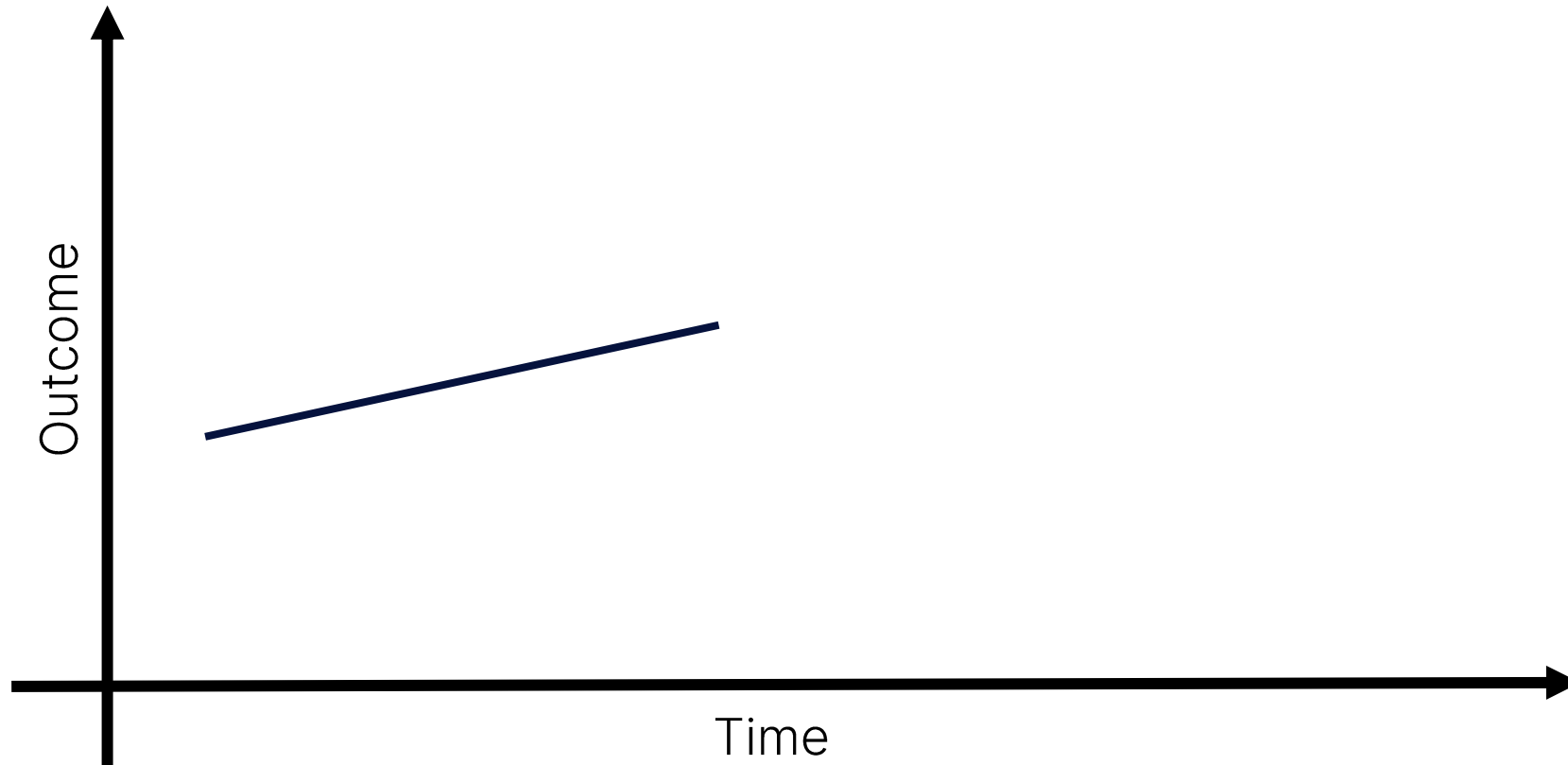
# What is difference in differences?

- Difference in differences (DID) is a causal inference method derived using the potential-outcomes framework.
- The intervention “treatment” variable can be binary or continuous.
- It estimates the average treatment effect on the treated (ATET) in a quasi-experimental pre/post design.
- ATETs can be homogenous or heterogeneous.
- DID can be applied to two types of observational data:
  - Repeated cross-sectional data
  - Panel data

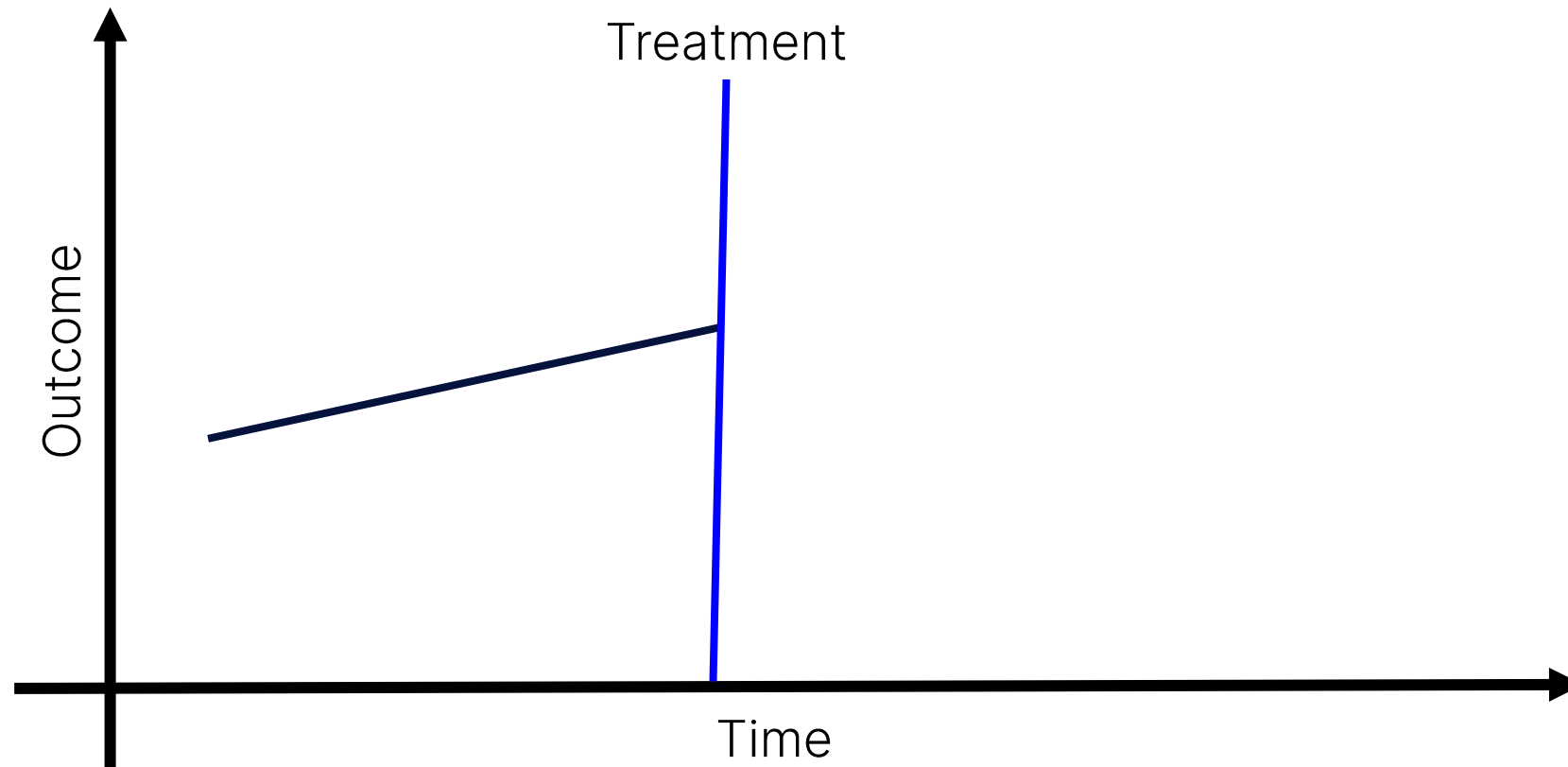
# Difference in differences (DID)



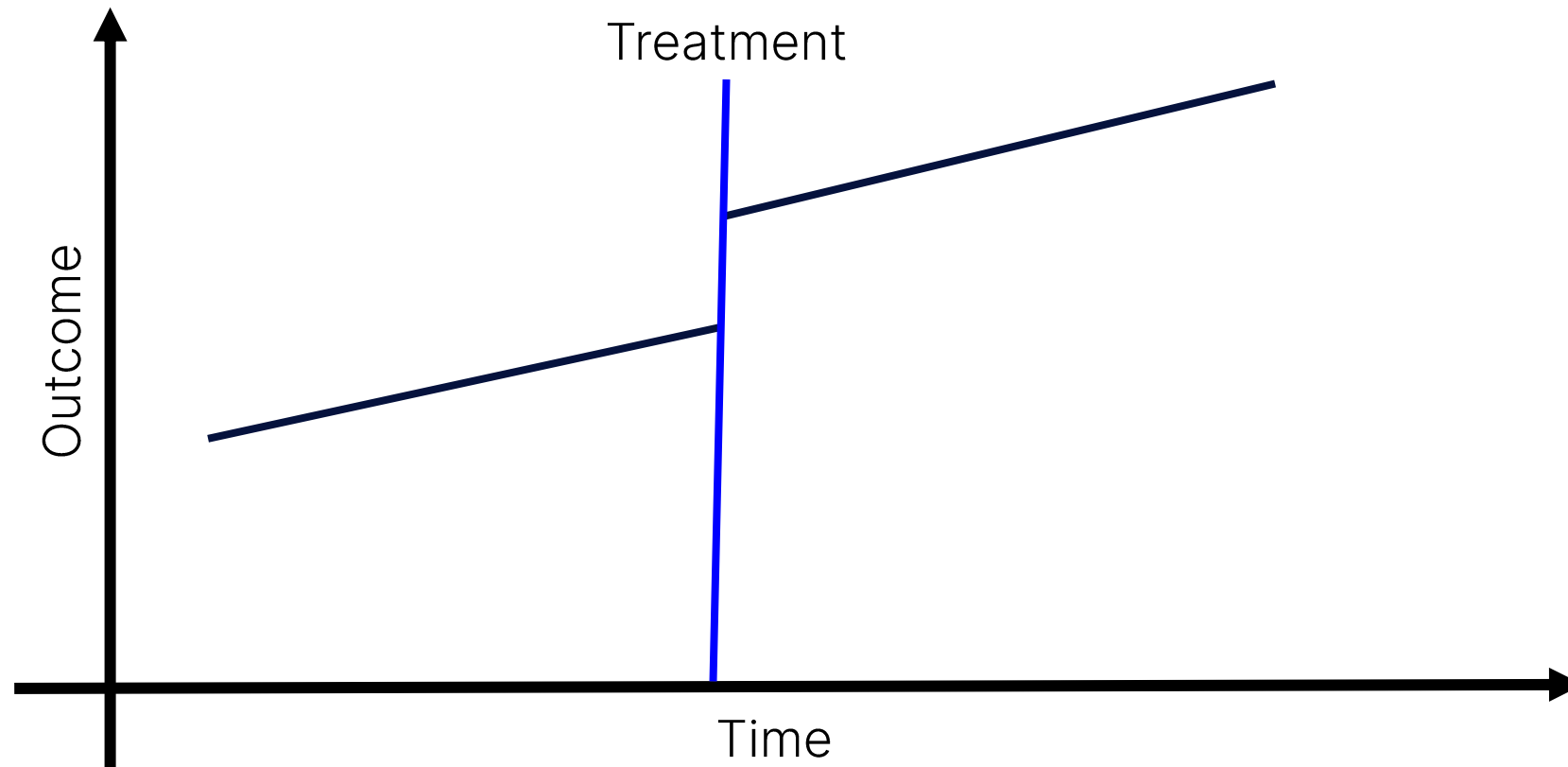
# Difference in differences (DID)



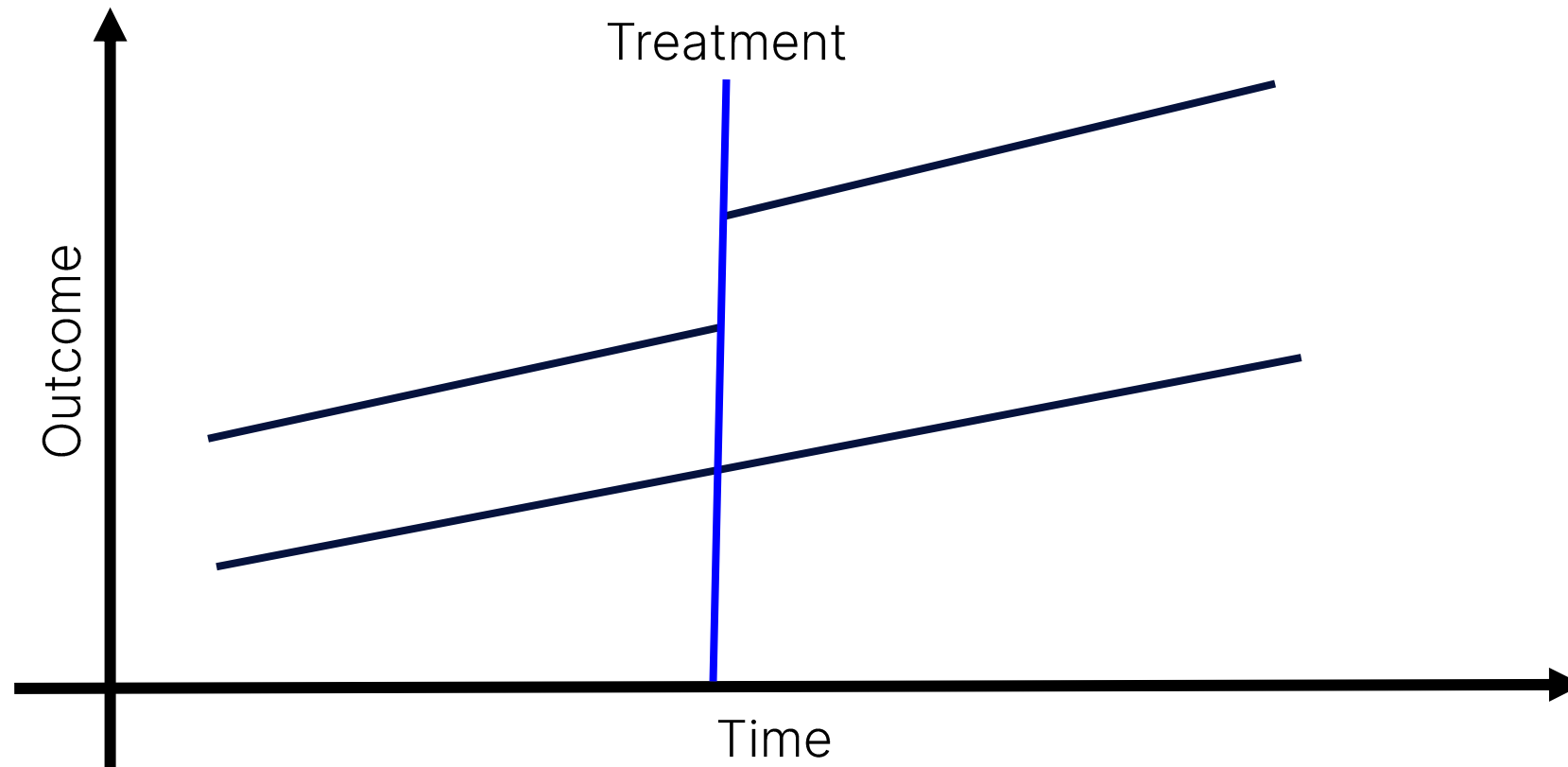
# Difference in differences (DID)



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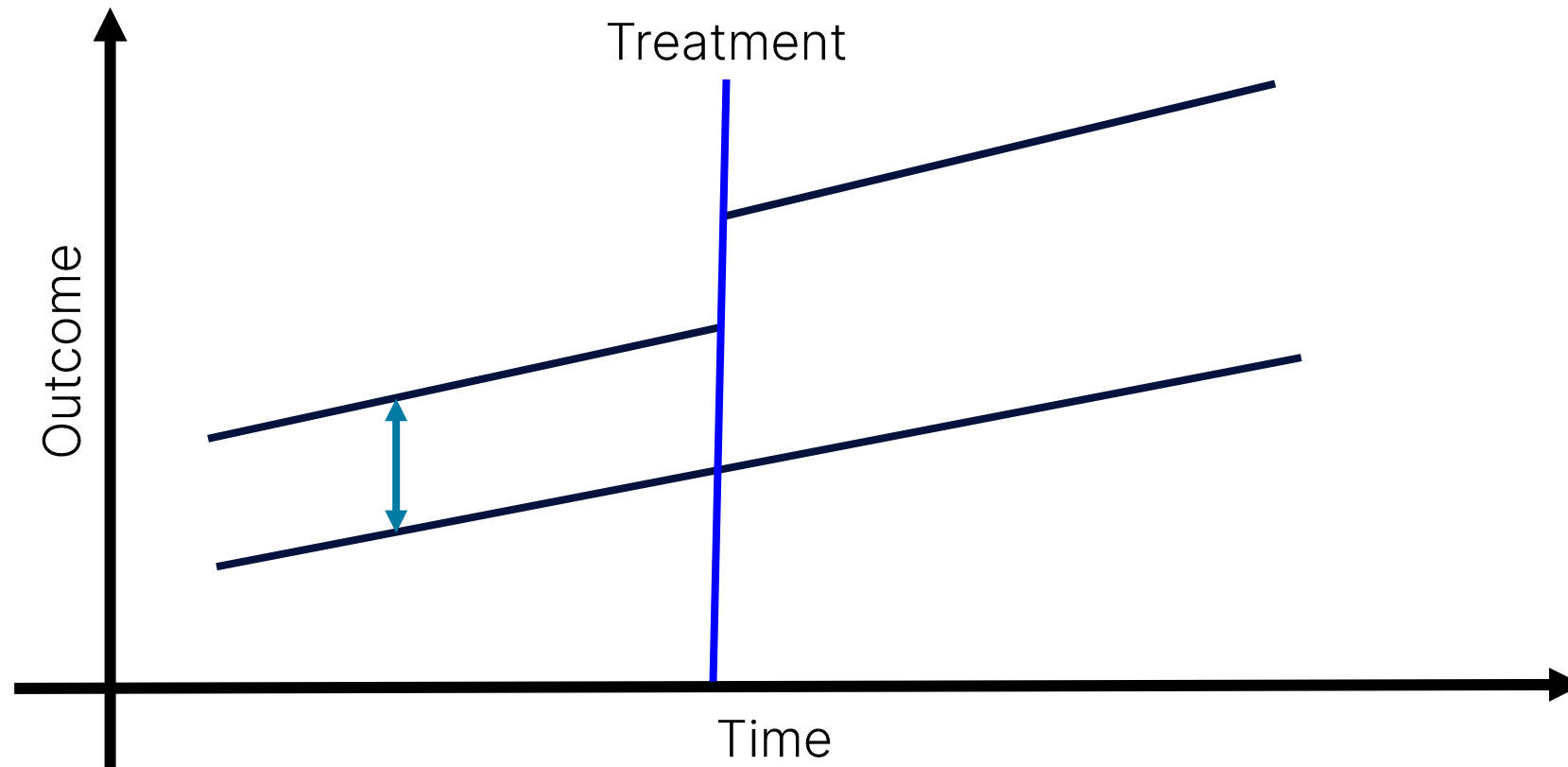


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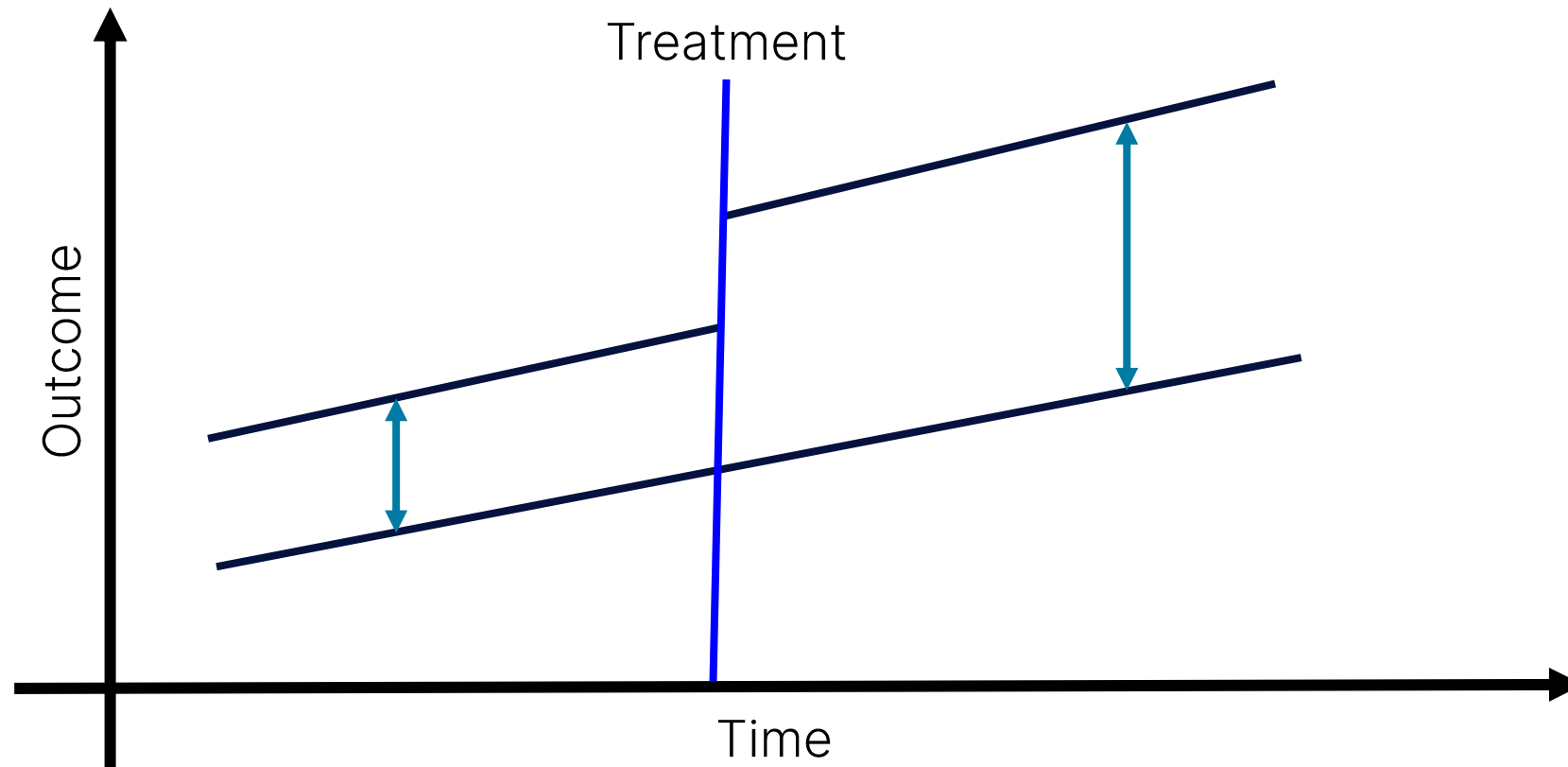




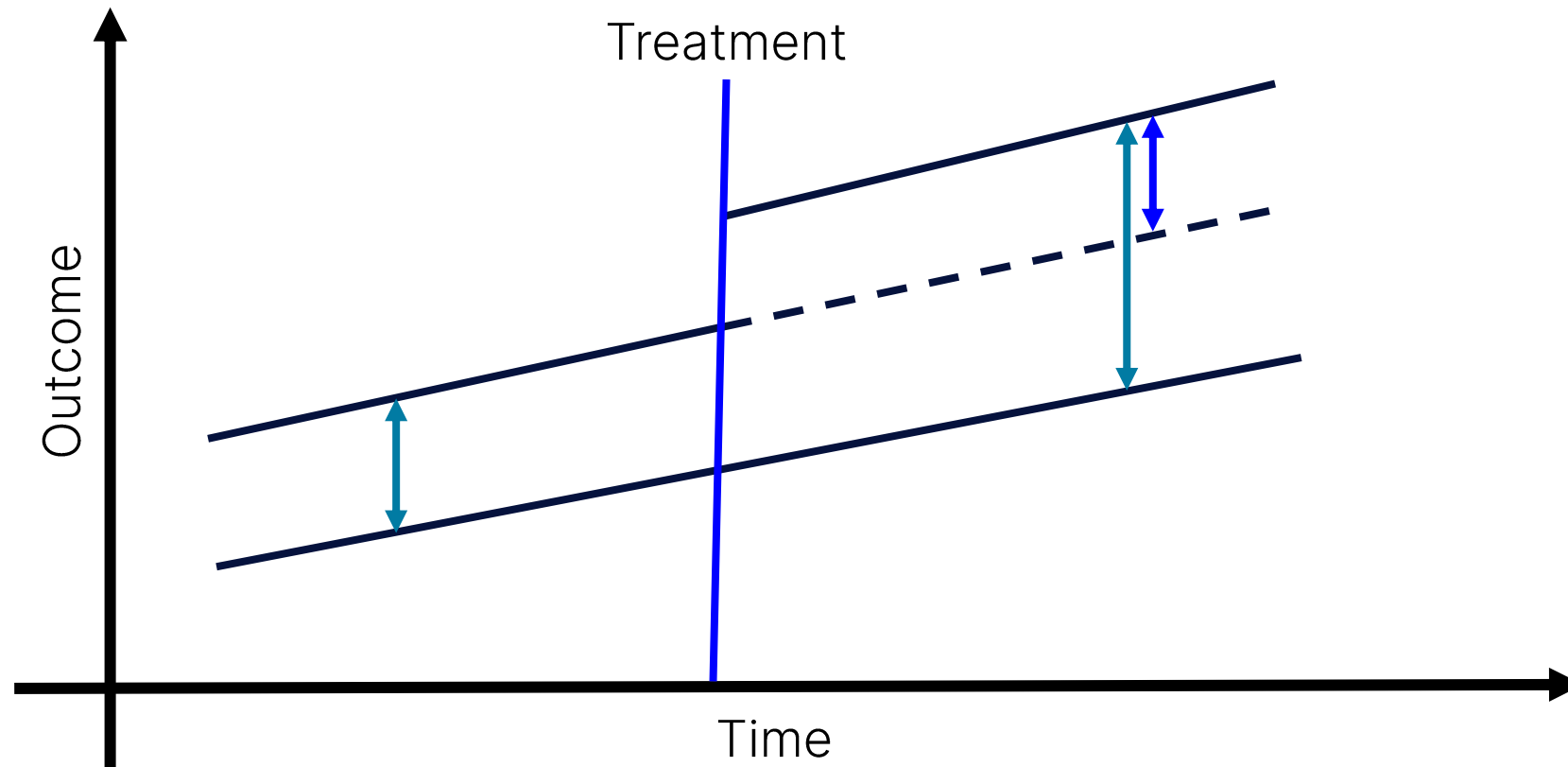
# Difference in differences (DID)



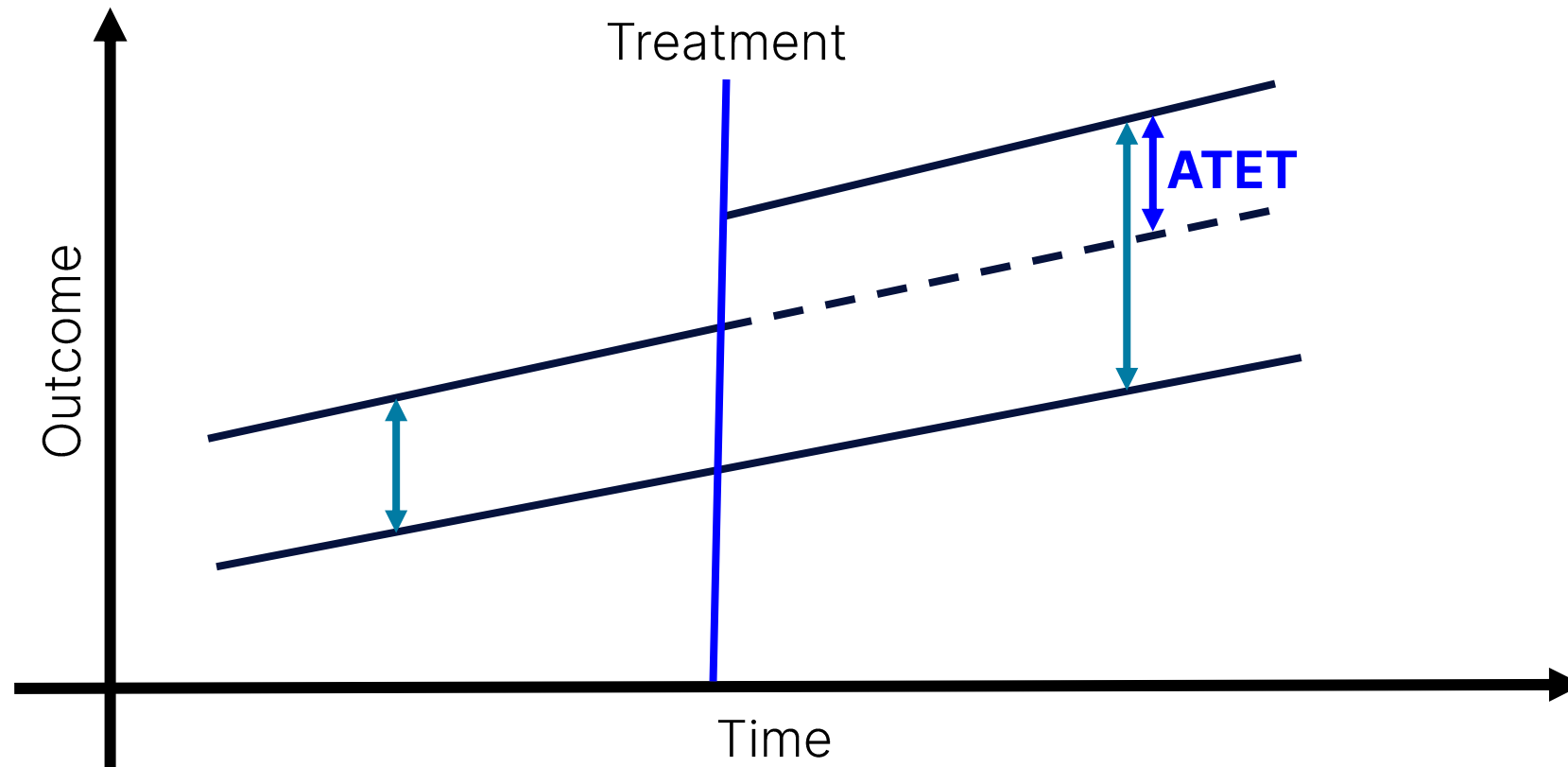
# Difference in differences (DID)



# Difference in differences (DID)



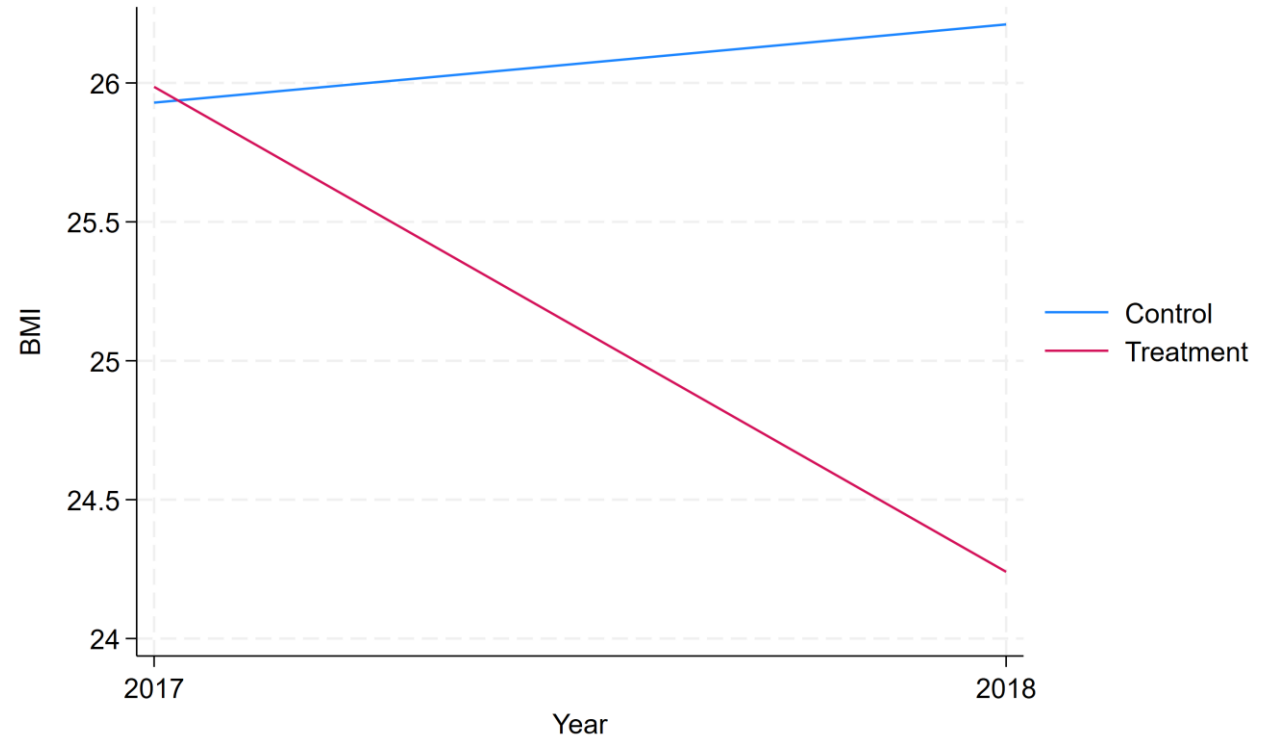
# Difference in differences (DID)



# Healthy Habits program

- Several school districts in the state have implemented a new program called Healthy Habits that incorporates more exercise time and augments the intake of fruits and vegetables at lunch time.
- Does the Healthy Habits program reduce students' body mass index (BMI) in 11 to 14 year olds?
- What is the average decrease in BMI for students who participate in the program?

# Study design – 2 x 2



# Repeated cross-sectional data

```
. use hhabit1  
(Fictional children BMI and school district data)
```

```
. list in 42/51
```

	district	year	hhabit	parksd	medu	girl	sports	bmi
42.	1	2017	No	6	8	Yes	Yes	26.92559
43.	1	2017	No	6	11	No	No	27.12172
44.	1	2017	No	6	8	Yes	Yes	24.77188
45.	1	2017	No	6	10	No	Yes	22.20161
46.	1	2017	No	6	9	No	Yes	18.3871
47.	1	2018	Yes	6	10	Yes	No	23.56391
48.	1	2018	Yes	6	13	No	No	28.9242
49.	1	2018	Yes	6	9	No	No	32.02917
50.	1	2018	Yes	6	0	No	Yes	21.86985
51.	1	2018	Yes	6	9	Yes	No	23.77382

# Repeated cross-sectional data

```
. bysort district: egen treatment = max(hhabit)

. dtable i.district bmi parksd medu i.girl i.sports if treatment, by(year)
  novarlabel sample( , statistics(frequency))
```

---

	2017	Year 2018	Total
N	149	145	294
district			
1	46 (30.9%)	44 (30.3%)	90 (30.6%)
2	52 (34.9%)	45 (31.0%)	97 (33.0%)
3	51 (34.2%)	56 (38.6%)	107 (36.4%)
bmi	25.986 (4.891)	24.240 (4.914)	25.125 (4.972)
parksd	5.658 (0.476)	5.614 (0.489)	5.636 (0.482)
medu	9.852 (2.820)	9.469 (2.438)	9.663 (2.641)
girl			
No	74 (49.7%)	78 (53.8%)	152 (51.7%)
Yes	75 (50.3%)	67 (46.2%)	142 (48.3%)
sports			
No	39 (26.2%)	49 (33.8%)	88 (29.9%)
Yes	110 (73.8%)	96 (66.2%)	206 (70.1%)

---



# Two-way fixed effects (TWFE)

$$y_{igt} = \gamma_g + \gamma_t + \mathbf{z}_{igt}\beta + D_{gt}\delta + \epsilon_{igt}$$

$i$  observation  
 $g$  group  
 $t$  time

$\gamma_g$  are the group fixed effects

$\gamma_t$  are the time fixed effects

$\mathbf{z}_{igt}$  are the covariates

$\beta$  are nuisance parameters

$D_{gt}$  is the treatment (varies by group and time)

$\delta$  is the average treatment effect on the treated (ATET)

# Two-way fixed effects (TWFE)

$$y_{igt} = \gamma_g + \gamma_t + \mathbf{z}_{igt}\beta + D_{gt}\delta + \epsilon_{igt}$$

$i$  observation  
 $g$  group  
 $t$  time

```
. regress bmi i.(district year girl sports hhabit), vce(cluster district)
```

# Two-way fixed effects (TWFE)

$$y_{igt} = \gamma_g + \gamma_t + \mathbf{z}_{igt}\beta + D_{gt}\delta + \epsilon_{igt}$$

$i$  observation  
 $g$  group  
 $t$  time

- . regress bmi i.(district year girl sports hhabit), vce(cluster district)
- . didregress (bmi i.girl i.sports) (hhabit), group(district) time(year)

# DID with regress

```
. regress bmi i.(district year girl sports hhabit), vce(cluster district)
```

```
Linear regression              Number of obs   =      2,815
                               F(3, 29)         =           .
                               Prob > F             =           .
                               R-squared            =      0.0446
                               Root MSE         =      4.8028
```

(Std. err. adjusted for 30 clusters in district)

bmi	Coefficient	Robust std. err.	t	P> t	[95% conf. interval]	
district						
1	0	(base)				
2	1.14331	.0084414	135.44	0.000	1.126046	1.160575
3	.2090071	.0234695	8.91	0.000	.1610065	.2570077
4	.1247919	.1658403	0.75	0.458	-.2143895	.4639733
5	.46288	.1566941	2.95	0.006	.1424046	.7833554
6	-.1512917	.1712865	-0.88	0.384	-.5016119	.1990286
7	.8683145	.1643517	5.28	0.000	.5321775	1.204451
8	.1420385	.1602934	0.89	0.383	-.1857983	.4698752
9	-.3473355	.1476729	-2.35	0.026	-.6493605	-.0453106
10	-.1224466	.1502409	-0.82	0.422	-.4297237	.1848305
11	-.0971328	.1411213	-0.69	0.497	-.3857582	.1914925
12	.2576859	.1569112	1.64	0.111	-.0632336	.5786054
13	1.21500	.1633451	8.10	0.000	.8820566	1.647122

18	-.5640486	.1529064	-3.69	0.001	-.8767772	-.25132
19	.4654356	.1526969	3.05	0.005	.1531354	.7777358
20	.5854752	.1537438	3.81	0.001	.2710338	.8999166
21	.2617872	.1628856	1.61	0.119	-.0713512	.5949257
22	.778589	.153641	5.07	0.000	.4643579	1.09282
23	1.334889	.1515312	8.81	0.000	1.024972	1.644805
24	.647442	.1552261	4.17	0.000	.329969	.9649149
25	.2337309	.1585146	1.47	0.151	-.0904679	.5579298
26	.0524395	.1578669	0.33	0.742	-.2704346	.3753136
27	-.0520582	.1613677	-0.32	0.749	-.3820921	.2779757
28	-.2484144	.1546527	-1.61	0.119	-.5647148	.0678859
29	.5323953	.1527937	3.48	0.002	.2198971	.8448934
30	.4268822	.1539278	2.77	0.010	.1120645	.7416998
year						
2017	0	(base)				
2018	.3300558	.2144101	1.54	0.135	-.1084622	.7685738
girl						
No	0	(base)				
Yes	.9947139	.1583773	6.28	0.000	.6707959	1.318632
sports						
No	0	(base)				
Yes	-1.377504	.2129341	-6.47	0.000	-1.813003	-.9420045
hhabit						
No	0	(base)				
Yes	-2.104804	.3071655	-6.85	0.000	-2.733028	-1.47658
_cons	26.03178	.1716093	151.69	0.000	25.6808	26.38276

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_cons	26.03178	.1716093	151.69	0.000	25.6808	26.38276

# DID with didregress

```
. didregress (bmi i.girl i.sports) (hhabit), group(district) time(year)
```

Difference-in-differences regression  
Data type: Repeated cross-sectional

Number of obs = 2,815

(Std. err. adjusted for 30 clusters in district)

		Robust				
	bmi	Coefficient	std. err.	t	P> t	[95% conf. interval]
ATET	hhabit					
(Yes vs No)		-2.104804	.3071655	-6.85	0.000	-2.733028 -1.47658

Note: ATET estimate adjusted for covariates, group effects, and time effects.

# Auxiliary equations

```
. didregress (bmi i.girl i.sports) (hhabit), group(district) time(year) aeq
```

Difference-in-differences regression  
Data type: Repeated cross-sectional

Number of obs = 2,815

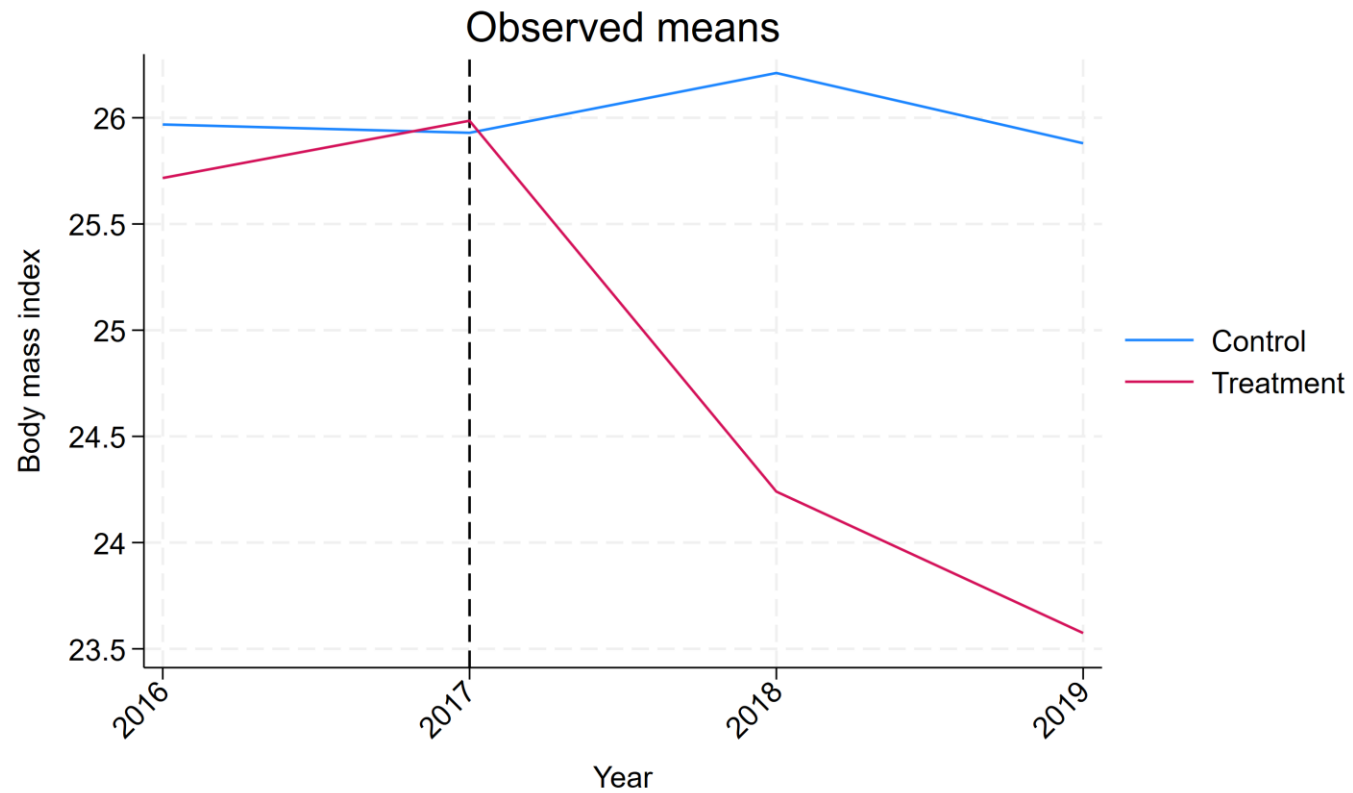
(Std. err. adjusted for 30 clusters in district)

bmi	Coefficient	Robust std. err.	t	P> t	[95% conf. interval]	
ATET						
hhabit (Yes vs No)	-2.104804	.3071655	-6.85	0.000	-2.733028	-1.47658
Controls						
girl						
No	0	(base)				
Yes	.9947139	.1583773	6.28	0.000	.6707959	1.318632
sports						
No	0	(base)				
Yes	-1.377504	.2129341	-6.47	0.000	-1.813003	-.9420045
year						
2017	0	(base)				
2018	.3300558	.2144101	1.54	0.135	-.1084622	.7685738



# Add additional time points

```
. use hhabit2
```



# Add additional time points

```
. didregress (bmi i.girl i.sports) (hhabit), group(district) time(year)
```

Difference-in-differences regression

Number of obs = 5,649

Data type: Repeated cross-sectional

(Std. err. adjusted for 30 clusters in district)

		Robust				
	bmi	Coefficient	std. err.	t	P> t	[95% conf. interval]
ATET						
hhabit						
(Yes vs No)		-2.00429	.2231442	-8.98	0.000	-2.460671 -1.547909

Note: ATET estimate adjusted for covariates, group effects, and time effects.

# Standard error considerations

- Treatment is administered at the group level
- Default standard errors are cluster-robust SEs at the group level (Bertrand, Dufflo, Mullainathan, 2004)
- When too few groups (<40-50):
  - Bias-corrected standard errors (HC2) with the Bell and McCaffrey (2002) degrees-of-freedom adjustment
  - Aggregation method
  - Wild-cluster bootstrap to obtain p-values and confidence intervals

# Bias-corrected robust standard error

- Cluster generalization of HC2
  - Scale residuals inverse of square of diagonals from projection matrix
  - Higher weight to residuals of observations with high leverage
- Bell and McCaffrey (2002) suggest a degrees of freedom adjustment

# Bias-corrected robust standard error

```
. didregress (bmi i.girl i.sports) (hhabit), group(district) time(year) vce(hc2)
```

Difference-in-differences regression

Number of obs = 5,649

No. of clusters = 30

Data type: Repeated cross-sectional

		Robust HC2				[95% conf. interval]	
bmi		Coefficient	std. err.	t	P> t		
ATET	hhabit						
(Yes vs No)		-2.00429	.2512349	-7.98	0.008	-2.91349	-1.095089

Note: ATET estimate adjusted for covariates, group effects, and time effects.

# Aggregation method

- Regress out observation-level covariates
- Aggregate at the  $g,t$  level
- Fit DID regression model to the aggregated data

$$y_{igt} = z_{1igt}\beta_1 + C_{gt} + \epsilon_{igt}$$

$$\hat{C}_{gt} = z_{2gt}\beta_2 + D_{gt}\delta + v_{gt}$$

- Use option `aggregate(standard)` to use the standard method with constant  $\beta_1$
- Use option `aggregate(dlang, varying)` to use the method proposed by Donald and Lang (2007) to estimate  $\beta_1$  separately by group

# Aggregation method

```
. didregress (bmi i.girl i.sports) (hhabit), group(district) time(year) aggregate(dlang, varying)
```

Difference-in-differences regression Number of obs = 120

Data type: Repeated cross-sectional

Aggregation: Donald-Lang, varying coefficients

bmi	Coefficient	Std. err.	t	P> t	[95% conf. interval]	
ATET hhabit (Yes vs No)	-1.569087	.9738948	-1.61	0.110	-3.498185	.3600115

Note: ATET estimate adjusted for covariates, group effects, and time effects.

# Aggregation method

```
. didregress (bmi i.girl i.sports) (hhabit), group(district) time(year) aggregate(dlang)
```

Difference-in-differences regression  
Data type: Repeated cross-sectional  
Aggregation: Donald-Lang

Number of obs = 120

bmi	Coefficient	Std. err.	t	P> t	[95% conf. interval]	
ATET hhabit (Yes vs No)	-1.971313	.4335124	-4.55	0.000	-2.830018	-1.112609

Note: ATET estimate adjusted for covariates, group effects, and time effects.



# Wild cluster bootstrap

1. Fit unrestricted model to estimate t-statistic for ATET,  $t$
2. Fit restricted model (assume ATET=0) and predict outcome,  $\hat{y}_{igt}$ , and residuals,  $\hat{\epsilon}_{igt}$
3. Repeat the following steps B times
  1. Multiply residual by randomly sampled weight function at the group level, i.e. from normal distribution,  $\tilde{\epsilon}_{igt,b} = \hat{\epsilon}_{igt} \times w_{g,b}$
  2. Use weighted residual to update predicted outcome,  $\tilde{y}_{igt,b} = \hat{y}_{igt} + \tilde{\epsilon}_{igt}$
  3. Fit unrestricted model to  $\tilde{y}_{igt,b}$  to estimate t-statistic for ATET,  $\tilde{t}_b$
4. Compare the observed  $t$  to the B bootstrap sampled  $\tilde{t}_b$ 's to compute p-values and 95% confidence intervals

# Wild cluster bootstrap

```
. didregress (bmi i.girl i.sports) (hhabit), group(district) time(year) wildbootstrap(rseed(273))
```

```
Performing 1,000 replications for p-value for constraint  
hhabit = 0 ...
```

```
DID with wild-cluster bootstrap inference
```

```
Number of obs = 5,649
```

```
Replications = 1,000
```

```
Data type: Repeated cross-sectional
```

```
Error weight: rademacher
```

bmi	Coefficient	t	P> t	[95% conf. interval]	
ATET hhabit (Yes vs No)	-2.00429	-8.98	0.018	-3.297128	-1.101595

Note: ATET estimate adjusted for covariates, group effects, and time effects.

# Wild cluster bootstrap

- Change error weight with option `errorweight()`

Error weight	Formula
rademacher	$\{-1 \ 1\}$ with probability $1/2$
mammen	$\{1 - \phi\}$ with probability $\frac{\phi}{\sqrt{5}}$ , $\{\phi\}$ otherwise where $\phi = \frac{1+\sqrt{5}}{2}$
webb	$\left\{\frac{-\sqrt{3}}{2} \ \frac{-\sqrt{2}}{2} \ \frac{-\sqrt{1}}{2} \ \frac{\sqrt{1}}{2} \ \frac{\sqrt{2}}{2} \ \frac{\sqrt{3}}{2}\right\}$ with probability $1/6$
normal	$\sim N(0,1)$
gamma	$\sim \Gamma\left(4, \frac{1}{2}\right)$

---

# Wild cluster bootstrap

```
. didregress (bmi i.sports i.girl) (hhabit) if district<11, group(district) time(year) wilddbstrap(rseed(432))
```

Performing 1,000 replications for p-value for constraint

**hhabit = 0 ...**

Computing confidence interval for **hhabit**

Lower bound: .....10 done (10)

Upper bound: .....10.....20... done (23)

note: upper-bound CI achieved  $1-F(-1.39) = 0.0240$ , but target is  $1-F(x) = .025$ .

note: at least one bootstrap t statistic matches the t statistic under the null; this prevents the CI bound from converging to the target.

DID with wild-cluster bootstrap inference

Number of obs = 1,880

Replications = 1,000

Data type: Repeated cross-sectional

Error weight: rademacher

bmi	Coefficient	t	P> t	[95.10% conf. interval]	
ATET					
hhabit (Yes vs No)	-2.106906	-8.16	0.002	-2.80338	-1.386673

Note: 95.10% confidence interval is wider than requested.

Note: ATET estimate adjusted for covariates, group effects, and time effects.

# Wild cluster bootstrap

```
. didregress (bmi i.sports i.girl) (hhabit) if district<11, group(district) time(year) wildbootstrap(rseed(432)
errorweight(normal))
```

Performing 1,000 replications for p-value for constraint  
hhabit = 0 ...

DID with wild-cluster bootstrap inference

Number of obs = 1,880

Replications = 1,000

Data type: Repeated cross-sectional

Error weight: normal

bmi	Coefficient	t	P> t	[95% conf. interval]	
ATET hhabit (Yes vs No)	-2.106906	-8.16	0.000	-2.676421	-1.448629

Note: ATET estimate adjusted for covariates, group effects, and time effects.

# Assumptions

- Conditional-mean independence (CMI) assumption
  - there are no unmeasured confounders
- Overlap assumption
  - each treated individual has a positive probability of not receiving the treatment
- Stable unit treatment value assumption (SUTVA)
  - composition of intervention and comparison groups is stable for repeated cross-sectional design
- Parallel trends
  - treatment and control groups have parallel trends in outcome prior to treatment

# Assessing SUTVA

```
. dtable parksd medu i.girl i.sports if treatment, by(year, nototal test) novarlabel
note: using test regress across levels of year for parksd and medu.
note: using test pearson across levels of year for girl and sports.
```

	Year				Test
	2016	2017	2018	2019	
N	137 (24.7%)	149 (26.9%)	145 (26.2%)	123 (22.2%)	
parksd	5.686 (0.466)	5.658 (0.476)	5.614 (0.489)	5.691 (0.464)	0.509
medu	10.036 (2.922)	9.852 (2.820)	9.469 (2.438)	10.008 (2.757)	0.281
girl					
No	58 (42.3%)	74 (49.7%)	78 (53.8%)	67 (54.5%)	0.169
Yes	79 (57.7%)	75 (50.3%)	67 (46.2%)	56 (45.5%)	
sports					
No	41 (29.9%)	39 (26.2%)	49 (33.8%)	33 (26.8%)	0.474
Yes	96 (70.1%)	110 (73.8%)	96 (66.2%)	90 (73.2%)	

# Assessing SUTVA

```
. dtable parksd medu i.girl i.sports if !treatment, by(year, nototal test) novarlabel  
note: using test regress across levels of year for parksd and medu.  
note: using test pearson across levels of year for girl and sports.
```

---

	Year				
	2016	2017	2018	2019	Test
N	1,289 (25.3%)	1,213 (23.8%)	1,308 (25.7%)	1,285 (25.2%)	
parksd	4.648 (1.192)	4.613 (1.212)	4.622 (1.177)	4.591 (1.202)	0.689
medu	10.049 (2.714)	9.967 (2.738)	9.997 (2.871)	9.886 (2.694)	0.503
girl					
No	650 (50.4%)	587 (48.4%)	691 (52.8%)	652 (50.7%)	0.173
Yes	639 (49.6%)	626 (51.6%)	617 (47.2%)	633 (49.3%)	
sports					
No	428 (33.2%)	420 (34.6%)	438 (33.5%)	439 (34.2%)	0.873
Yes	861 (66.8%)	793 (65.4%)	870 (66.5%)	846 (65.8%)	

---



# Assessing parallel trends

- Parallel-trends assumption: Both groups had mean outcomes that evolved similarly to each other over time prior to treatment
  - `estat ptrends`
  - `estat trendplots`
- Anticipatory treatment assumption: Neither group changed their behavior in anticipation of the treatment
  - `estat granger`
  - `estat grangerplot`

# Linear trend model

$$y_{igt} = \gamma_g + \gamma_t + \mathbf{z}_{igt}\beta + D_{gt}\delta + T_g \text{pre}_t t \theta + \epsilon_{igt}$$

$i$  observation  
 $g$  group  
 $t$  time

- Need two new variables:

```
. bysort district: egen treatment = max(hhabit)  
. generate pre = year<2018
```

- And include their three-way interaction with continuous time variable

```
. regress bmi i.hhabit i.district i.year treatment#pre#c.year, vce(cluster district)
```

# Linear trend model

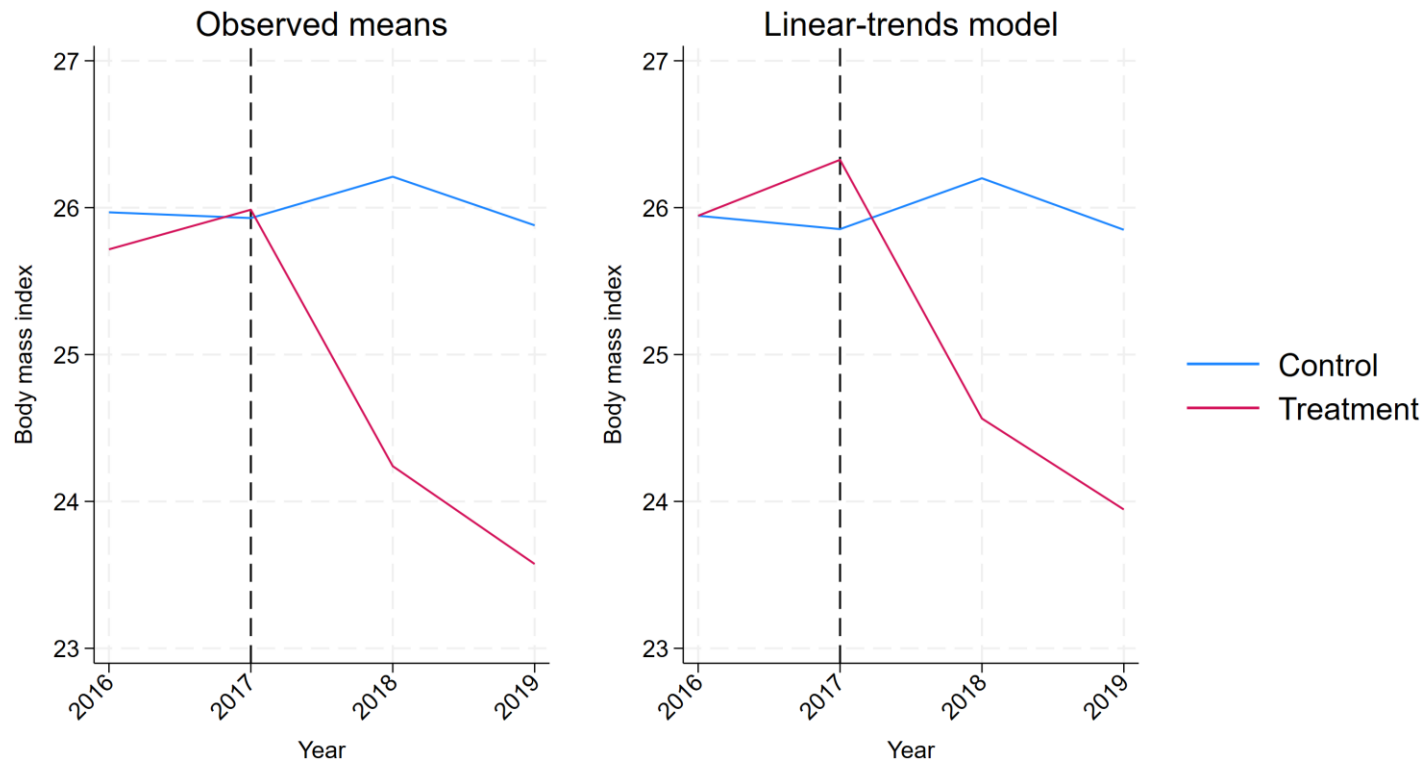
```
. estat trendplot
```

```
. estat ptrend
```

## Graphical diagnostics for parallel trends

Parallel-trends test (pretreatment time period)  
H0: Linear trends are parallel

F(1, 29) = 0.52  
Prob > F = 0.4749



# Granger causality model

$$y_{igt} = \gamma_g + \gamma_t + \mathbf{z}_{igt}\beta + D_{gt}\delta + T_g\text{pre}_t\gamma_t\theta + \epsilon_{igt}$$

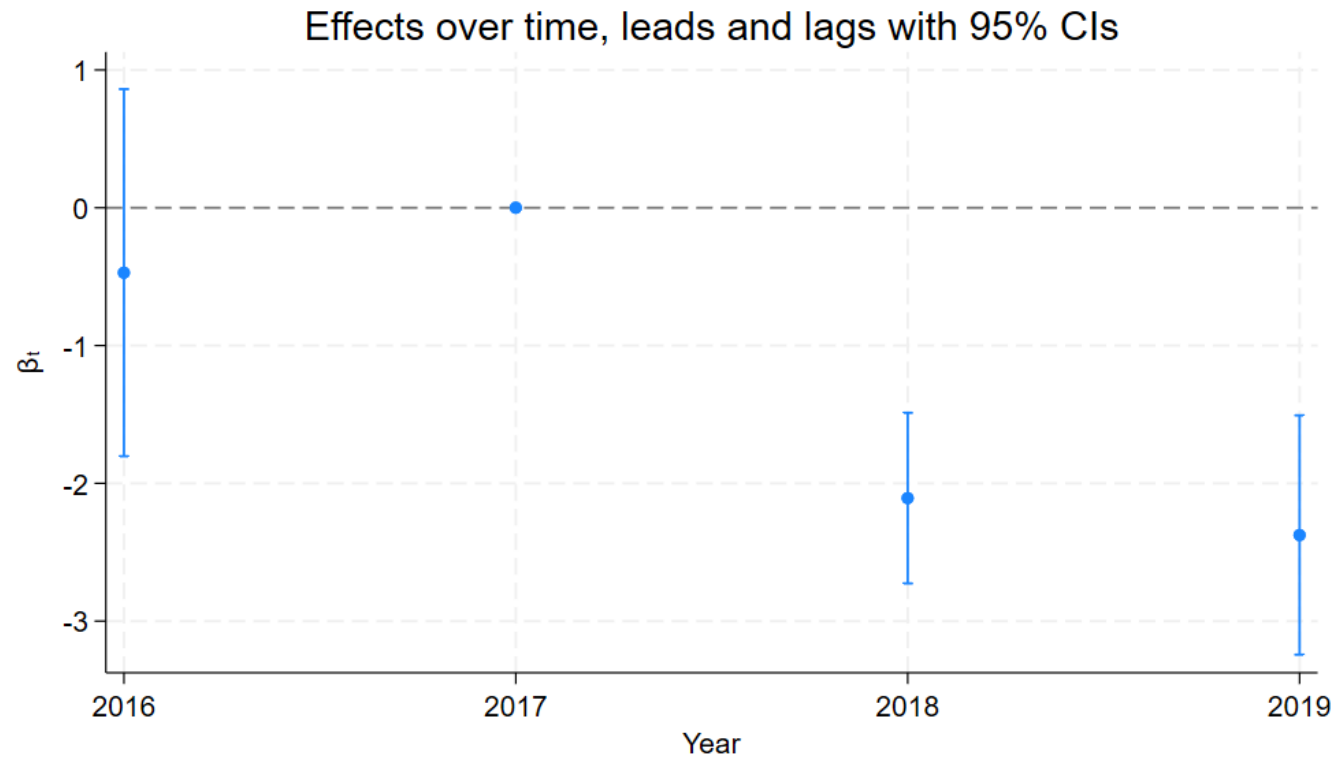
$i$  observation  
 $g$  group  
 $t$  time

- Include the three-way interaction with *discrete* time variable

```
. regress bmi i.hhabit i.district i.year treatment#pre#year, vce(cluster district)
```

# Granger causality model

```
. estat grangerplot
```



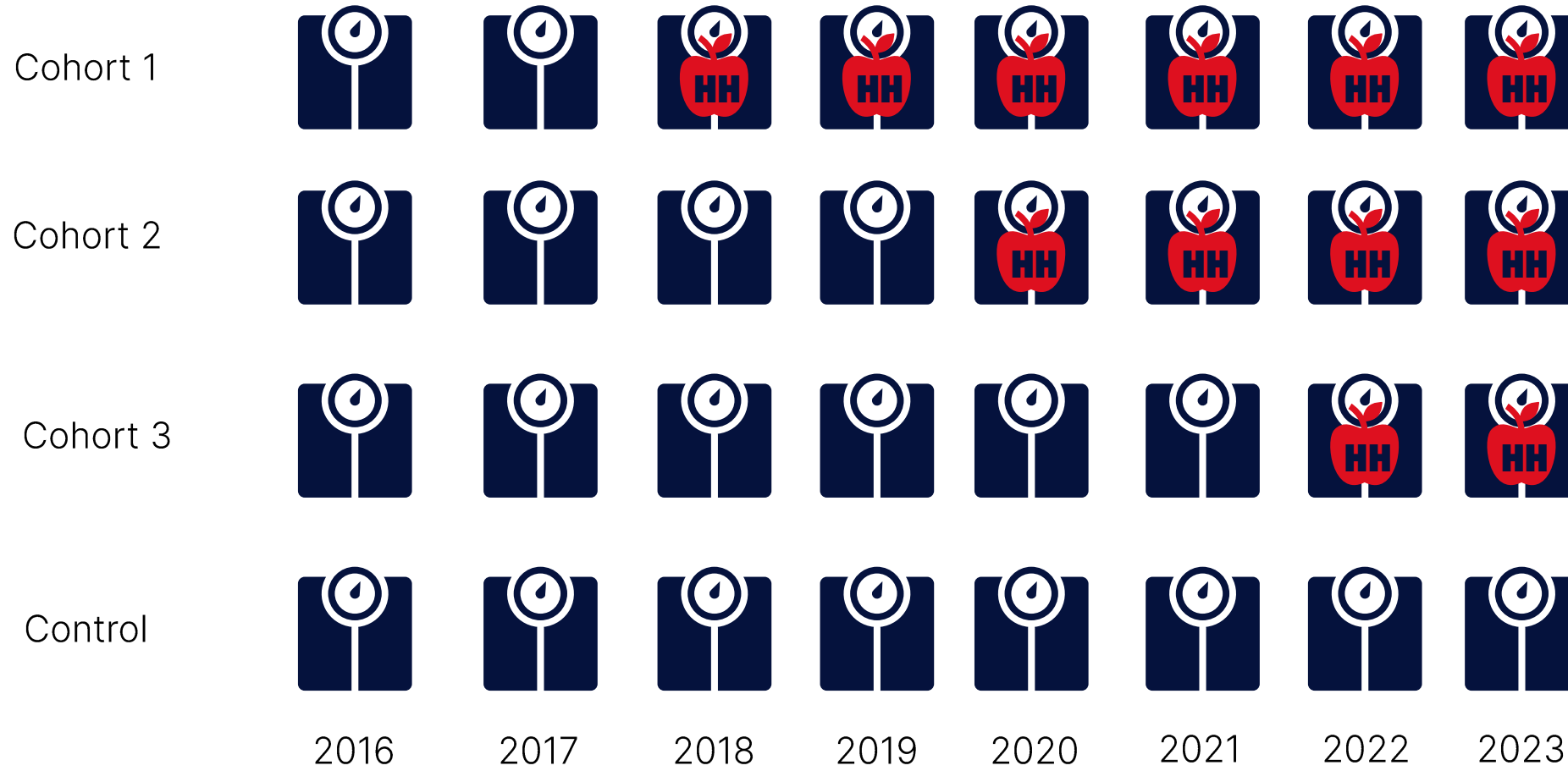
```
. estat granger
```

Granger causality test  
H0: No effect in anticipation of treatment

F(1, 29) = 0.52

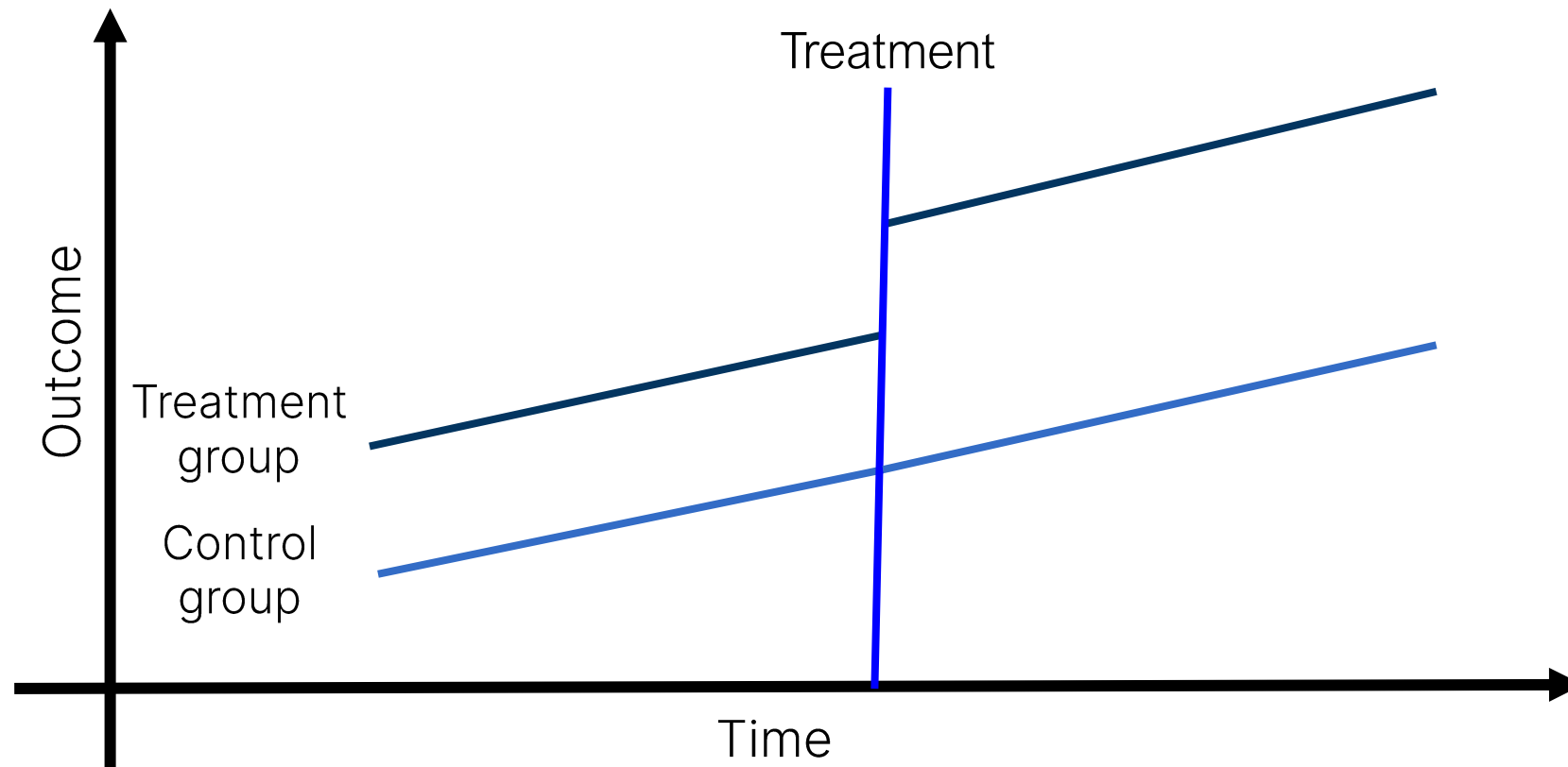
Prob > F = 0.4754

# The saga continues...



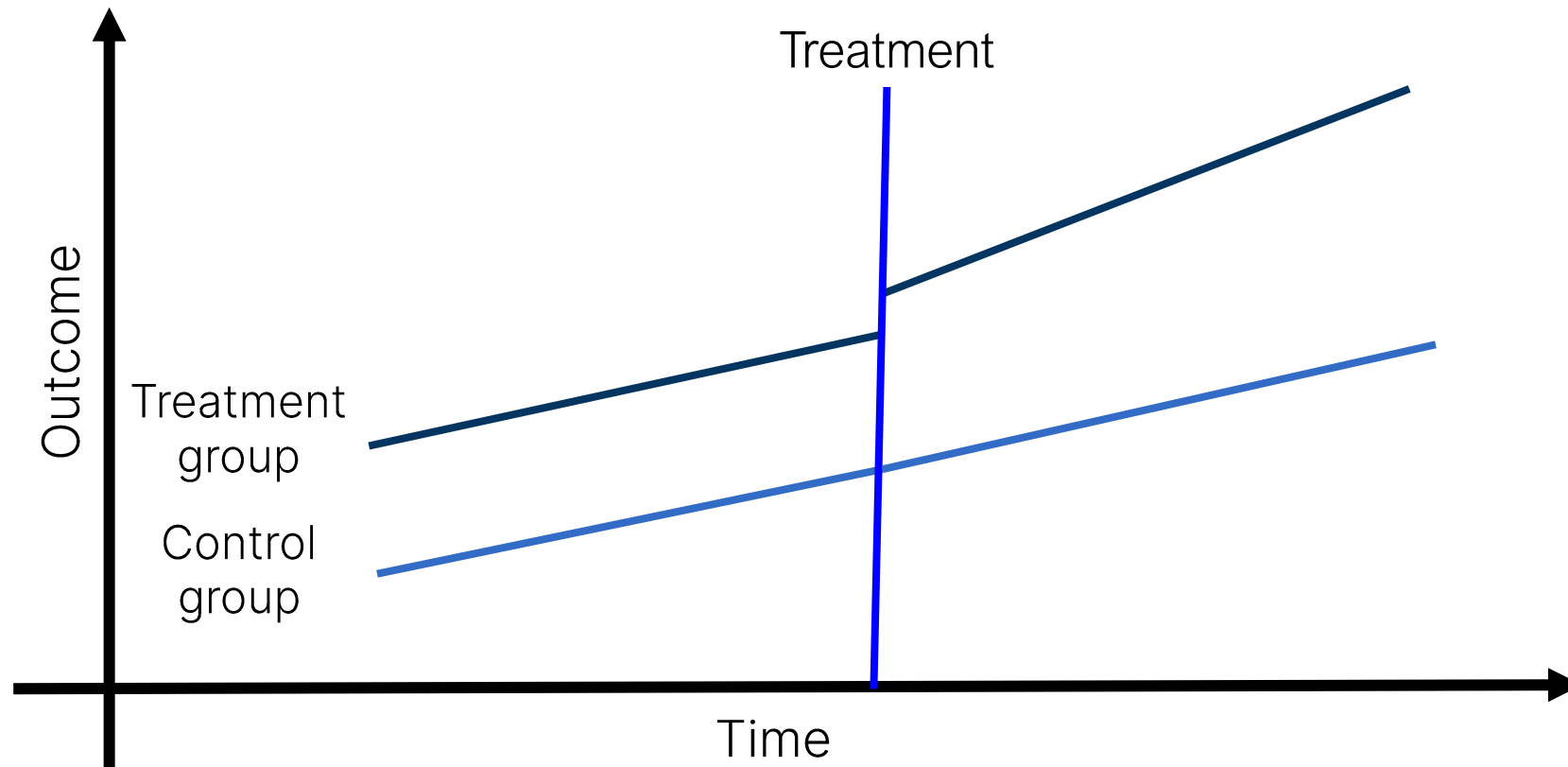


# DID

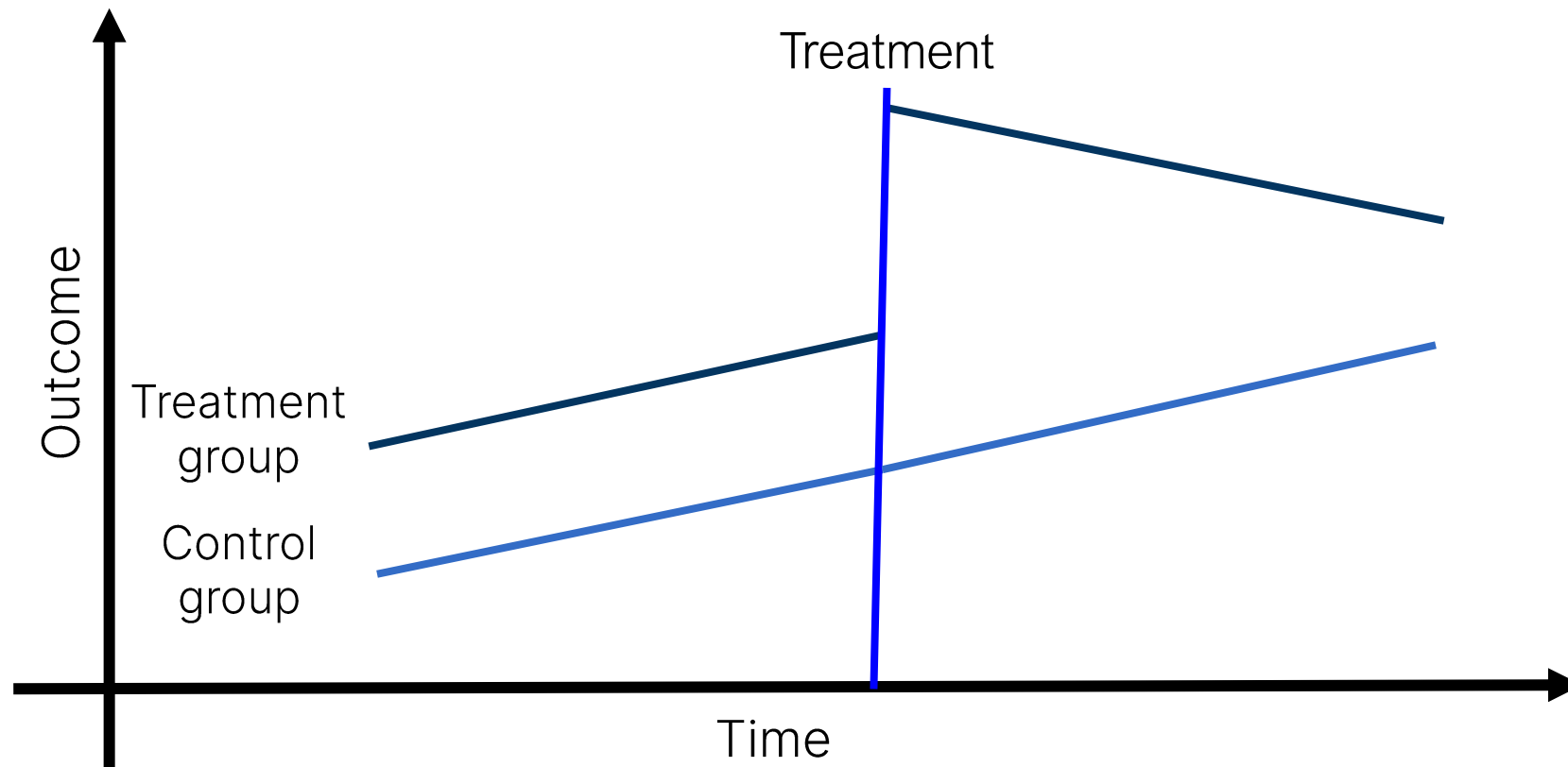




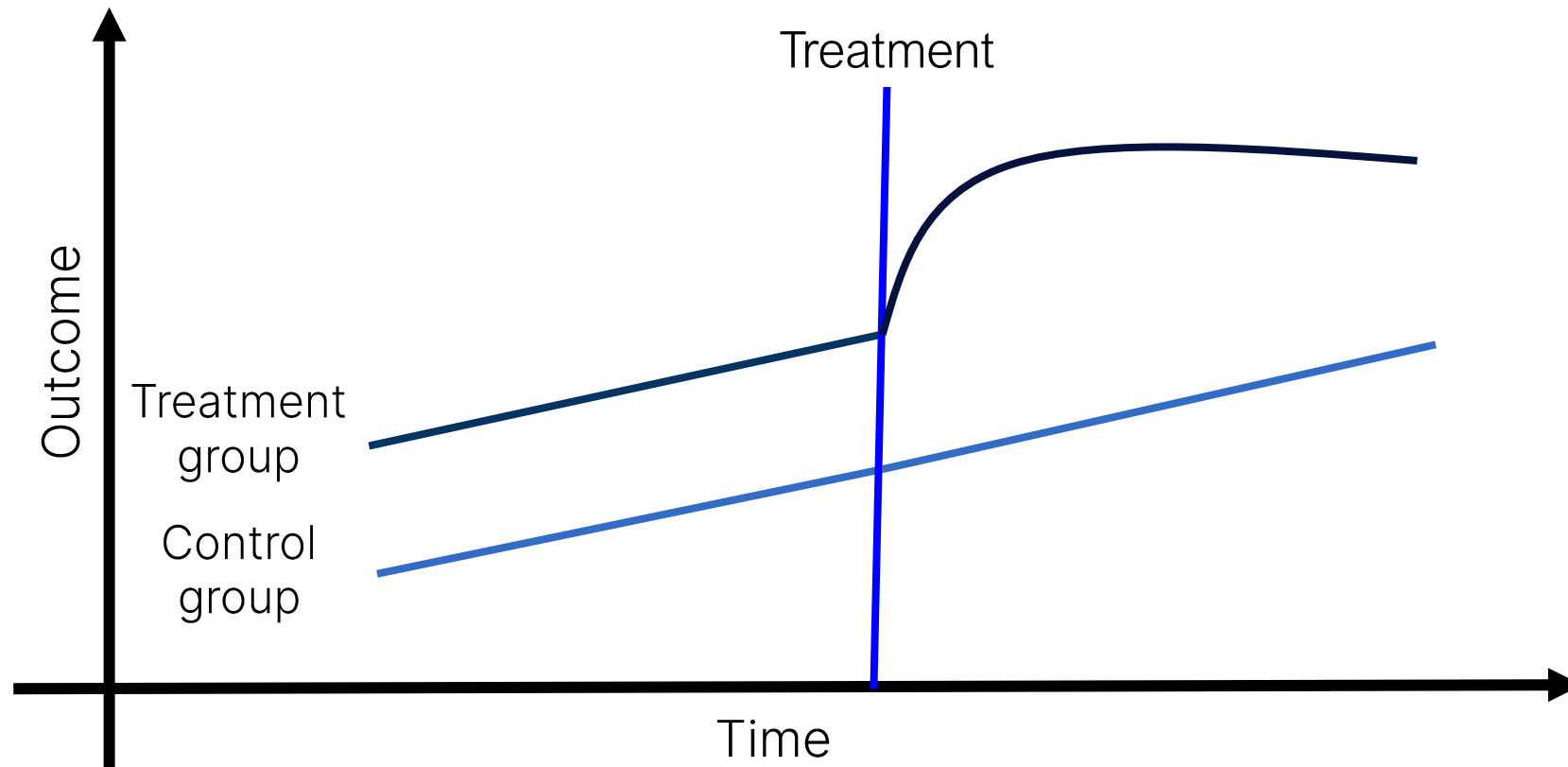
# Heterogeneous DID



# Heterogeneous DID

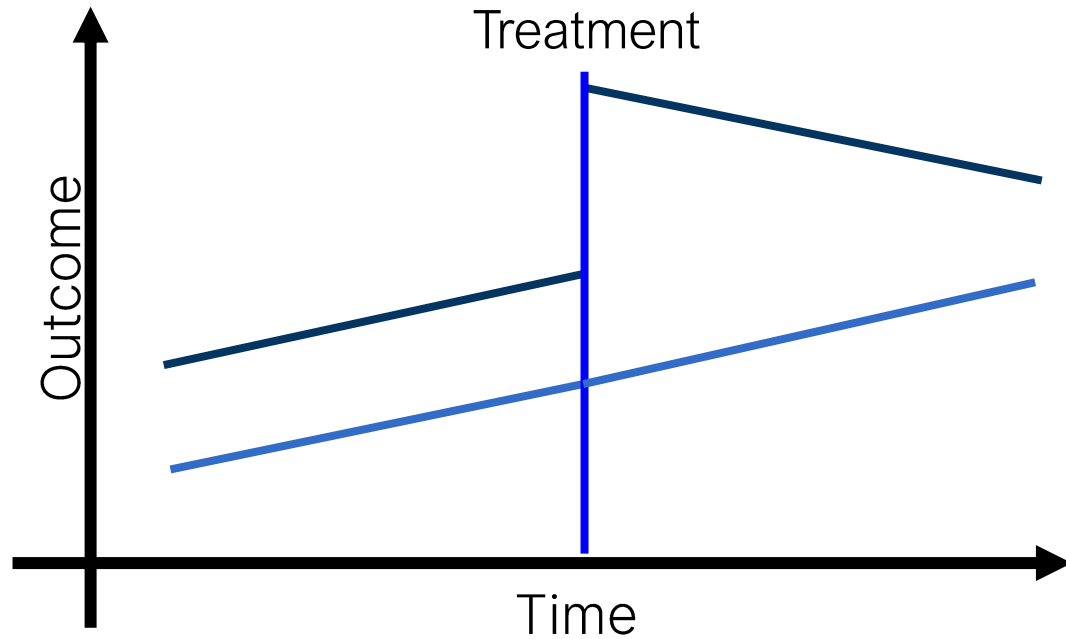


# Heterogeneous DID

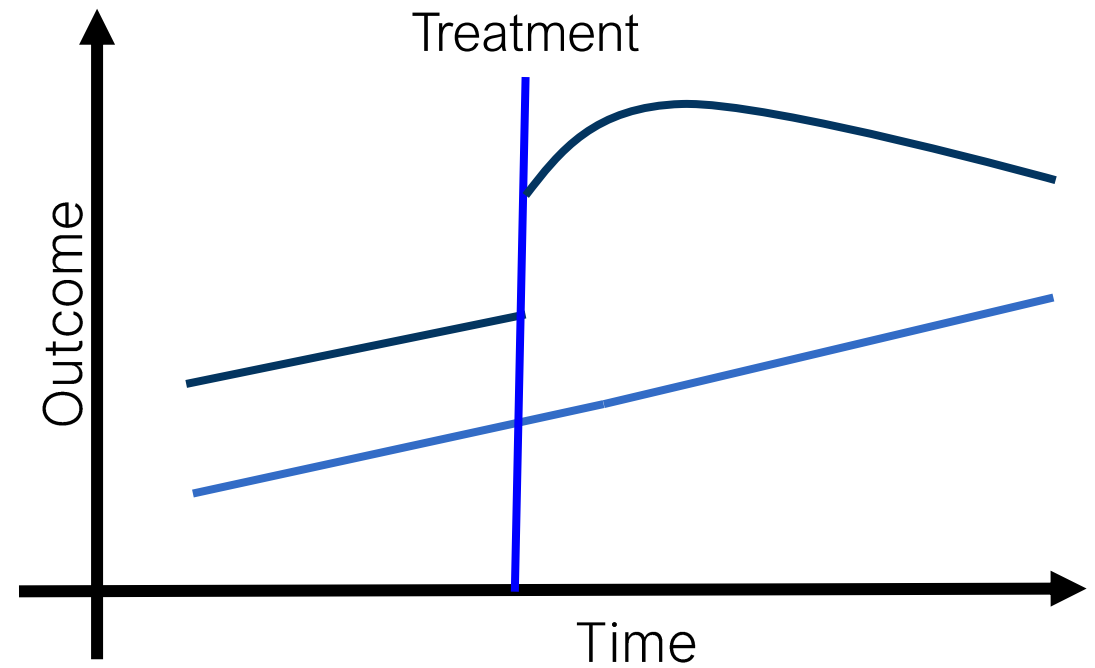


# Heterogeneous DID

2018 Cohort



2020 Cohort



# Heterogeneous TWFE

- First, create a cohort variable:

```
. bysort district hhabit: egen firstyear = min(year)
. bysort district: egen cohort = max(firstyear)
. replace cohort = 0 if !treatment
. egen tagme = tag(district year)
. list district year hhabit firstyear cohort if district==1 & tagme
```

	district	year	hhabit	firsty~r	cohort
1.	1	2017	No	2016	2018
2.	1	2016	No	2016	2018
102.	1	2022	Yes	2018	2018
103.	1	2024	Yes	2018	2018
104.	1	2020	Yes	2018	2018
106.	1	2023	Yes	2018	2018
107.	1	2018	Yes	2018	2018
110.	1	2019	Yes	2018	2018
119.	1	2021	Yes	2018	2018

# Heterogeneous TWFE

$$y_{igt} = \gamma_g + \gamma_t + \mathbf{z}_{igt}\beta + \gamma_{t|post}\gamma_C D_{Ct}\delta_{Ct} + \epsilon_{igt}$$

*i* observation  
*g* group  
*t* time  
*C* cohort

- Include the three-way interaction with post time and cohort

```
. regress bmi i.district i.year treatment#0.pre#year#cohort, vce(cluster district)
```

- Or use command `hdidregress`

```
. hdidregress twfe (bmi) (hhabit), group(district) time(year)
```

# Heterogeneous TWFE

```
. hdidregress twfe (bmi i.girl i.sports) (hhabit), group(district) time(year)
```

Treatment and time information

Time variable: year

Time interval: 2016 to 2024

Control:        `_did_cohort = 0`

Treatment:     `_did_cohort > 0`

	<code>_did_cohort</code>
Number of cohorts	4
Number of obs	
Never treated	11355
2018	1231
2020	2097
2022	2042

Heterogeneous treatment-effects regression

Data type:        Repeated cross-sectional

Estimator:       Two-way fixed effects

Treatment level: district

Control group:   Never treated

Heterogeneity:   Cohort and time

Number of obs = 16,725

Heterogeneous treatment-effects regression  
 Data type: Repeated cross-sectional  
 Estimator: Two-way fixed effects  
 Treatment level: district  
 Control group: Never treated  
 Heterogeneity: Cohort and time

Number of obs = 16,725

(Std. err. adjusted for 40 clusters in district)

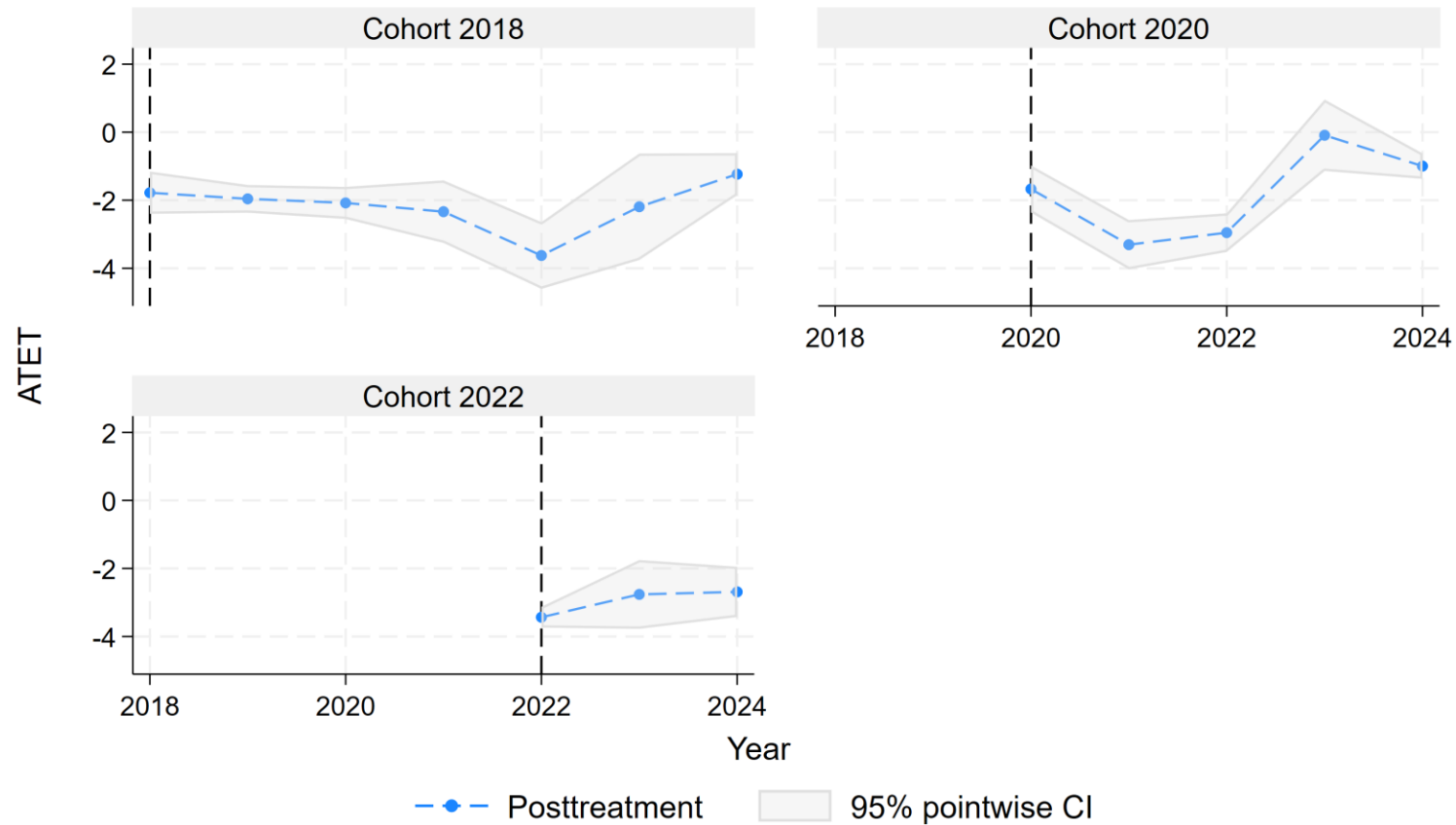
Cohort	ATET	Robust std. err.	t	P> t	[95% conf. interval]	
2018						
year						
2018	-1.778337	.3072316	-5.79	0.000	-2.399771	-1.156902
2019	-1.958607	.2017833	-9.71	0.000	-2.366752	-1.550461
2020	-2.078508	.2329257	-8.92	0.000	-2.549645	-1.607371
2021	-2.335608	.4538528	-5.15	0.000	-3.253612	-1.417604
2022	-3.623921	.484945	-7.47	0.000	-4.604815	-2.643027
2023	-2.191493	.7730937	-2.83	0.007	-3.755223	-.6277634
2024	-1.233747	.3058222	-4.03	0.000	-1.852331	-.6151629
2020						
year						
2020	-1.66891	.3419441	-4.88	0.000	-2.360557	-.9772632
2021	-3.306659	.3584147	-9.23	0.000	-4.031621	-2.581697
2022	-2.952263	.2806694	-10.52	0.000	-3.51997	-2.384556
2023	-.0899332	.518215	-0.17	0.863	-1.138122	.9582555
2024	-.9960773	.1854916	-5.37	0.000	-1.37127	-.6208851
2022						
year						
2022	-3.43286	.1509675	-22.74	0.000	-3.738221	-3.127499
2023	-2.762743	.4995317	-5.53	0.000	-3.773141	-1.752344
2024	-2.688314	.3648448	-7.37	0.000	-3.426282	-1.950346

Note: ATET computed using covariates.



# Visualize ATETs by cohort and time

```
. estat atetplot
```



# Aggregations

<code>estat aggregation, overall</code>	aggregate ATETs within cohorts and time periods; the default
<code>dynamic</code>	aggregate ATETs within exposures to the treatment
<code>time</code>	aggregate ATETs within time periods
<code>cohort</code>	aggregate ATETs within cohorts
<code>graph</code>	display the aggregation plot

# Aggregations

```
. estat aggregation
```

Overall ATET

Number of obs = 16,725

(Std. err. adjusted for 40 clusters in district)

bmi	ATET	Robust std. err.	t	P> t	[95% conf. interval]	
hhabit (Yes vs No)	-2.226743	.18228	-12.22	0.000	-2.595439	-1.858047

# Aggregations

```
. estat aggregation, dynamic graph
```

Duration of exposure ATET

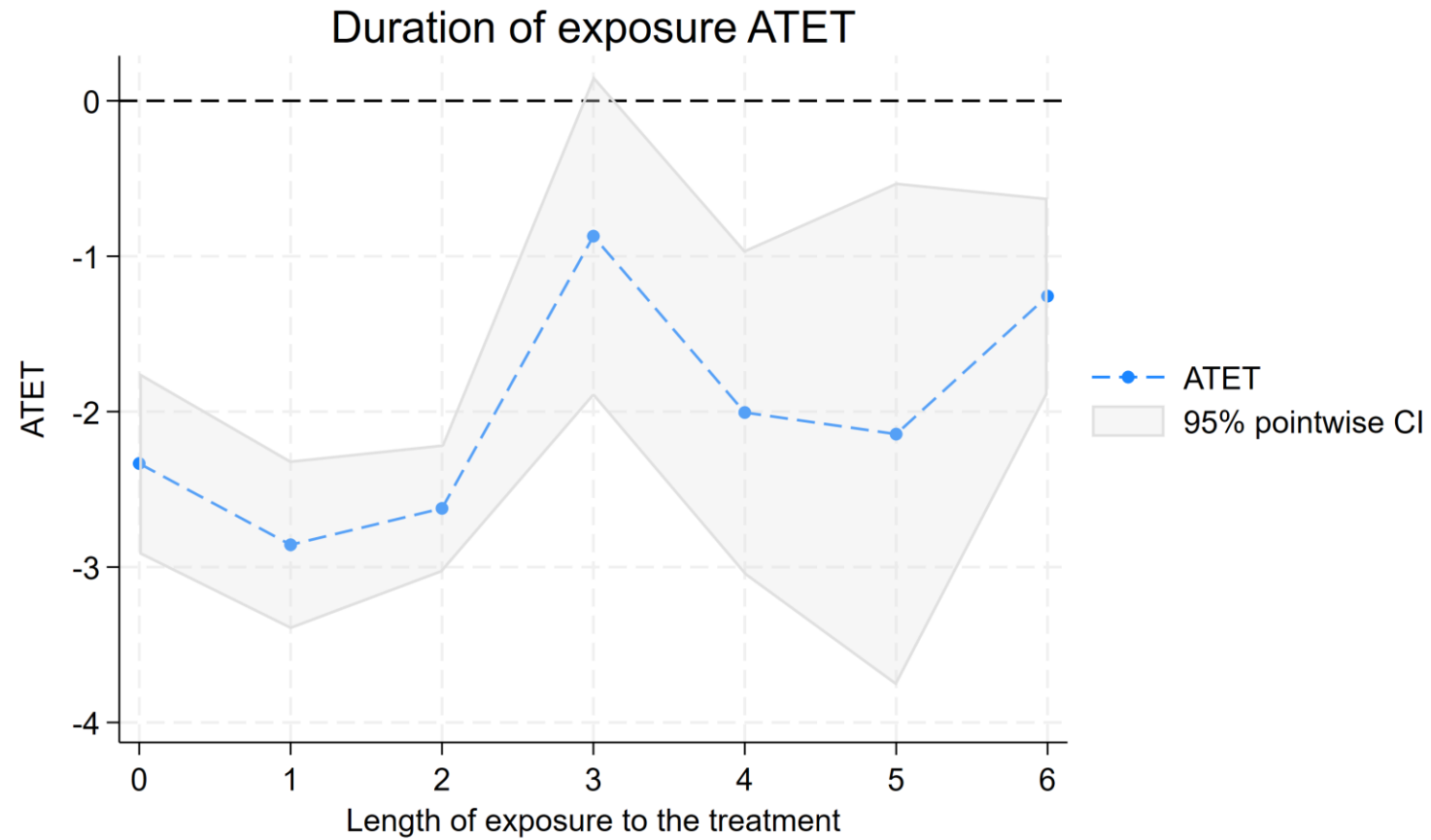
Number of obs = 16,725

(Std. err. adjusted for 40 clusters in district)

Exposure	ATET	Robust std. err.	t	P> t	[95% conf. interval]	
0	-2.333457	.288695	-8.08	0.000	-2.917398	-1.749516
1	-2.856789	.2685596	-10.64	0.000	-3.400002	-2.313576
2	-2.622523	.2035392	-12.88	0.000	-3.03422	-2.210826
3	-.8707736	.5103908	-1.71	0.096	-1.903136	.1615892
4	-2.005368	.5174664	-3.88	0.000	-3.052043	-.9586937
5	-2.144589	.8003042	-2.68	0.011	-3.763357	-.5258206
6	-1.256447	.3128612	-4.02	0.000	-1.889268	-.6236253

Note: Exposure is the number of periods since the first treatment time.

# Aggregations



# Aggregations

```
. estat aggregation, time graph
```

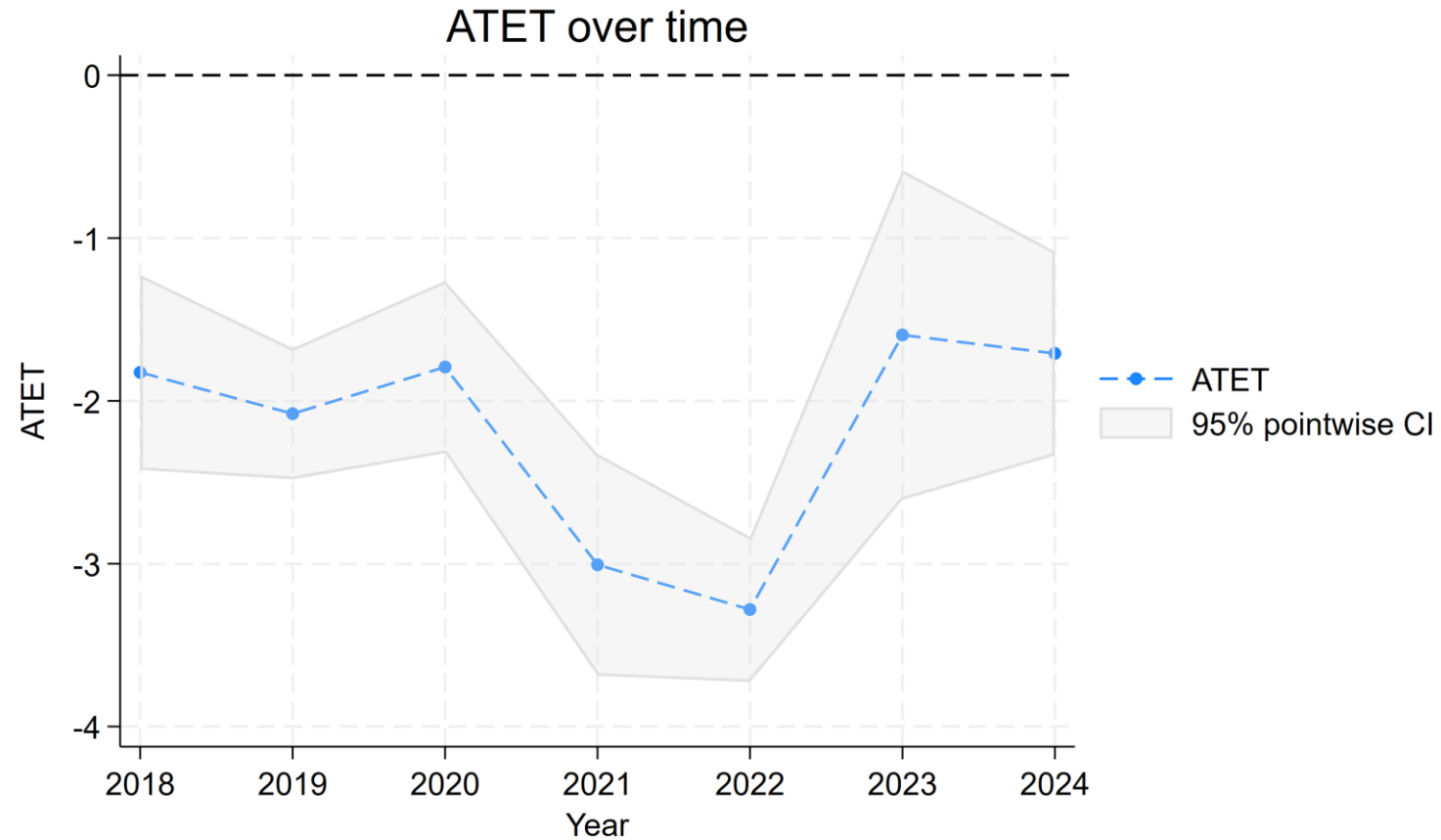
ATET over time

Number of obs = 16,725

(Std. err. adjusted for 40 clusters in district)

Time	ATET	Robust std. err.	t	P> t	[95% conf. interval]	
2018	-1.82485	.296194	-6.16	0.000	-2.423959	-1.225741
2019	-2.078951	.1990162	-10.45	0.000	-2.4815	-1.676403
2020	-1.791954	.2616474	-6.85	0.000	-2.321186	-1.262722
2021	-3.006472	.3371145	-8.92	0.000	-3.68835	-2.324593
2022	-3.281085	.220464	-14.88	0.000	-3.727016	-2.835155
2023	-1.595299	.5005498	-3.19	0.003	-2.607757	-.5828419
2024	-1.708546	.309853	-5.51	0.000	-2.335283	-1.081809

# Aggregations



# Aggregations

```
. estat aggregation, cohort
```

ATET over cohort

Number of obs = 16,725

(Std. err. adjusted for 40 clusters in district)

Cohort	ATET	Robust std. err.	t	P> t	[95% conf. interval]	
2018	-2.181784	.2444538	-8.93	0.000	-2.676238	-1.687329
2020	-1.851562	.1949819	-9.50	0.000	-2.24595	-1.457174
2022	-2.955739	.2123999	-13.92	0.000	-3.385358	-2.526119



# Constrain cohort effects

```
. hdidregress twfe (bmi i.girl i.sports) (hhabit), group(district) time(year) hettype(time)
```

```
Heterogeneous treatment-effects regression          Number of obs = 16,725  
Data type:      Repeated cross-sectional  
Estimator:      Two-way fixed effects  
Treatment level: district  
Control group:  Never treated  
Heterogeneity:  Time
```

(Std. err. adjusted for 40 clusters in district)

Cohort	ATET	Robust std. err.	t	P> t	[95% conf. interval]	
year						
2018	-1.894845	.3354881	-5.65	0.000	-2.573434	-1.216256
2019	-1.970336	.3191638	-6.17	0.000	-2.615906	-1.324766
2020	-2.084416	.2635516	-7.91	0.000	-2.6175	-1.551333
2021	-3.287492	.3853363	-8.53	0.000	-4.066908	-2.508075
2022	-3.354472	.2075074	-16.17	0.000	-3.774195	-2.934748
2023	-1.65903	.495814	-3.35	0.002	-2.661908	-.6561514
2024	-1.77315	.3215191	-5.51	0.000	-2.423484	-1.122817

Note: ATET computed using covariates.

.

# Constrain time effects

```
. hdidregress twfe (bmi i.girl i.sports) (hhabit), group(district) time(year) hettype(cohort)
```

```
Heterogeneous treatment-effects regression          Number of obs = 16,725  
Data type:      Repeated cross-sectional  
Estimator:      Two-way fixed effects  
Treatment level: district  
Control group:  Never treated  
Heterogeneity:  Cohort
```

(Std. err. adjusted for 40 clusters in district)

Cohort	ATET	Robust std. err.	t	P> t	[95% conf. interval]	
2018	-2.186048	.2379634	-9.19	0.000	-2.667375	-1.704722
2020	-1.83563	.1947801	-9.42	0.000	-2.22961	-1.44165
2022	-3.027139	.2130329	-14.21	0.000	-3.458038	-2.596239

Note: ATET computed using covariates.

.

# Other estimators

Estimator	Option	Outcome model	Treatment model
Two-way fixed effect	<code>twfe</code>	YES	NO
Regression adjustment	<code>ra</code>	YES	NO
Inverse-probability weighting	<code>ipw</code>	NO	YES
Augmented inverse-probability weighting	<code>aipw</code>	YES	YES

---

# Heterogeneous DID with AIPW

```
. hdidregress aipw (bmi i.girl i.sports) (hhabit parksd medu), group(district) time(year)
```

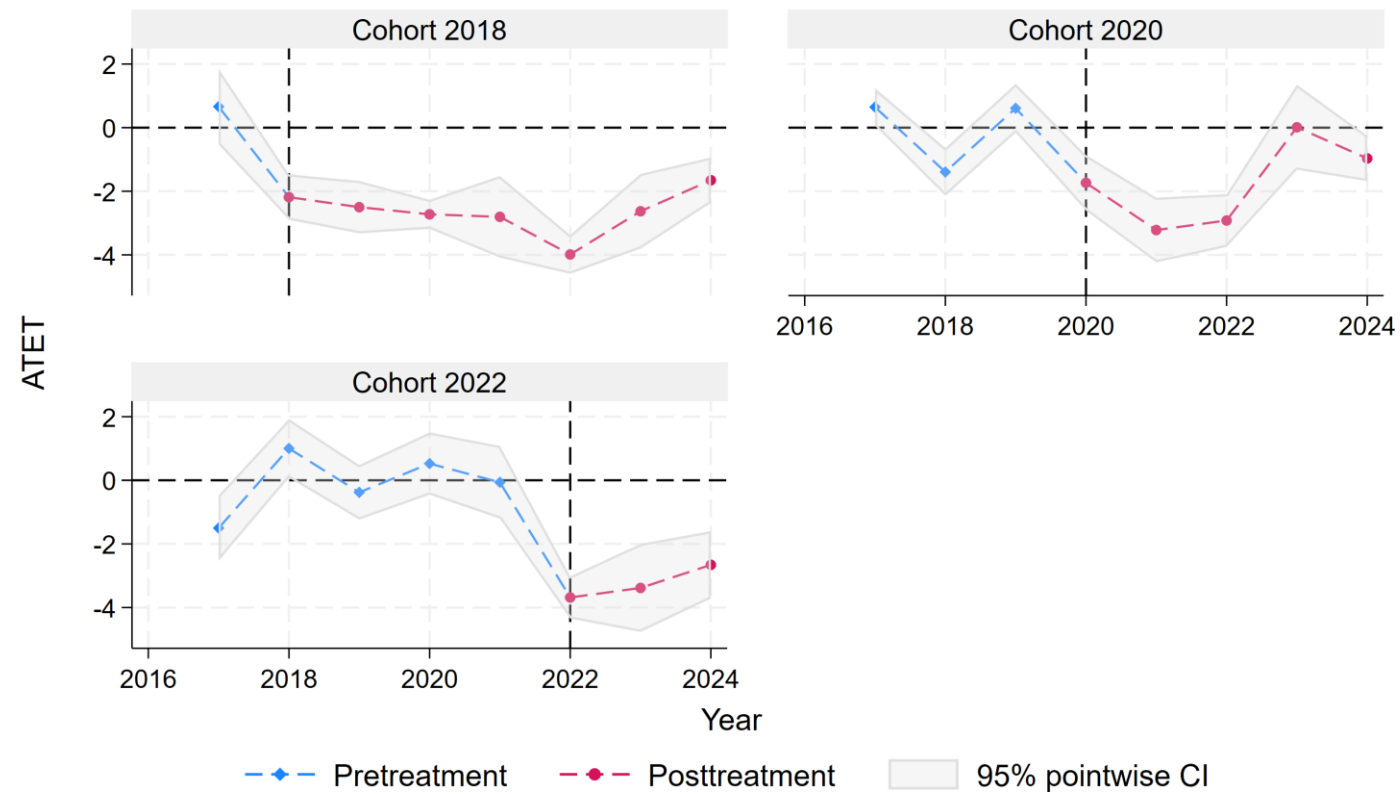
```
Heterogeneous treatment-effects regression          Number of obs = 16,725
Estimator:      Augmented IPW
Treatment level: district
Control group:  Never treated
```

(Std. err. adjusted for 40 clusters in district)

Cohort	ATET	Robust std. err.	z	P> z	[95% conf. interval]	
2018						
year						
2017	.6610372	.6068607	1.09	0.276	-.5283878	1.850462
2018	-2.184167	.3638684	-6.00	0.000	-2.897336	-1.470998
2019	-2.499109	.4201546	-5.95	0.000	-3.322597	-1.675621
2020	-2.724452	.232414	-11.72	0.000	-3.179975	-2.268929
2021	-2.803982	.6541718	-4.29	0.000	-4.086136	-1.521829
2022	-3.986615	.310742	-12.83	0.000	-4.595658	-3.377572
2023	-2.627951	.59908	-4.39	0.000	-3.802126	-1.453776
2024	-1.653263	.3645688	-4.53	0.000	-2.367805	-.9387212
2020						
year						
2017	.6468236	.3014149	2.15	0.032	.0560614	1.237586
2018	-1.39342	.3832857	-3.64	0.000	-2.144646	-.6421936
2019	.6095265	.3913018	1.56	0.119	-.157411	1.376464
2020	-1.734927	.4457704	-3.89	0.000	-2.608621	-.8612329
2021	-3.21896	.5187638	-6.21	0.000	-4.235718	-2.202201
2022	-2.010625	.4225200	-4.76	0.000	-2.742711	-1.278539
2023	-2.010625	.4225200	-4.76	0.000	-2.742711	-1.278539
2024	-2.010625	.4225200	-4.76	0.000	-2.742711	-1.278539

# Heterogeneous DID with AIPW

```
. estat atetplot
```



# Test for parallel trends

```
. estat ptrends
```

Parallel-trends test (pretreatment time period)

H0: Treatment effects in all the pretreatment periods are zero

```
chi2(9) = 33.93
```

```
Prob > chi2 = 0.0001
```



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