

# A Holistic Bayesian Approach to Addressing Measurement Reactivity with a Planned Missing Data Design

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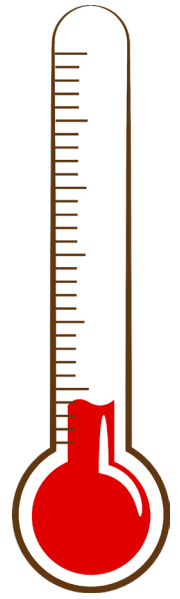
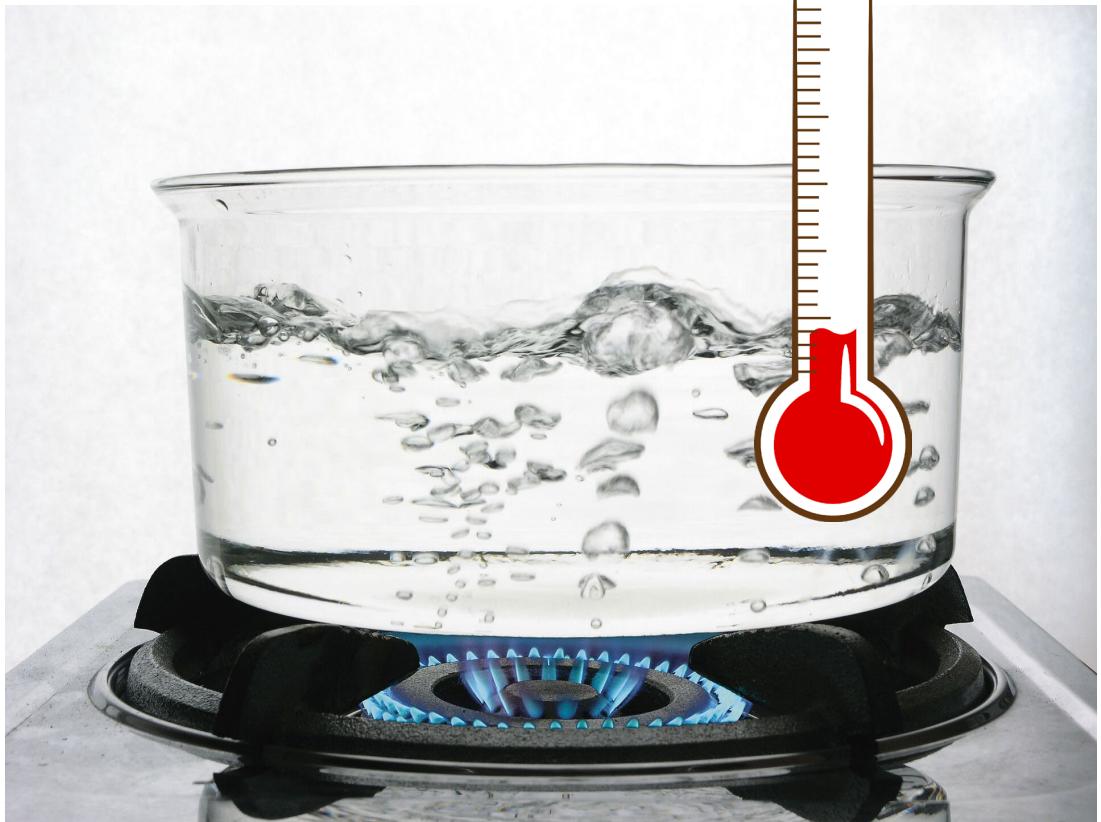
# Outline of Talk

1. Define **measurement reactivity** and show how it confounds research with **pre-test post-test designs**
2. Redefine measurement reactivity as a **missing data problem**
3. Show example in the literature for using **missing data imputation** to cure a **confounding effect of measurement reactivity**
4. Describe a classic **advice taking experimental design** as a simple example of a **pre-test post-test design**
5. Demonstrate how **advice changes people's minds** when their **prior beliefs are not directly measured** using fully Bayesian **joint imputation and measurement model**

# The Observer Effect

It is impossible to measure something without changing it

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# Measurement Reactivity

I wonder if I've grown at all....



This problem is hard, I wonder if he knows the answer....

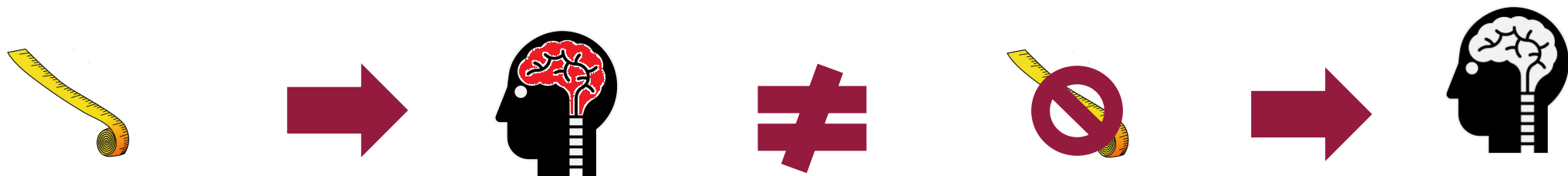


# When Does Measurement Reactivity Matter?

We often wish to understand the change that occurs as a result of a **treatment** or **intervention**

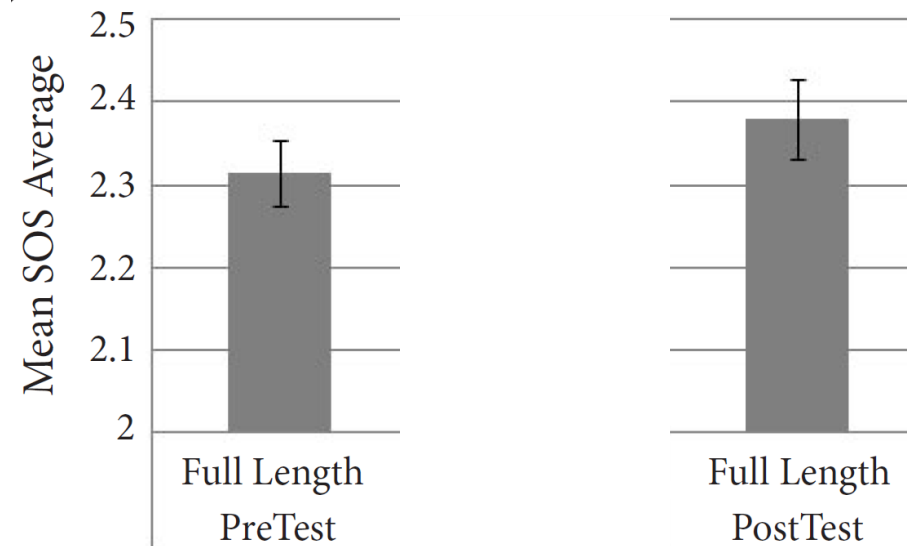


But existence of a pretest can change the effects of interventions and **threaten generalizability**



# Treatment Effects in Control Groups

- In RCTs with pre-tests and post-tests, results often show treatment-like effects even in control groups
- Such results reflect learning as a result of pre-test exposure
- Harel et al. (2011) showed increases in scores on a measure of attitudes toward suicide among control participants



# Missing Data Mechanisms

## Terms

- **missing:** Patterns of missing data
- **data:** Values of the actual data
- **x:** Set of completely observed predictors correlated with incomplete data *and* patterns of missingness

**Missing Completely At Random:** Patterns of missingness are independent of the data that is missing

- $P(\text{missing} | \text{data}, x) = P(\text{missing})$

**Missing at Random:** Patterns of missingness are *conditionally* independent of the data that is missing

- $P(\text{missing} | \text{data}, x) = P(\text{missing} | x)$

**Missing Not at Random:** Patterns of missingness are not independent of the data that is missing

- $P(\text{missing} | \text{data}, x) \neq P(\text{missing} | x)$

Imputation  
Possible

Imputation  
Not Possible

# Reframing Measurement Reactivity

Missing data mechanisms deal with whether *missingness depends on data*

Measurement reactivity is a *counterfactual* question about whether *data depends on missingness*

Measurement reactivity:  $P(\text{data} \mid \text{missing}) \neq P(\text{data})$



# Modeling Solution

- Find domains where measurement reactivity is a problem
  - $P(\text{data} | \text{missing}) \neq P(\text{data})$
- Ensure assumptions for imputation are met
  - MCAR:  $P(\text{missing} | \text{data}) = P(\text{missing})$ , or at least
  - MAR:  $P(\text{missing} | \text{data}, x) = P(\text{missing} | x)$
- Impute missing values
- Compare imputed missing values to observed values to understand patterns in data in the absence of measurement reactivity

# *Planned* Missing Data Design

Sometimes collectable data is left missing by design (Graham et al., 2006; Rhemtulla & Hancock 2016). Typically done to reduce

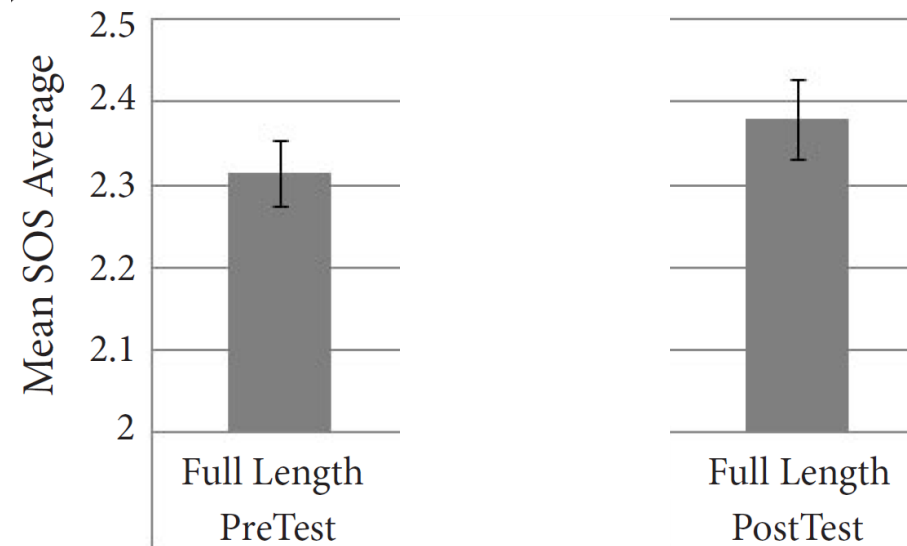
- Instrument length
- Participant burden
- Attrition

For measurement reactivity research, we do something similar

- Experimentally randomize missing data
- Impute
- Draw causal inference

# Treatment Effects in Control Groups

- In RCTs with pre-tests and post-tests, results often show treatment-like effects even in control groups
- Such results reflect learning as a result of pre-test exposure
- Harel et al. (2011) showed increases in scores on a measure of attitudes toward suicide among control participants
- However, when truncated pretests were used with multiple imputation for the missing data, no treatment effects were found in control group



# Multiple Imputation vs. Fully Bayesian Approach

## Multiple Imputation

- Fit imputation model
- Extract subset of posterior draws to represent uncertainty in missing values
- Fit measurement model on each draw
- Combine/average results of measurement models

## Fully Bayesian

- Fit imputation and measurement model with joint posterior
- Treat missing values as parameters / latent variables in measurement model
- Utilize full posterior uncertainty of missing values
- Only need to fit one model

# Fully Bayesian Imputation in Stan

## 3.1 Missing data

Stan treats variables declared in the `data` and `transformed data` blocks as known and the variables in the `parameters` block as unknown.

An example involving missing normal observations could be coded as follows.<sup>10</sup>

```
data {
  int<lower=0> N_obs;
  int<lower=0> N_mis;
  array[N_obs] real y_obs;
}
parameters {
  real mu;
  real<lower=0> sigma;
  array[N_mis] real y_mis;
}
model {
  y_obs ~ normal(mu, sigma);
  y_mis ~ normal(mu, sigma);
}
```

Observed data

Missing data

The number of observed and missing data points are coded as data with non-negative integer variables `N_obs` and `N_mis`. The observed data are provided as an array data variable `y_obs`. The missing data are coded as an array parameter, `y_mis`. The ordinary parameters being estimated, the location `mu` and scale `sigma`, are also coded as parameters. The model is vectorized on the observed and missing data; combining them in this case would be less efficient because the data observations would be promoted and have needless derivatives calculated.

<https://mc-stan.org/docs/stan-users-guide/missing-data.html>

# Advice Taking: Judge-Advisor Systems (JAS)

I wonder if the Mets are likely to make the playoffs



Agent or “judge” tasked with forming an objective belief

They Stink!  
Maybe a 10% chance!



Consults an “advisor” for new information

| FANGRAPHS |               |
|-----------|---------------|
| Team      | Make Playoffs |
| Mets      | 42.2%         |

Can consult many advisors, who need not be people

Who am I kidding, I'll go with 25%



Judge makes decision

# Studying Belief Change

To study belief-change as a result of advice, researchers will typically ask a judge what they believe **before** and **after** getting advice

Pre-test:  
Initial judgment ( $J_1$ )

I think maybe  
75%!



Treatment:  
Advice ( $J_a$ )

They Stink!  
Maybe a 10%  
chance!



Post-test:  
Final judgment ( $J_2$ )

Maybe Stephen A  
is right, I'll go  
with 25% instead



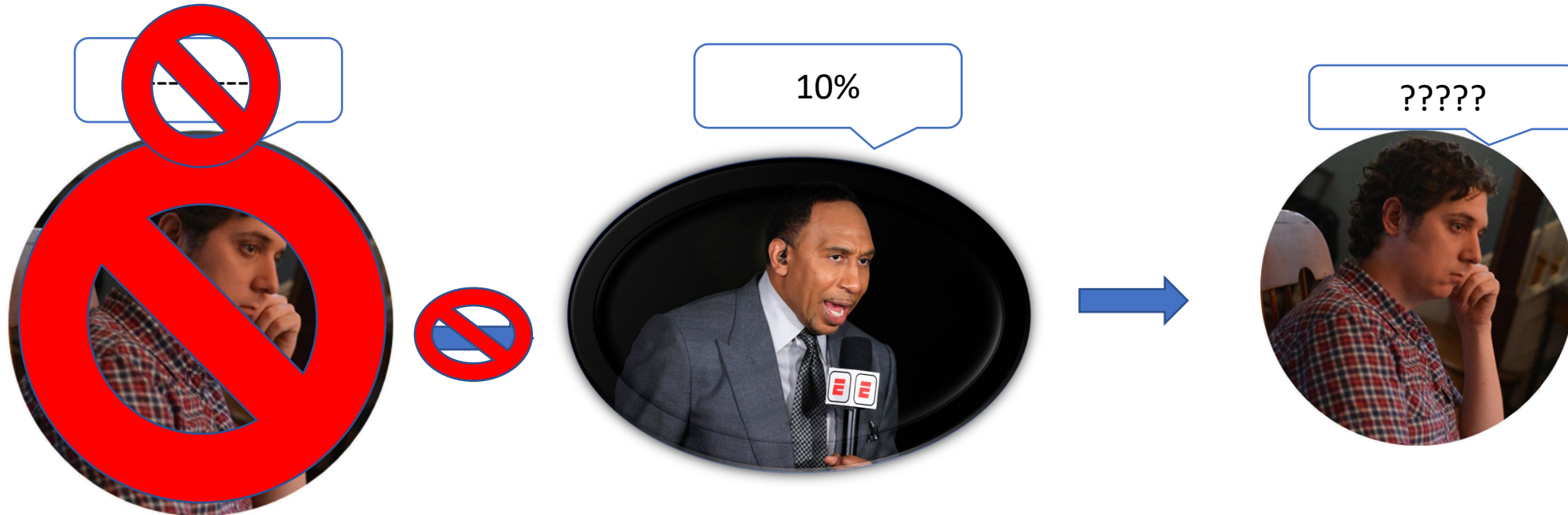
## Weight of Advice (WOA)

$$WOA = \frac{J_2 - J_1}{J_a - J_1} = \frac{25\% - 75\%}{10\% - 75\%} = .77$$

Implies

$$J_2 = WOA(J_a) + (1 - WOA)(J_1) = 25\% = .77(10\%) + .23(75\%)$$

# What happens if the pre-test is removed?

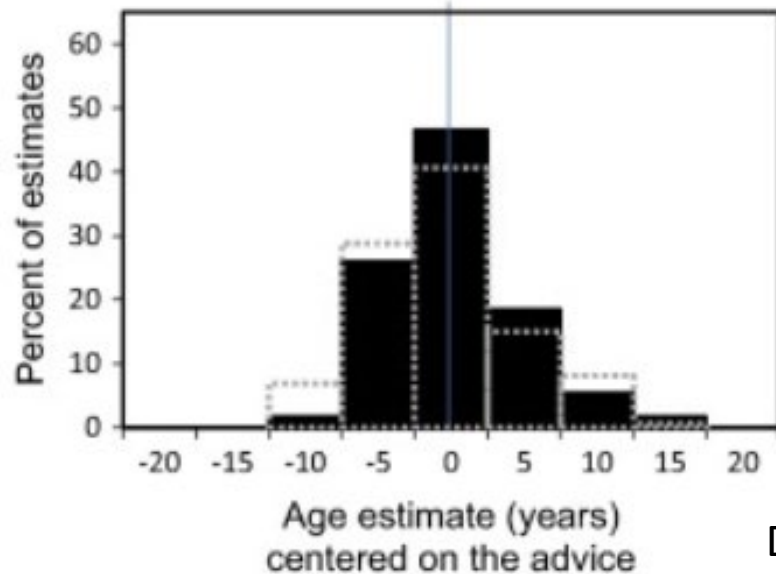




# Pushing Away vs Staying Close

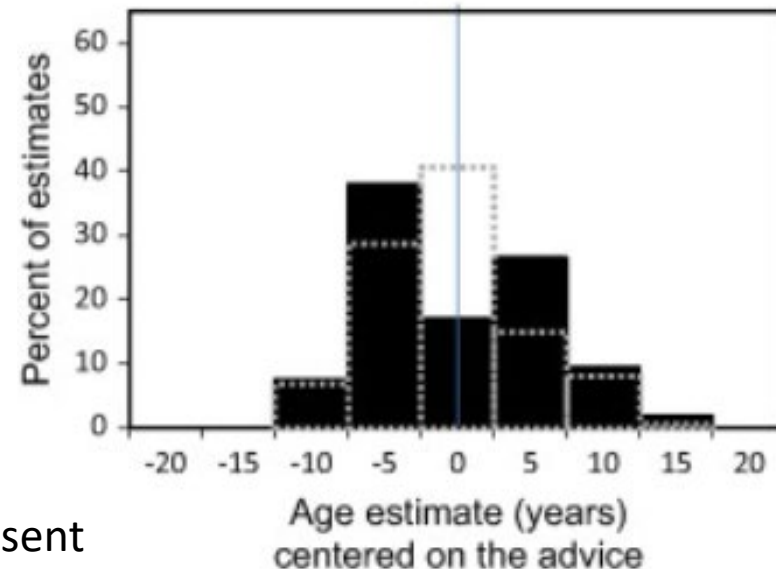
(Rader et al., 2015)

**Untruncated JAS:** Final Judgment ( $J_2$ ) when initial judgment ( $J_1$ ) **WAS** elicited



No “push away” effect

**Truncated JAS:** Final Judgment ( $J_2$ ) when initial judgment ( $J_1$ ) **WAS NOT** elicited



“Push away” effect

Difference implies  $P(\text{data} | \text{missing}) \neq P(\text{data})$

# Solution: Planned Missing Data Experimental Design

## Untruncated Condition

| Type | Item # | $J_1$    | $J_2$    |
|------|--------|----------|----------|
| mis  | 1      | Observed | Observed |
| mis  | 2      | Observed | Observed |
| mis  | :      | Observed | Observed |
| mis  | $m$    | Observed | Observed |
| obs  | 1      | Observed | Observed |
| obs  | 2      | Observed | Observed |
| obs  | :      | Observed | Observed |
| obs  | $o$    | Observed | Observed |

## Partially Truncated Condition

| Type | Item # | $J_1$           | $J_2$    |
|------|--------|-----------------|----------|
| mis  | 1      | <b>MISSING!</b> | Observed |
| mis  | 2      | <b>MISSING!</b> | Observed |
| mis  | :      | <b>MISSING!</b> | Observed |
| mis  | $m$    | <b>MISSING!</b> | Observed |
| obs  | 1      | Observed        | Observed |
| obs  | 2      | Observed        | Observed |
| obs  | :      | Observed        | Observed |
| obs  | $o$    | Observed        | Observed |

Imputation  
Model

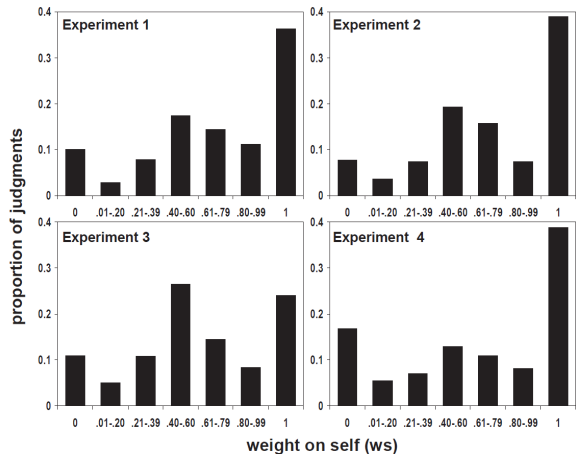
$$J_{1,mis,un} = \alpha' z_{un} + J_{1,obs,un} B + \epsilon_{un}$$

$$\delta_{mis,tr} = \alpha' z_{tr} + J_{1,obs,tr} B$$

$$\hat{J}_{1,mis,tr} \sim MVN(\delta_{mis,tr}, \hat{\Sigma}_{\epsilon_{un}})$$

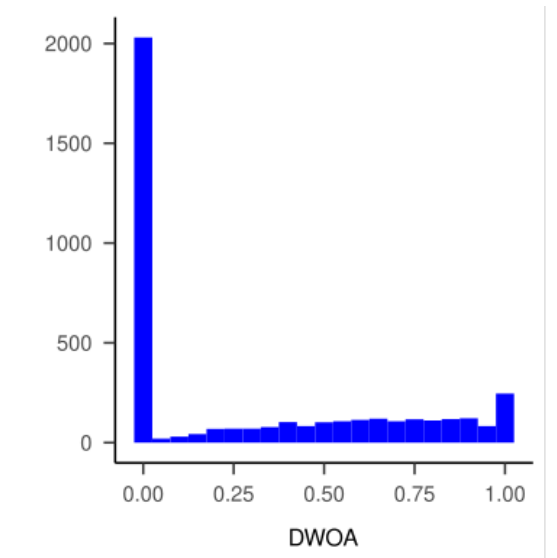
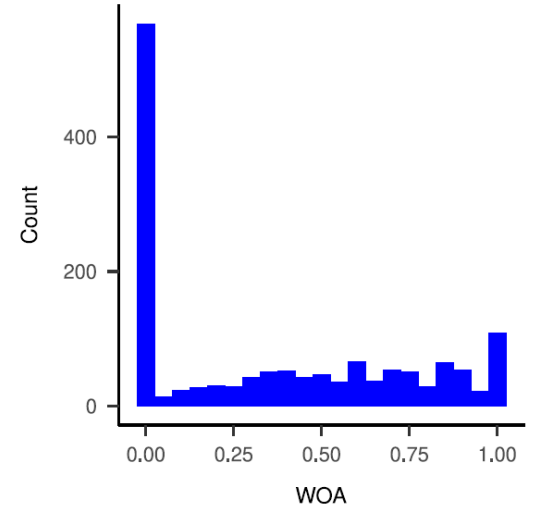
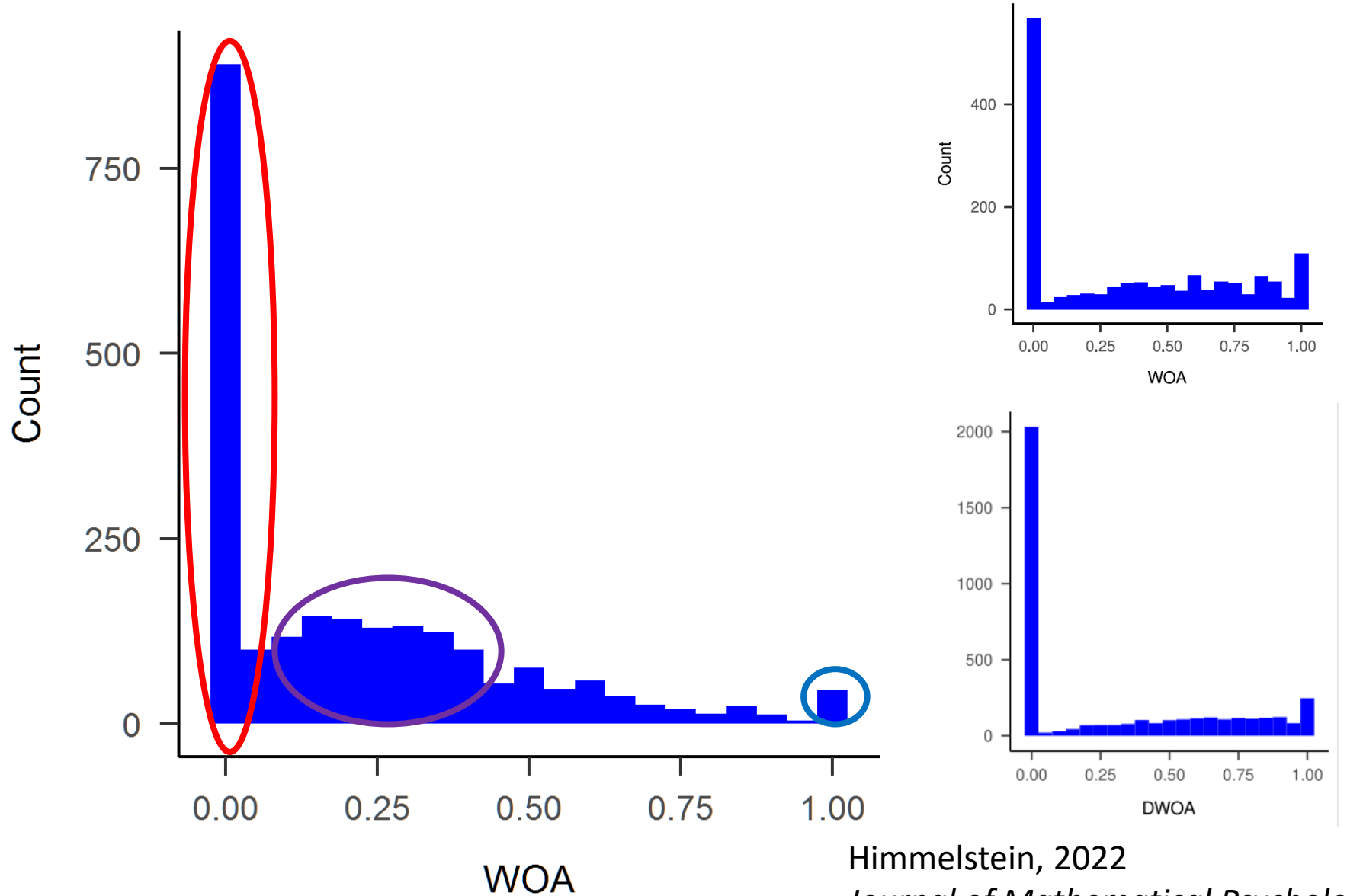
# The Empirical Distribution of Untruncated WOA

Three  
Distinct  
Modes



(Soll & Larrick, 2009)

Note:  $ws = 1 - WOA$



Himmelstein, 2022  
*Journal of Mathematical Psychology*

## Untruncated JAS: Dual Hurdle Model

- Treats  $WOA \in [0,1]$  as response variable
- Mixture between
  - **Categorical model:** Decline ( $WOA = 0$ ), adopt ( $WOA = 1$ ), or compromise ( $0 < WOA < 1$ )
  - **Continuous model:** averaging judgment conditional on compromise

## Truncated JAS: Hurdled Mixture Model

- Treats  $J_2 \in (-\infty, \infty)$  as response variable
- Adopt ( $J_2 = J_a$ ) remains categorical submodel
- Imputed  $\hat{J}_1$  contain measurement error, meaning  $\hat{J}_1 = J_2$  exactly will not occur naturally
- Replace discrete *decline* observations with a Gaussian component that incorporates measurement error

## Measurement Model: Hurdled Mixture

$$J_2 \begin{cases} = J_{a,mis,tr} \text{ with } P_{adopt} = \eta \\ = \hat{J}_1 \text{ with } P_{decline} = \tau(1 - \eta) \\ = \hat{C} \text{ with } P_{compr} = (1 - \tau)(1 - \eta) \end{cases}$$

$$\hat{J}_1 \sim MVN(\delta_{mis,tr}, \hat{\Sigma}_{\epsilon_{un}})$$

$$\hat{\mu} = \omega J_a + (1 - \omega) \hat{J}_1$$

$$\hat{C} \sim N(\hat{\mu}, \hat{\sigma}_{compr})$$

# Simulation Study

Simulated population correlation matrices  $\Sigma_{J_1}$  for  $K$  items. Each iteration varied:

- $\gamma$ : Proportion of variance accounted by first eigenvalue of  $\Sigma_{J_1}$
- $K$ : Total number of items
- $N$ : Total number of subjects
- $S$ : Proportion of items in partially truncated condition
- $D$ : Rate of declining advice
  - **Hypothesis 1:** ~47.5%, people anchor on prior belief, like untruncated JAS (big mode at  $J_1 = J_2$ )
  - **Hypotheses 2:** ~2.5%, people do not anchor on prior belief, unlike untruncated JAS (no big mode at  $J_1 = J_2$ )

# Simulation Study

- **Step 1:** Simulate  $N$  judges and their prior belief ( $J_1$ ) values for each of the  $K$  items based on  $\Sigma_{J_1}$
- **Step 2:** Simulate advice  $J_a$  for each item (based on  $\Sigma_{J_1}$  with a small amount of residual error)
- **Step 3:** simulate for each judge a parameter that defines
  - Stage 1 tendencies (discrete decision,  $\eta_j$ )
  - Stage 2 tendencies (continuous averaging judgment,  $\omega_j$ )

# Simulation Study

- **Step 4:** For each combination  $N$  and  $K$  simulate a realized outcome for discrete decision ( $\eta_{j,k}$ ) and averaging judgment ( $\omega_{j,k}$ )
- **Step 5:** For each judge and item, define  $J_2$  as follows
  - If discrete decision = decline,  $J_2 = J_1$
  - If discrete decision = adopt,  $J_2 = J_a$
  - If discrete decision = compromise,  $J_2 = \omega_{j,k} J_a + (1 - \omega_{j,k}) J_1$
- **Step 6:** Randomly select  $\frac{N}{3}$  simulated judges and  $\frac{K}{S}$  simulated items and delete  $J_1$  in all cases

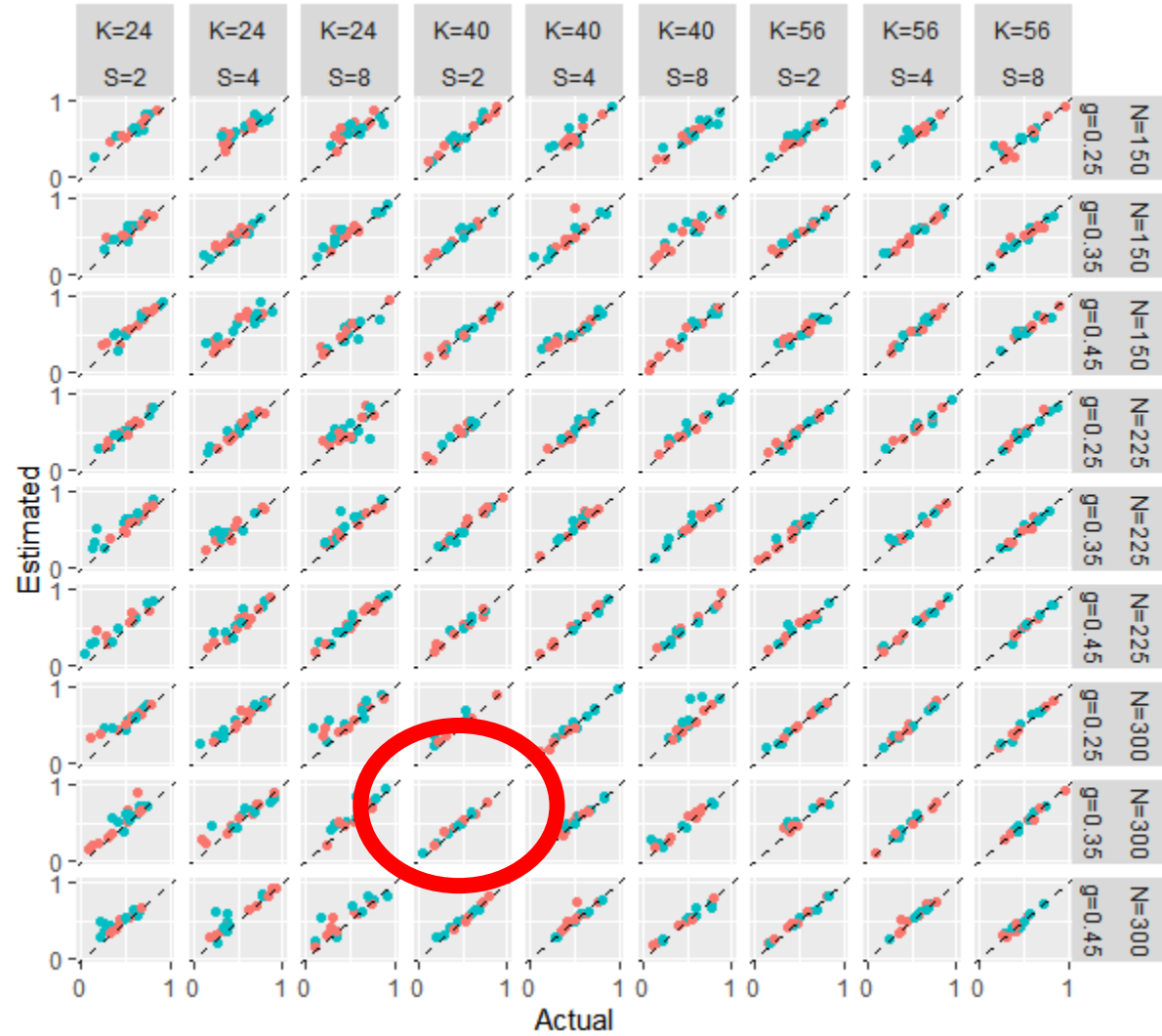
# Study 1: Simulated Schultze (2015) data

- **Step 7:** Use observed values from remaining  $N - \frac{N}{3}$  judges to estimate imputation model that uses remaining  $K - \frac{K}{S}$  untruncated items to predict  $\frac{K}{S}$  items that were artificially truncated
- **Step 8:** Impute  $J_1$  values that were deleted in **Step 6** based on imputation model estimated in **Step 7**
- **Step 9:** estimate hurdled mixture model using values imputed in **Step 8** and compare the estimated  $\hat{\eta}$  and  $\hat{\omega}$  to actual simulated  $\eta$  and  $\omega$

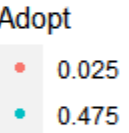
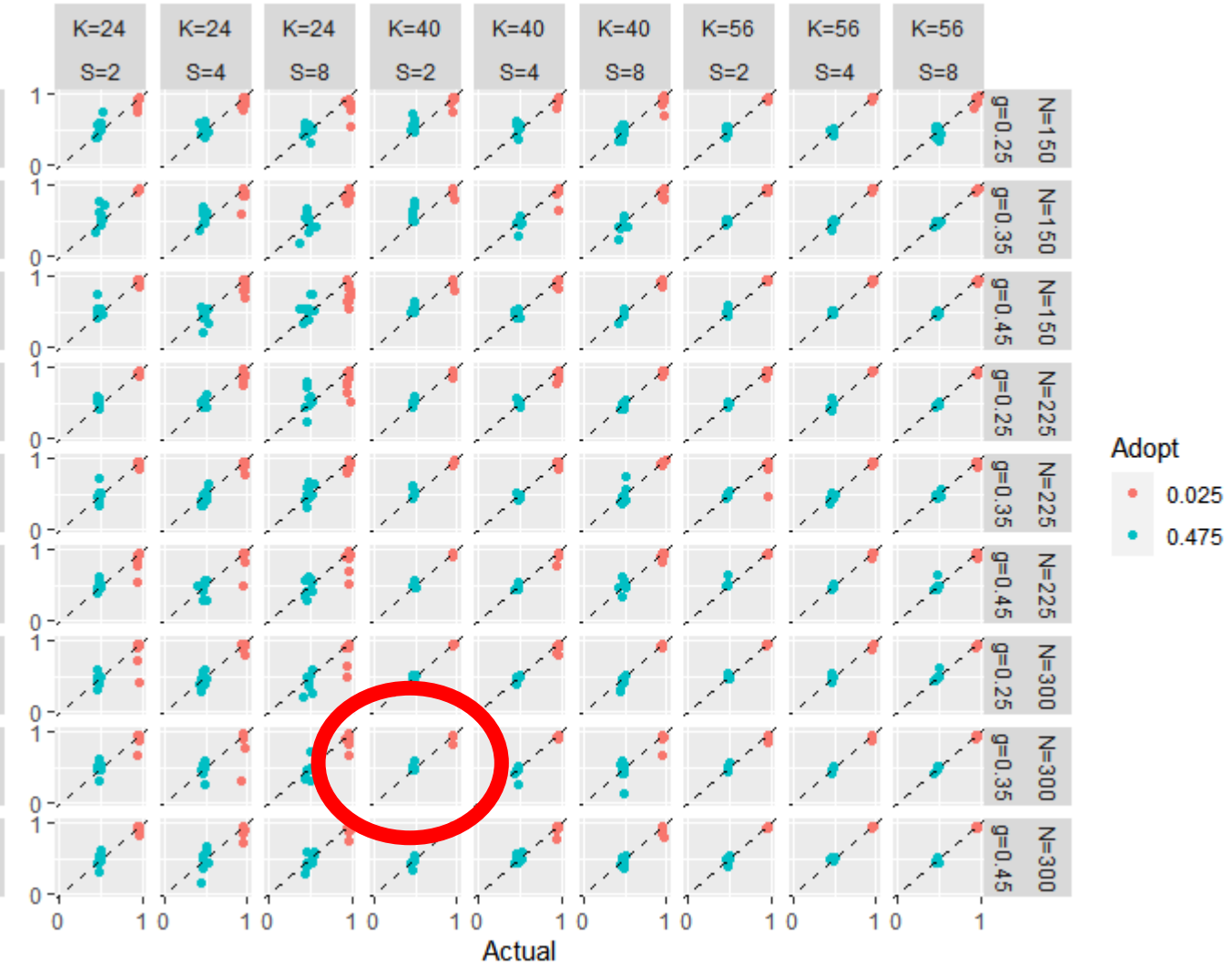


# Results

## WOA



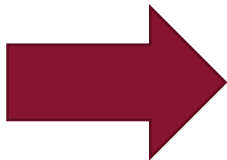
## Compromise



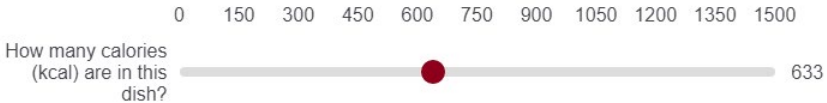
# A Real Truncated JAS

- Schultze et al. (2015) design was a JAS in which subjects estimated the number of calories in different foods.
- Took images from 40 different **Blue Apron** recipes to use in a JAS in three conditions
  - Untruncated (N = 196)
  - Partially Truncated (N = 148)
  - Fully truncated (N = 52, control condition to compare against partially truncated condition)

# Untruncated Item Example



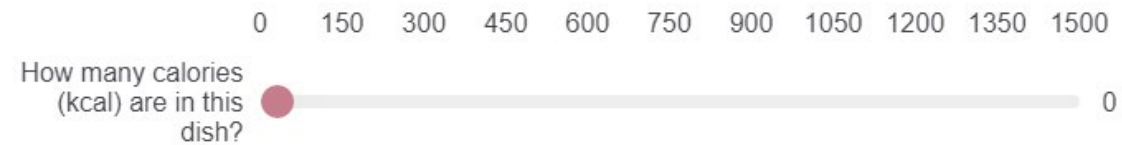
Your advisor made the following estimate: **450 Calories**



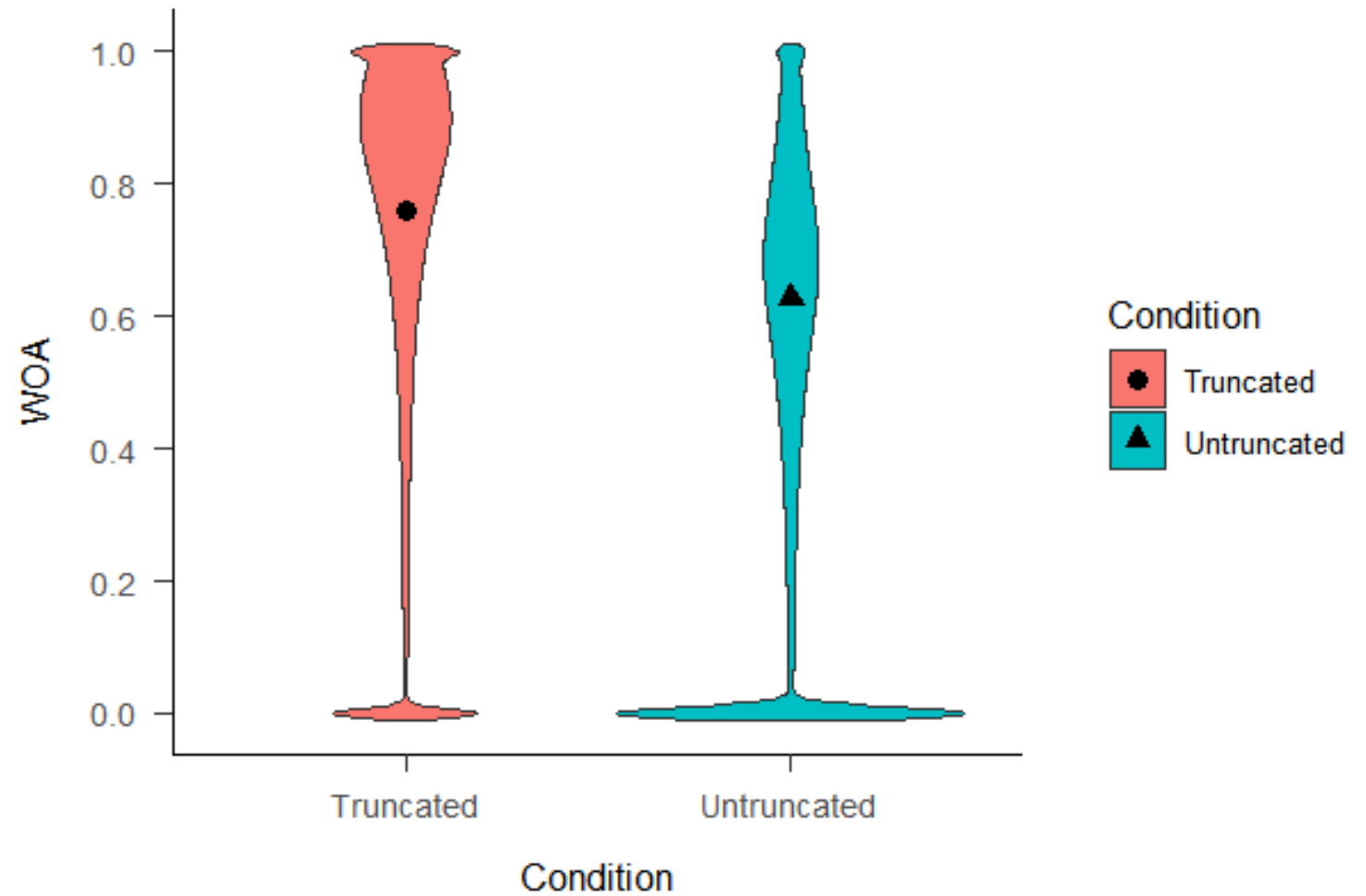
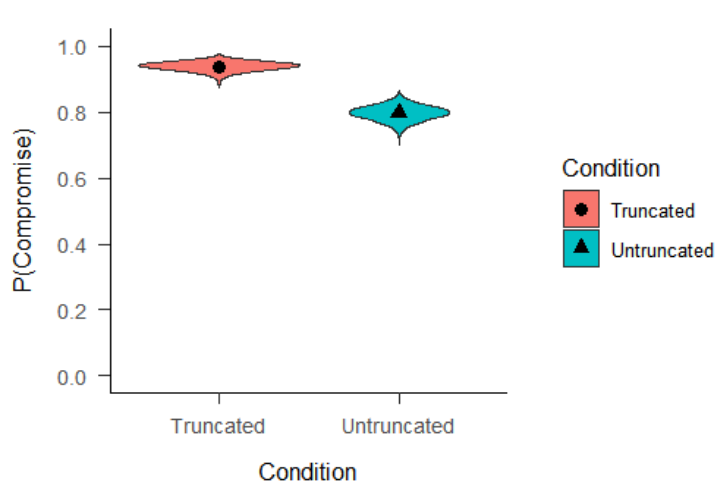
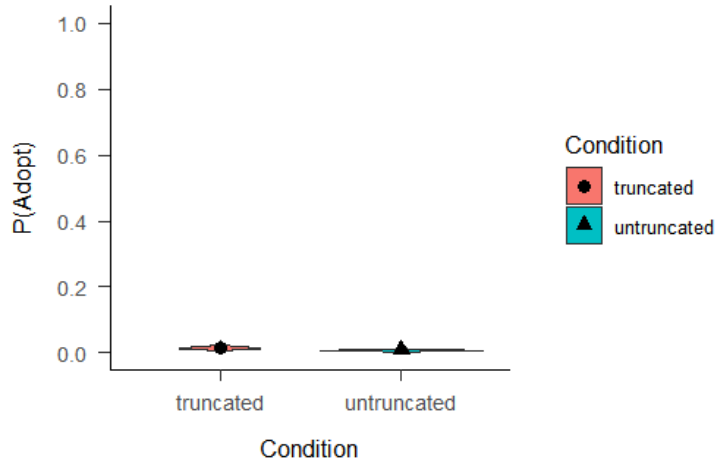
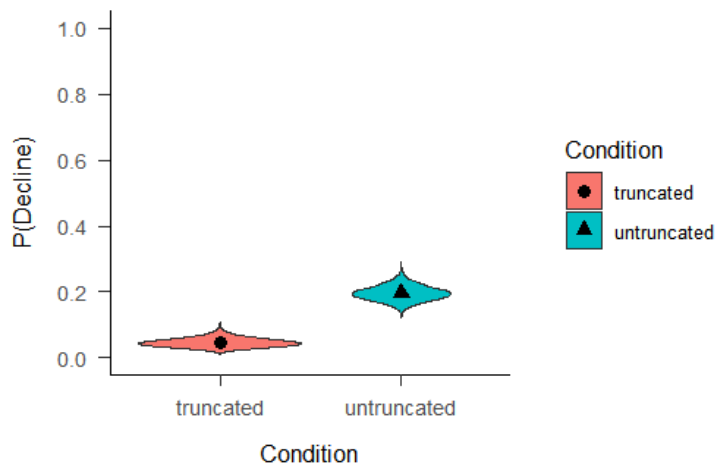
# Truncated Item Example



Your **advisor** made the following estimate: **600 Calories**



# Result: Persuasion Without Prior Belief



# Individual Differences in Decision Strategy



# Conclusions

## Methodological Insights

- Planned missing data designs can cure or reduce measurement reactivity
- Fully Bayesian imputation reliably recaptures population level parameters in planned missing data designs

## Belief Revision

- People do not anchor on their prior when it is not explicitly elicited
- People weigh advice *more* heavily in the absence of measurement reactivity
  - Possible reason for difference from Rader et al. (2015): people believed their personal knowledge was more informative in those tasks than it was for calorie estimation

# References

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